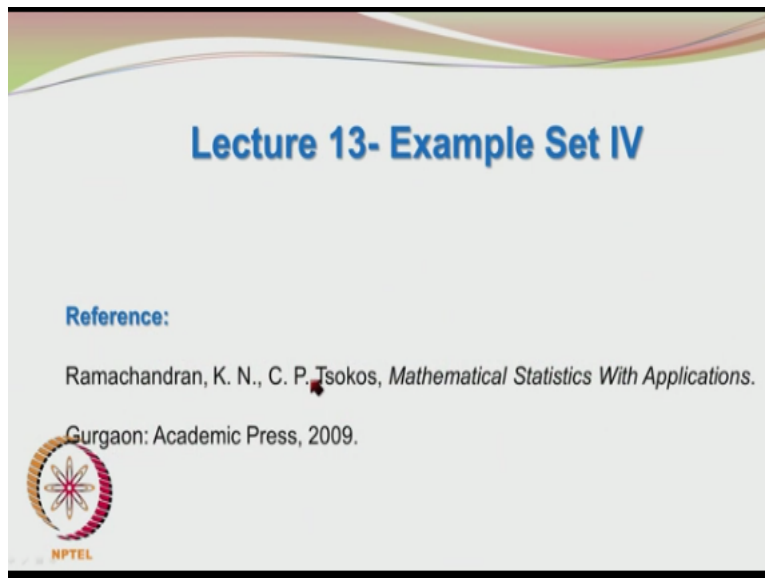


**Statistics for Experimentalists**  
**Prof. Kannan. A**  
**Department of Chemical Engineering**  
**Indian Institute of Technology – Madras**

**Lecture – 14**  
**Example Set – IV Part A**

Hello, welcome back in today's lecture, we will be solving the few problems. The reference I used for solving one of problems is the book written by Ramachandran and Tsokos, mathematical statistics with applications, academic press, published in 2009. It has an interesting set of both examples and problems. So, the topics covered in this example set are properties of random samples, applications of the central limit theorem and maximum likelihood estimation of the parameters and also the method of moments.

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The first example is 2 random samples come from 2 different populations  $P_1$  and  $P_2$ . The 2 samples are also of different sizes 9 and 25. If the 2 samples distributions however are to have the same standard deviation, what should be the ratio of their respective population standard deviations? So, we are asked to find the ratio of the population standard deviations such that the 2 unequal size to samples have the same standard deviation.

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## Topics Covered in this Example Set

- ❖ Properties of Random Samples
- ❖ Applications of the Central Limit Theorem
- ❖ Maximum Likelihood Estimation of Parameters




So, depending upon the size of the sample you can have different sampling distributions. You also know that the sampling distributions of the mean are centered around the population parameter  $\mu$  itself, but have a lesser spread given by  $\sigma^2/n$ , there  $\sigma^2$  is the variance of the population from which the random sample was taken  $n$  is the size of the sample taken. So, using this information we can do the following.

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## Example 1: Properties of a Random Sample

**Two random samples come from two different populations  $P_1$  and  $P_2$ . The two samples are also of different sizes viz. 9 and 25.**



I have given a table here. In this table, you can see the population parameters listed,  $P_1$ ,  $P_2$ ,  $\mu_1$ ,  $\mu_2$ , the 2 population means  $\sigma_1$ ,  $\sigma_2$ , the 2 population standard deviations and of course the population would hypothetically comprise of infinite size or very, very large size and when you go to the sample, again the sample probable distribution of the means will have a mean value


of  $\mu_1$  and  $\mu_2$  for sample1 and sample2 corresponding to the 2 populations from which they were taken.

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**Example 1: Properties of a Random Sample**

If the two sample distributions however are to have the same standard deviation, what should be the ratio of their respective population standard deviations?


The answers are summarized in the following table

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Standard deviations  $\sigma_1/\sqrt{n_1}$ ,  $\sigma_2/\sqrt{n_2}$  and what should be the ratio of the  $\sigma_1/\sigma_2$  such that these 2 are equal. So, the question is very simple. So,  $\sigma_1/\sqrt{n_1} = \sigma_2/\sqrt{n_2}$  and then, we have  $\sigma_1/\sqrt{n_1} \cdot \sqrt{n_1}$  would be root of 9, so that is not difficult to get  $\sigma_1/3$ ,  $\sigma_2/\sqrt{n_2}$ . What is  $n_2$ ? 25. Again that is easy to get, root of 25 was 5. So, you have  $\sigma_2/\sigma_1$  is 1.67.


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Parameter	Population Parameters		Sample Distribution Parameters	
	$P_1$	$P_2$	Sample 1	Sample 2
Mean	$\mu_1$	$\mu_2$	$\mu_1$	$\mu_2$
Standard Deviation	$\sigma_1$	$\sigma_2$	$\sigma_1/\sqrt{n_1}$	$\sigma_2/\sqrt{n_2}$
Size	$\infty$	$\infty$	9	25

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And  $\sigma_2^2 / \sigma_1^2$ , the ratio of the 2 population variances would be  $25/9$ , which is 2.78 rather than doing the mental mathematics, let us do with the calculator  $25/9$  that is 2.777 so on. So, you can truncate it to 2.78. So, the second population variance was 2.78 times more than the first population variance. But the second sample distribution was identical to that of the first as the second sample size was also higher by 2.78 times the first.

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$$\frac{\sigma_1}{\sqrt{n_1}} = \frac{\sigma_2}{\sqrt{n_2}}$$
$$\frac{\sigma_1}{3} = \frac{\sigma_2}{5}$$
$$\frac{\sigma_2}{\sigma_1} = \frac{5}{3} = 1.67$$

So, when you normalize the variances of the 2 different populations by the sample sizes taken. In this case, we were equal because we sample size taken from the second population was higher than the first. So this sort of balanced out the higher variance of the second population, okay. Let us go to the next example again the simple example you have 2 random samples,  $\bar{X}_1$  and  $\bar{X}_2$ , they come from 2 independent normal populations  $N_1, \mu_1, \sigma_1^2$  and  $N_2, \mu_2, \sigma_2^2$ .

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$$\frac{\sigma_2^2}{\sigma_1^2} = \frac{25}{9} = 2.78$$

The second population variance was 2.78 times more than the first population variance. But the second sample distribution variance was identical to that of the first as the second sample size was also higher by 2.78 times the first.



The 2 samples are also of different sizes namely  $n_1$  and  $n_2$ . So find the mean and variance of the following linear combinations  $\bar{X}_1 - \bar{X}_2$ , then  $\bar{X}_1 + \bar{X}_2$ . Very nicely, the problem statement gives us all we require. It says that the 2 parent populations are normal and they are also independent of one another. So, when you take a random sample out of these 2 populations, we have to get the random sample means that is easy.

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## Example 2: Properties of a Random Sample

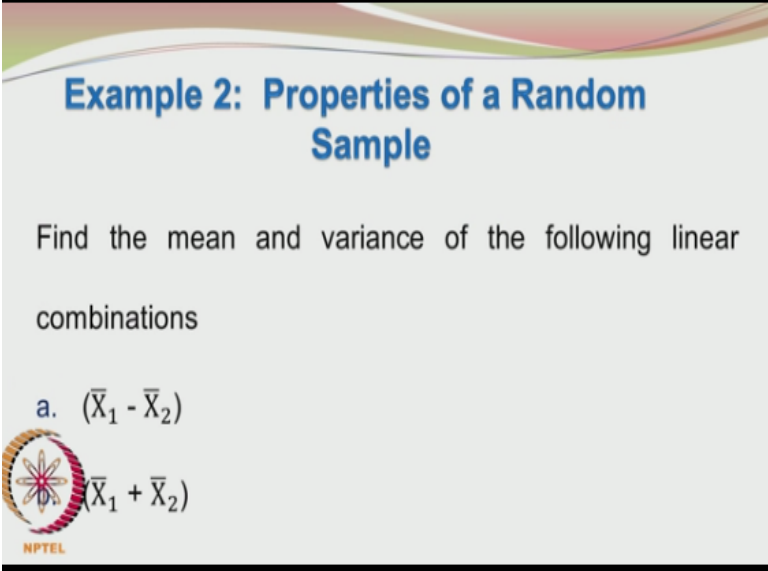
Two random samples ( $\bar{X}_1$  and  $\bar{X}_2$ ) come from two independent normal populations  $N_1 (\mu_1, \sigma_1^2)$  and  $N_2 (\mu_2, \sigma_2^2)$ . The two samples are also of different sizes viz.  $n_1$  and



So, you will have  $X_1 + X_2 + \dots + X_n$  divided by  $n$  and again  $X_2$  would be from the second population. Again, you add up all the attributes or values of the random sample elements and then divided by that particular sample size. So, you will get sample mean 1 and then you will

also get sample mean 2. The important result is, suppose you take random variables  $X_1$ ,  $X_2$ , they come from independent normally distributed populations.


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**Example 2: Properties of a Random Sample**

Find the mean and variance of the following linear combinations

a.  $(\bar{X}_1 - \bar{X}_2)$

  $(\bar{X}_1 + \bar{X}_2)$

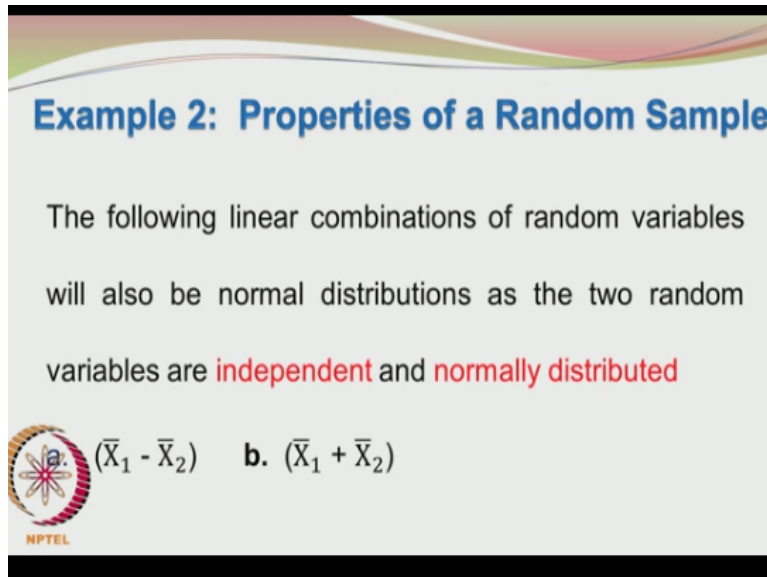
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Then, a linear combination of  $X_1$  and  $X_2$  would also be a normal distribution that is an important result. Now, we are having  $\bar{X}_1$  and  $\bar{X}_2$ .  $\bar{X}_1$  in turn is formed by taking the elements of the first sample adding all the attributes of those sample elements divide it by the sample size. Similarly, do for the second random sample. So, now you are going to combine these 2. So, rather than thinking of them as  $X$ , all the elements divided by  $n_1$ .

Then, all the elements of the second random sample divided by  $n_2$ . You think of  $\bar{X}_1$  and  $\bar{X}_2$  as random variables themselves and they are coming from 2 independent populations. So, the distributions of  $X_1$  and  $X_2$  are independent of each other and if you think on these lines, it is easier to proceed further. So, now you have to find mean and variance of the 2 linear combinations.

Why I gave this example as we encounter such crises very frequently even when different kinds of problems, okay. So, the following linear combinations of random variables will also be normal distributions as the 2 random variables are independent and normally distributed. So, these would also be normal distributions. So, this would be one normal distribution. This would be another normal distribution.

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**Example 2: Properties of a Random Sample**

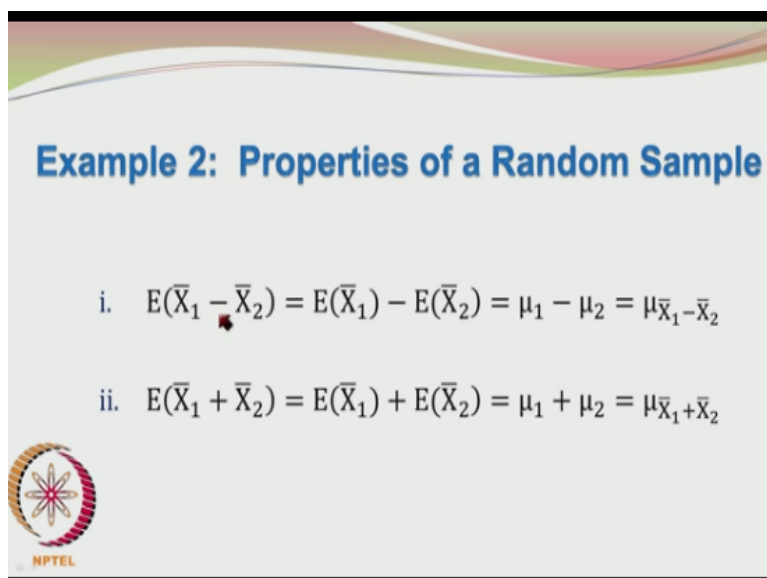
The following linear combinations of random variables will also be normal distributions as the two random variables are **independent** and **normally distributed**

a.  $(\bar{X}_1 - \bar{X}_2)$     b.  $(\bar{X}_1 + \bar{X}_2)$

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What are the mean and variances of such normal distributions for the 2 cases? So, the expected value for  $\bar{X}_1 - \bar{X}_2$  would be expected value of  $\bar{X}_1$  – expected value of  $\bar{X}_2$  that would be  $\mu_1 - \mu_2$  and that is represented as  $\mu$  of  $\bar{X}_1 - \bar{X}_2$ ,  $\mu$  of the probability distribution formed by  $\bar{X}_1 - \bar{X}_2$ . Again, you have expected value of  $\bar{X}_1 + \bar{X}_2$  that would be expected value of  $\bar{X}_1$  + expected value of  $\bar{X}_2$  that =  $\mu_1 + \mu_2$ , which is represented by  $\mu$  of  $\bar{X}_1 + \bar{X}_2$ .

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**Example 2: Properties of a Random Sample**

i.  $E(\bar{X}_1 - \bar{X}_2) = E(\bar{X}_1) - E(\bar{X}_2) = \mu_1 - \mu_2 = \mu_{\bar{X}_1 - \bar{X}_2}$

ii.  $E(\bar{X}_1 + \bar{X}_2) = E(\bar{X}_1) + E(\bar{X}_2) = \mu_1 + \mu_2 = \mu_{\bar{X}_1 + \bar{X}_2}$

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So, the linear combinations of the probability distribution of  $\bar{X}_1$  and  $\bar{X}_2$  would also result in a normal distribution which is centered at  $\mu_1 - \mu_2$ . Well, you can ask what will happen if

$\mu_1 > \mu_2$ , no problem, it is a positive value. If  $\mu_1 < \mu_2$ , it is a negative value. So what, let the resulting probability distribution be centered on a negative value, there is no harm in that. So, again if you look at expected value of  $\bar{X}_1 + \bar{X}_2$  that would be  $E$  of  $X_1 + E$  of  $X_2$ , which is  $\mu_1 + \mu_2$ .

So, when I am taking the linear combination of independent random variables which are normally distributed, I am going to get a resulting probability distribution which is also normally distributed and having the mean at the some of the means of the 2 probability distributions, I am adding. So, this is again quite straight forward. Let us look at the variance, the variance is quite interesting.


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**Example 2: Properties of a Random Sample**

iii.  $V(\bar{X}_1 - \bar{X}_2) = V(\bar{X}_1) + V(\bar{X}_2) = \sigma^2_{\bar{X}_1 - \bar{X}_2}$

iv.  $V(\bar{X}_1 + \bar{X}_2) = V(\bar{X}_1) + V(\bar{X}_2) = \sigma^2_{\bar{X}_1 + \bar{X}_2}$

When  $\bar{X}_1$  and  $\bar{X}_2$  are independent random samples the co  
variances in iii. and iv. vanish



The expected value was sign dependent depending upon what was a sign used here, but when you look at the variance, variance of  $\bar{X}_1 - \bar{X}_2$  is variance of  $\bar{X}_1 +$  variance of  $\bar{X}_2$ . Variance of  $\bar{X}_1 + \bar{X}_2$  is variance of  $\bar{X}_1 +$  variance of  $\bar{X}_2$ . So, the negative sign or positive sign does not matter. The negative sign or positive sign would really matter when you look at the covariance and here in the first case, it will be  $-$  covariance of  $\bar{X}_1$  and  $\bar{X}_2$ .

Here, it will be  $+$  of covariance of  $\bar{X}_1$ ,  $\bar{X}_2$ , but the covariance will vanish because  $\bar{X}_1$  and  $\bar{X}_2$  are independent. So, we simply have variance of  $\bar{X}_1 +$  variance of  $\bar{X}_2$  in both the cases. So, summarizing the results from this example, we have the random variable  $X_1$ ,




having a  $\mu_1$  as mean and  $\sigma_1^2/n_1$  as variance, so the standard deviation would be  $\sigma_1/\sqrt{n_1}$ .

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**EXAMPLE 2**

Random Variable	Mean	Variance	Standard Deviation	
$\bar{X}_1$	$\mu_1$	$\frac{\sigma_1^2}{n_1}$	$\frac{\sigma_1}{\sqrt{n_1}}$	
$\bar{X}_2$	$\mu_2$	$\frac{\sigma_2^2}{n_2}$	$\frac{\sigma_2}{\sqrt{n_2}}$	
$\bar{X}_1 - \bar{X}_2$	$\mu_1 - \mu_2$	$\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}$	$\sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}$	<b>Independent Distributions</b>
$\bar{X}_1 + \bar{X}_2$	$\mu_1 + \mu_2$	$\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}$	$\sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}$	



$\bar{X}_2$  for the second case, again I think it is better if I sort of go back a little bit. What is  $\bar{X}_2$ ? This is the random sample taken from a second population. The second population is normally distributed. So, you take the elements of size  $n_2$ , then add the attributes or values of these elements divided by  $n_2$ , you will get  $\bar{X}_2$  and similarly, you can take many such random samples from the second population and each one would have a different average value.

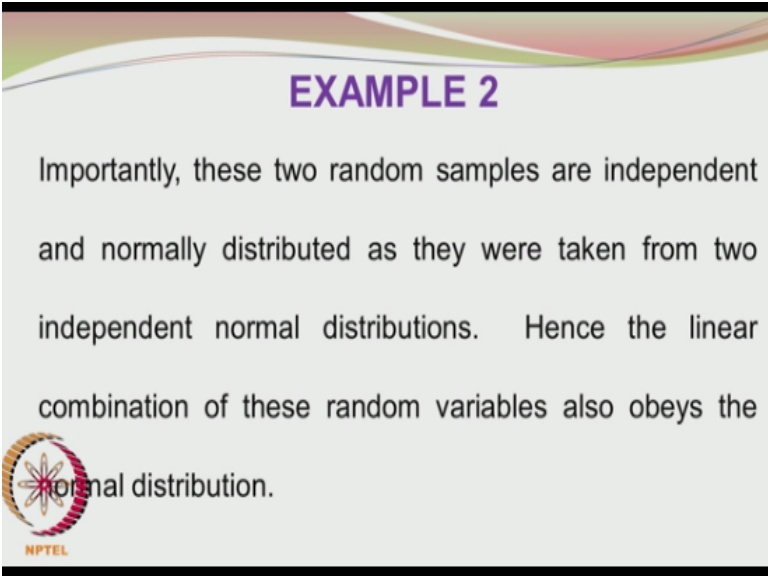
So, they will form a distribution of the sample means. This distribution of the sample means would be normal with mean at  $\mu_2$  and variance at  $\sigma_2^2/n_2$ . What is  $\mu_2$ ? It is not only the mean of the sampling distribution, but it is also the mean of the parent population from where the random sample was taken and  $\sigma_2^2$  again is the variance of the second parent population.

And the variance of the probability distribution of the sample means taken from the second population will be smaller and it will be given by  $\sigma_2^2/n_2$ . So, the standard deviations of course would be  $\sigma_1/\sqrt{n_1}$  for the first case,  $\sigma_2/\sqrt{n_2}$  for the second case and the  $\bar{X}_1 - \bar{X}_2$ , a linear combination of the 2 random variables would have a mean of  $\mu_1 - \mu_2$ , we saw that it is sign dependent.

We just saw it a couple of slides back and the variance would be variance of  $\bar{X}_1$  which is  $\sigma_1^2/n_1$  + variance of  $\bar{X}_2$  which is  $\sigma_2^2/n_2$  and so, they are added up and we take the standard deviation, it would be square root of  $\sigma_1^2/n_1 + \sigma_2^2/n_2$ . When you take  $\bar{X}_1 + \bar{X}_2$  as the linear combination of the 2 random variables.


The 2 random sample means they will be distributed around  $\mu_1 + \mu_2$  at the center and having a variance or spread given by  $\sigma_1^2/n_1 + \sigma_2^2/n_2$ . The square root of that would be  $\sigma_1^2/n_1 + \sigma_2^2/n_2$ . This applies for independent distributions. When  $\bar{X}_1$  and  $\bar{X}_2$  are independent of each other, then this results I have shown here would apply, okay.

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**EXAMPLE 2**

Importantly, these two random samples are independent and normally distributed as they were taken from two independent normal distributions. Hence the linear combination of these random variables also obeys the normal distribution.



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So importantly, I would like to re-emphasize, the 2 random samples are independent and normally distributed as they were taken from 2 independent normal distributions. Hence the linear combination of these random variables also obeys the normal distribution. Since it obeys the normal distribution, we can express this in the standard form so that we may use the probability tables.


So, when you expressed them in a standard normal form, it becomes quite straight forward,  $\bar{X}_1$  may be in turn normalized by subtracting  $\mu_1$ ,  $\bar{X}_1 - \mu_1$  divided by  $\sigma_1/\sqrt{n}$ . Let

me just correct that typo and so we have  $Z_1 = \bar{X}_1 - \mu_1$  divided by  $\sigma_1/\sqrt{n_1}$ .  $Z_2$  is  $\bar{X}_2 - \mu_2$  divided by  $\sigma_2/\sqrt{n_2}$  and if you look at the  $\bar{X}_1 - \bar{X}_2$ , you can create it as another random variable with mean  $\mu_1 - \mu_2$  and standard deviations square root of  $\sigma_1^2/n_1 + \sigma_2^2/n_2$ .

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**EXAMPLE 2**

Random Variable (X)	Standard Normal Form
$\bar{X}_1$	$Z_1 = \frac{\bar{X}_1 - \mu_1}{\sigma_1/\sqrt{n_1}}$
$\bar{X}_2$	$Z_2 = \frac{\bar{X}_2 - \mu_2}{\sigma_2/\sqrt{n_2}}$
$\bar{X}_1 - \bar{X}_2$	$\frac{(\bar{X}_1 - \bar{X}_2) - (\mu_1 - \mu_2)}{\sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}}$
$\bar{X}_1 + \bar{X}_2$	$\frac{(\bar{X}_1 + \bar{X}_2) - (\mu_1 + \mu_2)}{\sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}}$



So, the random variable combination  $\bar{X}_1 + \bar{X}_2$  may be expressed as shown here. So, this is a very nice way of putting it in a compact form and then we may use the standard normal probability tables to do the necessary calculations, right. Now let us look at example 3, the problem statement goes on like this. From historical data, the yields of power from a nuclear reactor supplied by XYZ Company are normally distributed.

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### Example 3: Power of Probability

From historical data, the yields of power from a nuclear reactor supplied by XYZ Company are normally distributed. This reactor supplied by this company is operated in several plants around the



This reactor supplied by this company is operated in several plants around the world. The population standard deviation based on process design specification is 0.7 gigawatt. The average power output of power from 6 random measurements taken at a plant using this reactor is 2 gigawatt. However, the XYZ Company had guaranteed an average power output of 2.3 gigawatt from its reactors.

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### Example 3: Power of Probability

The **population standard deviation** based on process design specification is 0.7 GW.

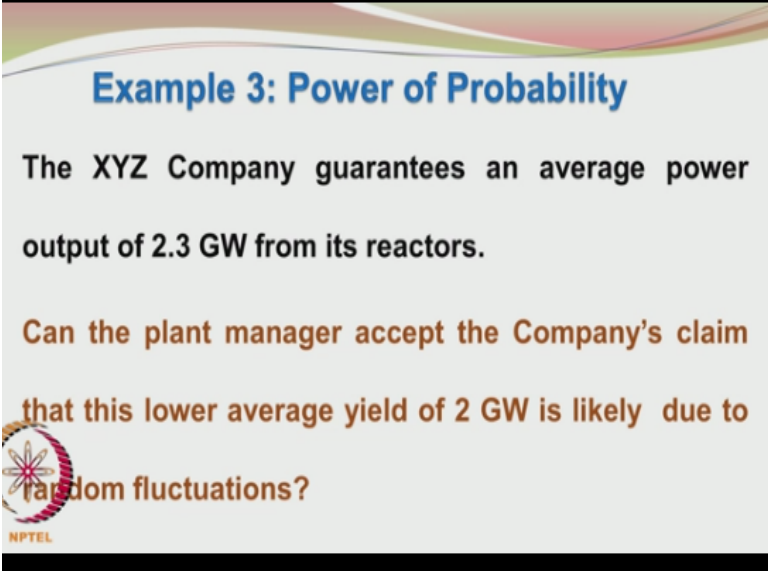
The average output of power from 6 random measurements taken at a plant using this reactor is 2



Obviously, the client organization using this reactor is getting an average power output of 2 gigawatt and it is concerned because it is supposed to produce 2.3 gigawatt, but it is producing only 2 gigawatt and that may lead to loss, okay and then the company is contacted, the company

says do not worry the thing is normal, it is only a random fluctuation or a random variation even if you are taken the means, the differences because of random fluctuation.


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**Example 3: Power of Probability**

The XYZ Company guarantees an average power output of 2.3 GW from its reactors.

Can the plant manager accept the Company's claim that this lower average yield of 2 GW is likely due to random fluctuations?


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But the company said if it is random fluctuation on the positive side, if we had got 2.6 gigawatt that would have been nice, but we are getting only 2 gigawatt whereas you are promising 2.3 gigawatt. So there is an issue here and we have to see what is the probability of the average power output from the plant being 2 gigawatt even though the actual mean value is 2.3 gigawatt. Coming again, what we have to do is there is a distribution of the sample means and the mean value is 2.3 gigawatt.

So from this sampling distribution of the means probable distribution, what is the probability of picking up a sample with a mean power output of 2 gigawatt? If the probability is quite high, then the probability of occurrence of such kind of event is quite high. So, we can only attributed to random effects. We cannot say anything more.

However, if the probability of picking up a sample of mean power output of 2 gigawatt is pretty low from sampling distribution of the mean of 2.3 gigawatt, then we have to question the supplier. So, we have to look at the sampling distribution of the mean. Since we are talking about the mean power output we are referring to the sampling distributions of the means and we also have a probability distribution.


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Parameter	Value
Population mean ( $\mu$ )	2.3 GW
Sample mean ( $\bar{x}$ )	2 GW
Population standard deviation ( $\sigma$ )	0.7 GW
Sample size	6

So, the population mean has given as 2.3 gigawatt, sample mean  $\bar{x}$ , I am using small  $\bar{x}$  because sample has been taken and its value is known that is 2 gigawatt only. Population standard deviation based on design specification is 0.7 gigawatt, having the same units as the mean power output and sample size is only 6. So, it is given that the population is normal and the value of  $\sigma$  is also known which makes life easier for us.

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**Solution:**

It is noted that the population standard deviation is known and the sample size is rather small at 6.

However, the population is normal and the value of  $\sigma$  is known.


So, we have to find out the probability of the power output being  $<$  or  $=$  2 megawatts from the given data and  $\bar{X}$ , I am normalizing it again  $\bar{X} - \mu / \sigma / \sqrt{n}$ ,  $2 - 2.3$  divided by  $0.7 / \sqrt{6}$ , let me sort of check it out. So, I should be doing  $-0.3 * \sqrt{6}$  divided by, so I am

getting -1.04978, -1.05 is okay and so, what is the probability that X bar would be  $\leq 2$ , which is equivalent to asking what is the probability of the standard normal variable  $Z \leq -1.05$ .

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**Solution:**

We can now find out the probability of the power output being less than or equal to 2 MW from the given data.




And the probability is 0.147, so the probability of the sampled mean being lower than or  $= 2$  gigawatt is rather high at 0.15, okay. So, the company is saying the mean power output is 2.3 gigawatt. It is not stopping there; it is also saying that the standard deviation of the normal distribution is 0.7 gigawatt. Now, we are taking about the sampling distribution of the means, the probability distribution of the sample means.

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Random Variable	Standard normal Form	Value
$\bar{X}$	$\frac{\bar{X} - \mu_1}{\frac{\sigma}{\sqrt{n}}}$	$\frac{2 - 2.3}{\frac{0.7}{\sqrt{6}}} = -1.05$
$P(\bar{X} \leq 2)$	$P(Z \leq -1.05)$	0.147

The probability of the sampled mean being lower than or equal to 2 GW is rather high at 0.15.

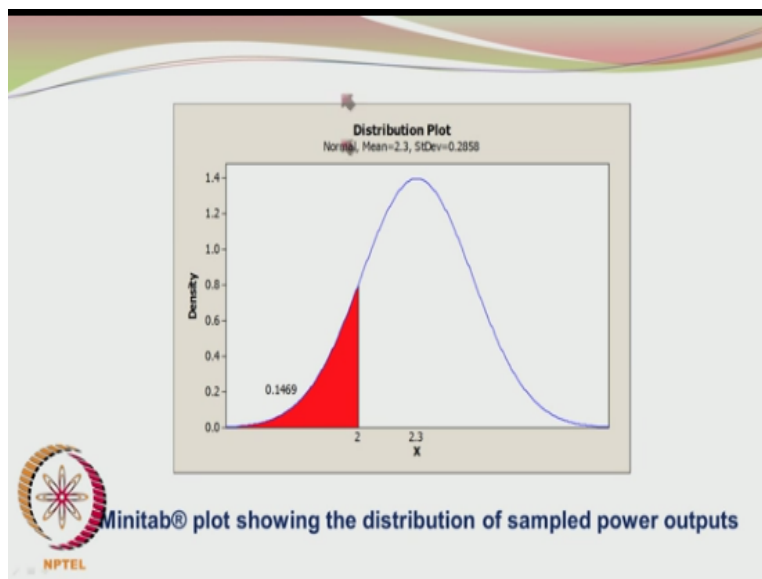


And the probability distribution of the sample means is centered again at 2.3 gigawatt and having the spread given by  $0.7/\sqrt{6}$ . So, what is  $0.7/\sqrt{6}$ ? 0.286. So, there is a spread of 0.286 gigawatt around this particular sampling distribution of the mean. So, the standard deviation is 0.286 gigawatt.

The company is getting only 2 gigawatt, so when we do the calculations for the probability of this occurrence, namely the occurrence of 2 gigawatt or lower when the sample is taken from a sampling distribution of the means centered at 2.3 gigawatt and standard deviation of 0.286 gigawatt. The probability comes to 0.147 which is rather high. So, you really cannot question the supplier because 0.15 is a good reasonable chance of occurrence of this kind of event.

So, if you do the plotting with the Minitab, this is the normal distribution centered at 2.3 gigawatt and having the standard deviation of  $0.7/\sqrt{6}$  which is 0.286 gigawatt. So, this is the spread and I am looking at the probability of occurrence of 2 gigawatt or lower from this probability distribution and I am finding the probability, the area under the curve in the shaded region which comes to 0.147.

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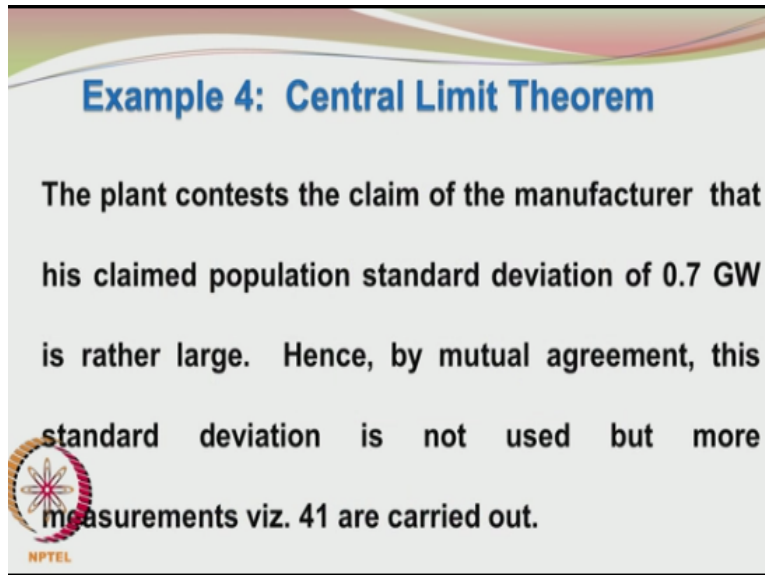


So, moving onto the next example, the plant contests the claim of the manufacturer that his claimed population standard deviation of 0.7 gigawatt is rather large. Hence, by mutual agreement, this standard deviation is not used but more measurements names 41 are carried out,




okay. So that 0.7 gigawatt is thrown out of the window and you are no longer even thinking of the population being normally distributed that is not mentioned in the problem statement.

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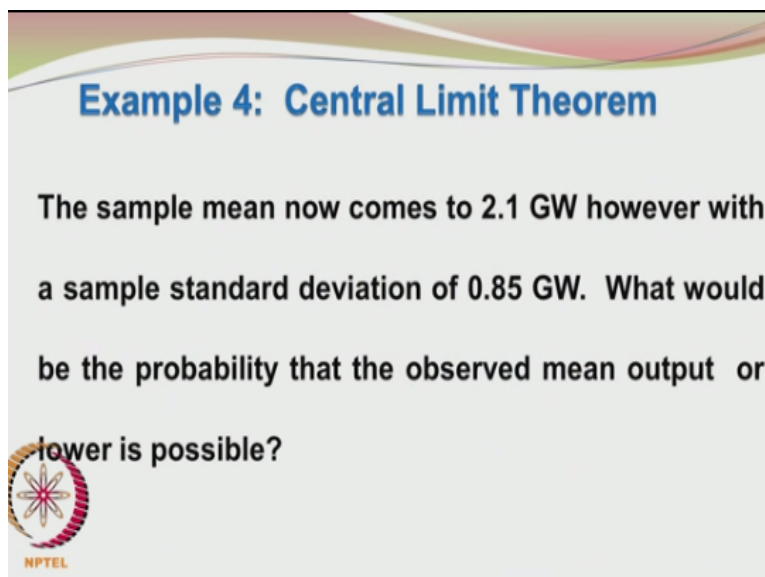
**Example 4: Central Limit Theorem**

The plant contests the claim of the manufacturer that his claimed population standard deviation of 0.7 GW is rather large. Hence, by mutual agreement, this standard deviation is not used but more measurements viz. 41 are carried out.




Whereas in the previous problem statement, it was given that the population was normally distributed. But you are also taking a large sample size of 41. The sample mean now comes to a slightly higher 2.1 gigawatt, but the sample standard deviation is 0.85 gigawatt. The sample standard deviation is even higher than the design specification value of 0.7 gigawatt. What would be the probability that the observed mean output or lower is possible?

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**Example 4: Central Limit Theorem**

The sample mean now comes to 2.1 GW however with a sample standard deviation of 0.85 GW. What would be the probability that the observed mean output or lower is possible?



What is the probability of this occurrence? That you can get a sample mean of 2.1 gigawatt or lower that is what we have to find now. So, conditions are slightly changed. Population mean value  $\mu$  is 2.3 gigawatt, sample mean  $\bar{X}$  is 2.1 gigawatt, sample standard deviation  $s$  is 0.85 gigawatt, sample size is 41. We are no longer using the population standard deviation of 0.7 gigawatt.

**(Refer Slide Time: 26:18)**



Parameter	Value
Population mean ( $\mu$ )	2.3 GW
Sample mean ( $\bar{x}$ )	2.1 GW
Sample standard deviation ( $s$ )	0.85 GW
Population-standard deviation ( $\sigma$ )	0.7 GW (not used)
Sample size	41

So we are not supposed to use sigma, but we can use  $s$ , the sample standard deviation. When  $s$  is used that is permitted because the sample size is quite large, we can even continue with the normal distribution according to the central limit theorem. The central limit theorem says that irrespective of the population probability distribution characteristics. If a large sample is taken typically  $> 30$ , then the resulting sampling distribution of the means is also normal.

**(Refer Slide Time: 26:39)**

It is noted that the population standard deviation is not used but the sample size is a healthy 41.

Hence we may substitute  $s$  for  $\sigma$  and find out the probability of the power output being less than or equal to 2.1 GW




from the sampling distribution.

In the present case, the parent population we do not have to worry about because the sample size is quite large and so the central limit theorem will apply and so the sampling distribution of the means is going to be normal and since we are going to use  $s$ , because  $\sigma$  is not available for use. The  $s$  value may be substituted for  $\sigma$  in the calculations and we are also having a large sample size of 41 to account for it.

So, the problem calculations are quite straight forward instead of using  $\sigma$ , here we use  $s$ , we have  $\bar{X} - \mu/s/\sqrt{n}$  and that is  $2.1 - 2.3$  that is  $-0.2 * \sqrt{41}$  divided by  $0.85$  that comes to  $-1.5066$ ,  $-1.51$ . So, the probability of  $\bar{X} < 2$  is equivalent to probability of  $Z < -1.51$  and the probability has now considerably reduced to  $0.066$ . So the results show that the probability of the sample having power output  $< \text{or} = 2.1$  gigawatt may occur only  $6.6\%$  of the time or the probability value is  $0.066$ ,

**(Refer Slide Time: 28:02)**

Random Variable	Standard normal Form	Value
$\bar{X}$	$\frac{\bar{X} - \mu}{\frac{s}{\sqrt{n}}}$	$\frac{2.1 - 2.3}{\frac{0.85}{\sqrt{41}}} = -1.5066$
$P(\bar{X} \leq 2)$	$P(Z \leq -1.5066)$	0.066




So, we stop here and left to the 2 parties to take it from here, okay. So showing this on the normal probability distribution, here we have a standard deviation of 0.13275, how did that come about that was s used is 0.85 divided by root 41. So, 0.85/root 41 is 0.13275. The mean value hypothesized or taken as 2.3 gigawatt, so that is what we have here. So, the probability of occurrence of 2.1 gigawatt are lower is given by the area of the shaded portion and that is 0.066.

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The results show that the probability of the sample having power output less than or equal to 2.1 GW may occur only 6.6% of the time.

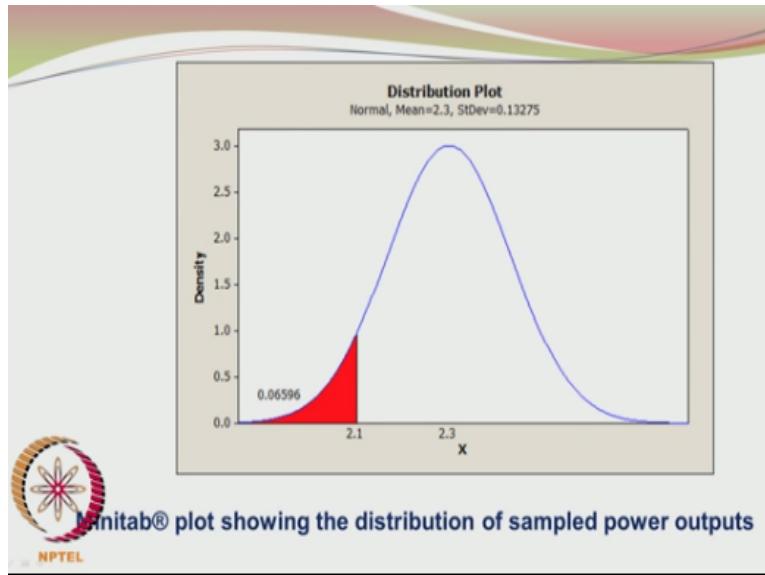
It is left to the two parties to take it from here.



So, the probability is 0.066. So, the 2 probability distributions are plotted as shown in this figure generated from Minitab, so you are having 2 probability distributions. The first one is centered at 2.3 gigawatt and has a standard deviation of 0.2858. How did this 0.2858 come about, it was 0.7

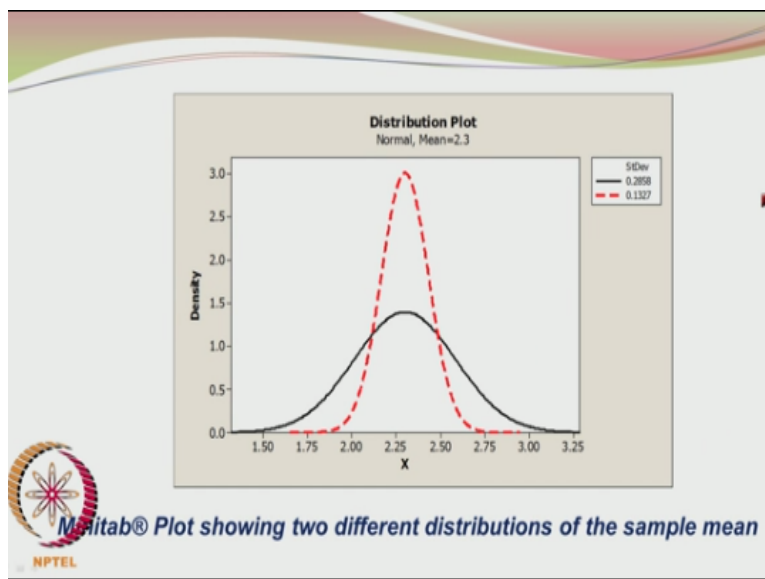
gigawatt divided by root 6. The design specification of sigma was 0.7 gigawatt and the sample size was 6 in the first case.

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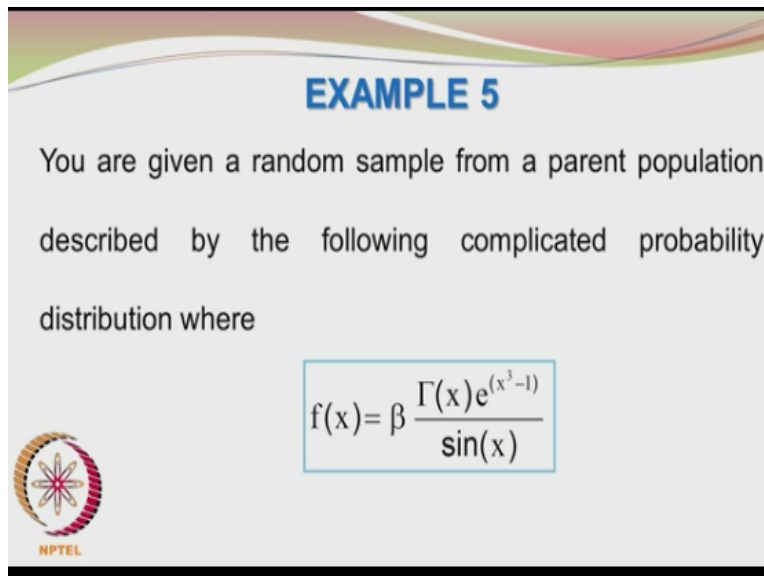
So, we are having 0.7 divided by root 6 which is 0.2858. The second distribution shown is having a lesser spread and it is also centered at 2.3 gigawatt. It is based on a sample standard deviation of 0.85 gigawatt which is higher than 0.7 gigawatt and still the spread this smaller because of the larger sample size. So, instead of using  $\sigma/\sqrt{n}$ , we are using  $s/\sqrt{n}$ , where  $s$  is 0.85 gigawatt and  $n$  is 41.

**(Refer Slide Time: 29:50)**




So, 0.85 divided by root 41 is 0.1327 which is more than half of the earlier spread value of 0.2858. So you can see as lesser spread here and also the probability value declined. The next problem is you are given a random sample from a parent population described by the complicated probability distribution function where  $f(x)$  is  $\beta \cdot \Gamma(x) \cdot e^{x^3-1} \cdot \frac{1}{\sin(x)}$ .

**(Refer Slide Time: 31:36)**



**EXAMPLE 5**

You are given a random sample from a parent population described by the following complicated probability distribution where

$$f(x) = \beta \frac{\Gamma(x) e^{(x^3-1)}}{\sin(x)}$$


So,  $\beta$  is an adjustable constant such that the probability distribution is a valid one, you know what it means. The area under the curve for any probability distribution function should be one. For continuous probability density functions described by a smooth curve and the area under such curves should be  $= 1$ . So, we adjust the parameter  $\beta$  such that this is a valid probability distribution function.

**(Refer Slide Time: 32:14)**

## EXAMPLE 5

$\beta$  is an adjustable constant such that the probability distribution is a valid one. Let mean and standard deviation of this distribution be  $\phi$  and  $\psi$ . If the sample size was chosen as 64,



Let the mean and standard deviation of this distribution be  $\phi$  and  $\psi$  that means the variance of this distribution  $\psi$  squared. If the sample size was chosen as 64, find the mean and standard deviation of the sampling distribution of the means. What is the form of the sample mean distribution? And what is the probability that the sample mean will be within 0.15 standard deviations of the population mean?

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## EXAMPLE 5

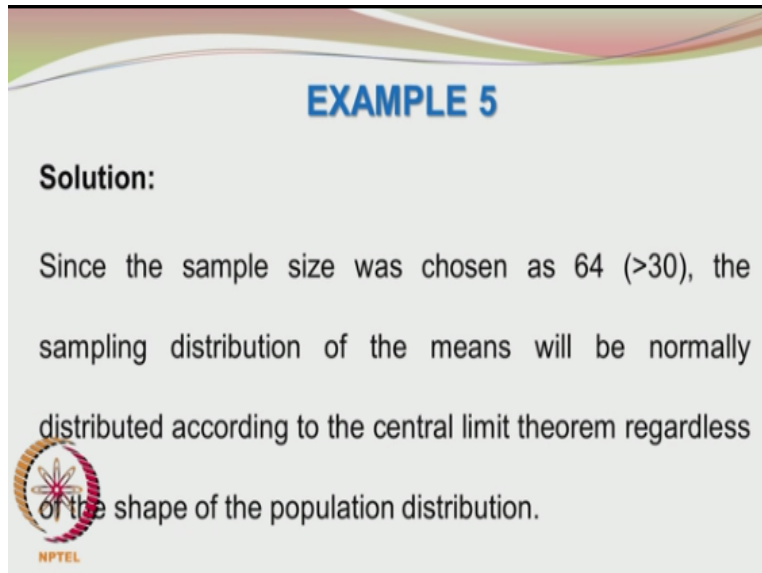
- a. Find the mean and standard deviation of the sampling distribution of the means.
- b. What is the form of the sample mean distribution?
- c. What is the probability that the sample mean will be within 0.15 standard deviations of the population mean?



So, since the sample size is quite large at 64 which is  $> 30$ , the sampling distribution of the means will be normally distributed according to the central limit theorem regardless of the shape of the parent population distribution. So, now the problem is quite straight forward. The mean of

this distribution of sample means will be  $\phi$  and the standard deviation will be  $\psi/\sqrt{64}$  which is  $0.125\psi$ .

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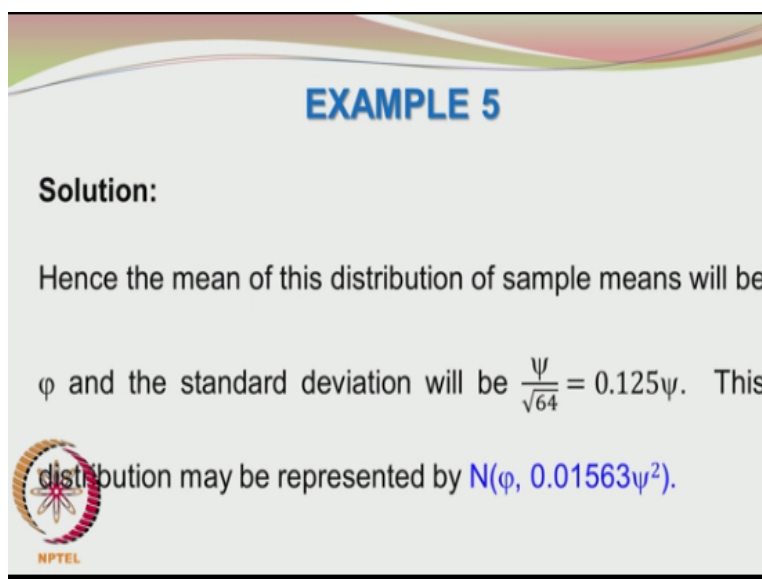
**EXAMPLE 5**

**Solution:**

Since the sample size was chosen as 64 (>30), the sampling distribution of the means will be normally distributed according to the central limit theorem regardless of the shape of the population distribution.

So,  $1/\sqrt{64}$  is  $1/8$  which is  $0.125$ . So the standard deviation of the sampling distribution of the means would be  $0.125\psi$ . This distribution may be represented by a normal distribution of mean  $\phi$  and variance which will be square of this  $0.01562\psi^2$ , okay.  $0.125^2$  square, let us confirm  $0.125^2$  is  $0.015625$ , so that is fine. What is the probability that the sample mean will be within  $0.15$  standard deviations from the population mean.

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**EXAMPLE 5**

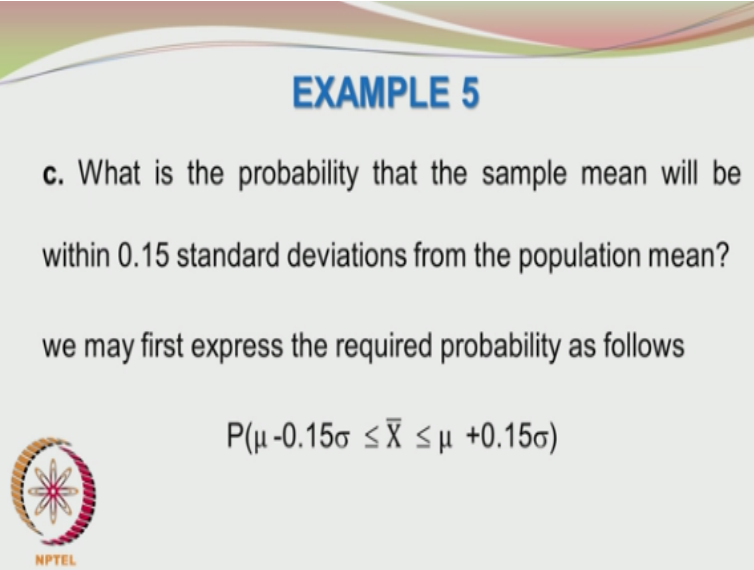
**Solution:**

Hence the mean of this distribution of sample means will be  $\phi$  and the standard deviation will be  $\frac{\psi}{\sqrt{64}} = 0.125\psi$ . This distribution may be represented by  $N(\phi, 0.01563\psi^2)$ .



So, the problem can be expressed in the following way. Probability of the value of the random sample being 0.15 sigma distant from the population mean. So, probability of mu which is the population mean and also the random sample probability distribution mean  $\mu - 0.125 \sigma < \text{or} = \bar{X} < \text{or} = \mu + 0.15 \sigma$ . So, the random sample which we take may have a value either lower than mu or higher than mu.


**(Refer Slide Time: 34:01)**



**EXAMPLE 5**

c. What is the probability that the sample mean will be within 0.15 standard deviations from the population mean?

we may first express the required probability as follows

$$P(\mu - 0.15\sigma \leq \bar{X} \leq \mu + 0.15\sigma)$$


And it may lie either on the right hand side of mu or on the left hand side of mu. So, now it is easy to normalize and how do we normalize, we just subtract mu from x bar and divided by  $\sigma/\sqrt{n}$ , we do it in all the other 2 sides of the inequality and then we get  $-0.15 \sigma/\sigma/\sqrt{n}, +0.15 \sigma/\sigma/\sqrt{n}$  and this works out probability of  $-1.2 < \text{or} = Z < \text{or} = 1.2$  and this comes to 0.77, okay that can be read of from the standard normal probability charts.

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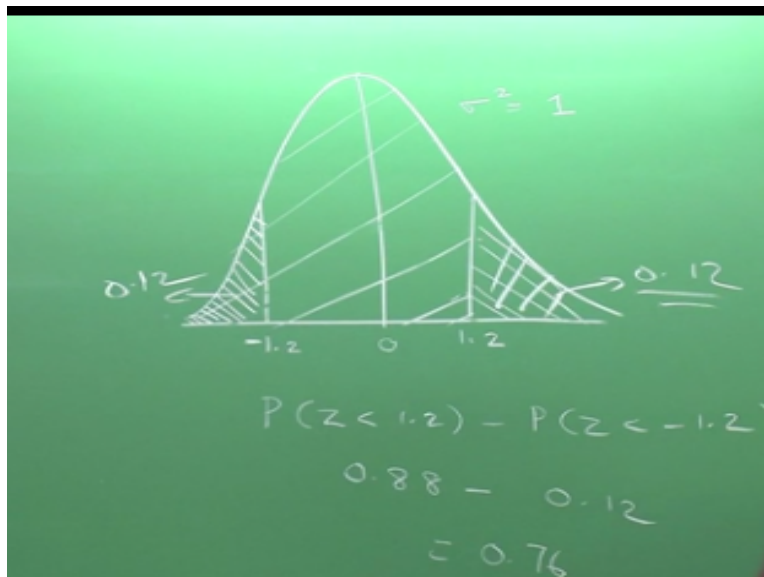
## EXAMPLE 5

$$\begin{aligned} &= P\left(\frac{-0.15\sigma}{\sigma/\sqrt{n}} \leq \frac{\bar{X} - \mu}{\sigma/\sqrt{n}} \leq \frac{+0.15\sigma}{\sigma/\sqrt{n}}\right) \\ &= P(-0.15\sqrt{64} \leq Z \leq 0.15\sqrt{64}) \\ &= P(-1.2 \leq Z \leq 1.2) = 0.77 \end{aligned}$$



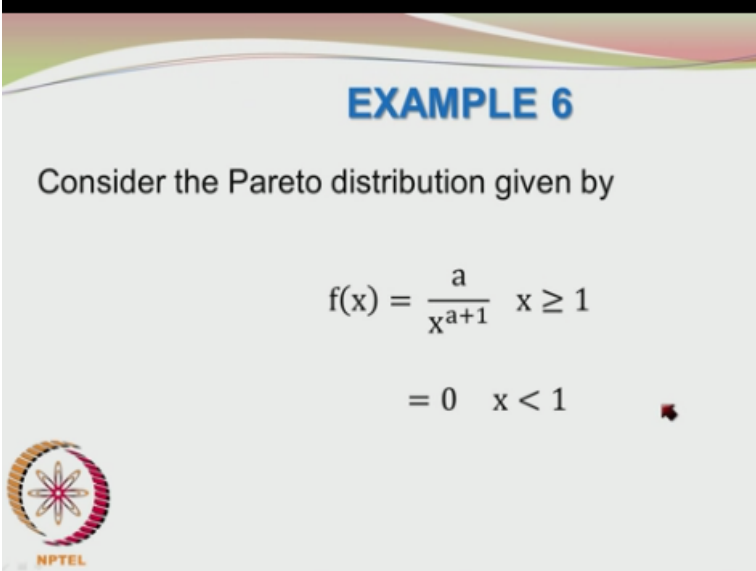
I hope now you are comfortable using these charts and you should be able to figure out how we get this 0.77. I just illustrate this on the board, so you have the standard normal distribution which is having a mean value of 0 and variance sigma square = 1 and we have to find the area under the curve 1.2, -1.2. So, what we can do is probability of  $Z < 1.2$  – probability of  $Z < -1.2$ .

**(Refer Slide Time: 36:06)**




So, first what we do is we find the area under the entire curve and then, from this total area, we subtract out this area and we get the required probability. If I remember right this comes to around 0.88 and then this would be 0.12. If the entire area is around 0.88, then this area would be 0.12 and by symmetry, this area would also be = this area would be 0.12. So,  $0.88 - 0.12$  is 0.76, I am just doing it from memory and you can also see the answer is to 0.77.

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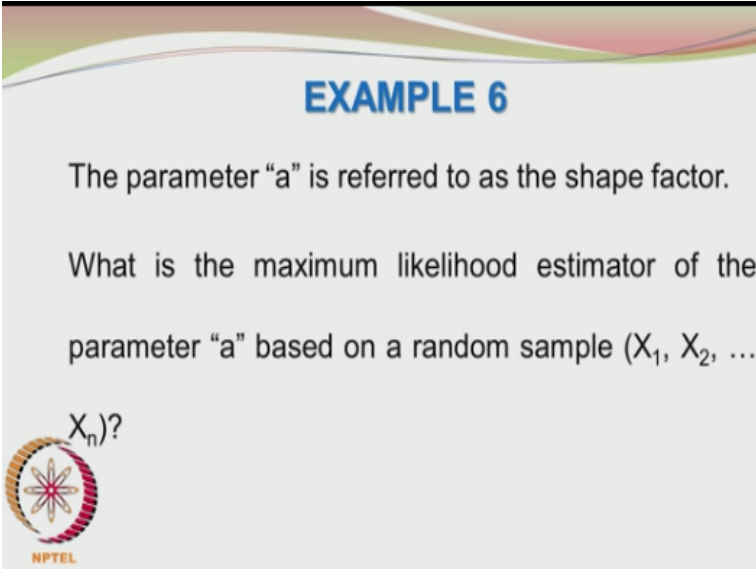
**EXAMPLE 6**

Consider the Pareto distribution given by

$$f(x) = \frac{a}{x^{a+1}} \quad x \geq 1$$
$$= 0 \quad x < 1$$


Let us move onto the next problem. Here, we have the Pareto distribution, quite an interesting function. This was the problem I had taken from the Ramachandran and Tsokos book and  $f$  of  $x = a/x$  power  $a+1$ ,  $x > \text{or} = 1 = 0$  for  $x < 1$ . So, the parameter “ $a$ ” is referred to as the shape factor. What is the maximum likelihood estimator of the parameter “ $a$ ” based on the random sample  $X_1, X_2, \dots, X_n$ .


(Refer Slide Time: 38:10)



**EXAMPLE 6**

The parameter “ $a$ ” is referred to as the shape factor.

What is the maximum likelihood estimator of the parameter “ $a$ ” based on a random sample  $(X_1, X_2, \dots, X_n)$ ?

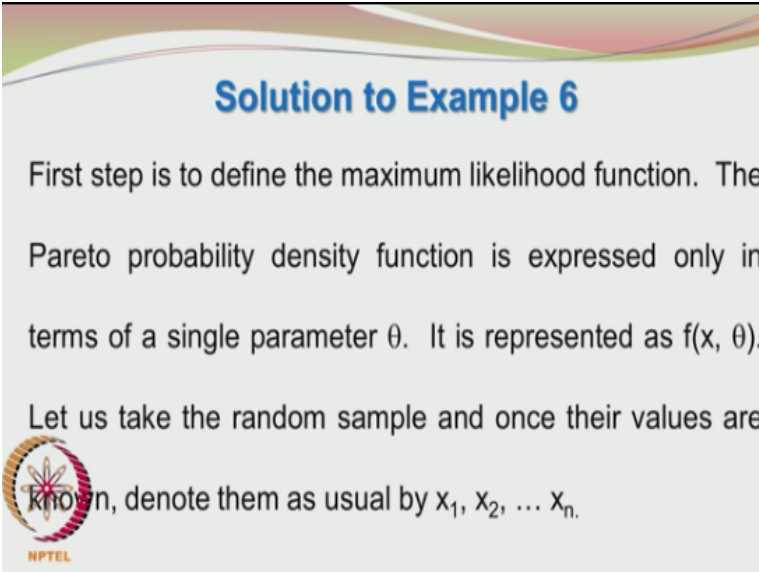


Some of you may ask we do not know the value of  $a$  and we do not know whether this is a valid probability density function, so finding the area under the curve from 1 to infinity  $a/x$  power  $a+1$   $dx$  should tell us the value of  $a$ . So, what is the additional need for finding the value of  $a$ . I leave

it to you, okay. The hint is you cannot find out  $a$  using this method for the simple reason that no matter what value of  $a$ , you plug in there, the integral 1 to infinity will be  $= 1$ .

I mean do the integration, you can find out this will be  $X$  power  $-a-1$ , so it will be  $-1/x$  power  $a$  and  $a$  would cancel out and so, when you go from 1 to infinity, it would be  $1-0$ , 1 power  $a$  is always going to be 1. I requested to do the integrations yourself and confirm that no matter what the value of  $a$  is, the  $a$  will cancel out and so, this area under the curve will always be  $= 1$ . So, let us move onto the actual problem.


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**Solution to Example 6**

First step is to define the maximum likelihood function. The Pareto probability density function is expressed only in terms of a single parameter  $\theta$ . It is represented as  $f(x, \theta)$ .

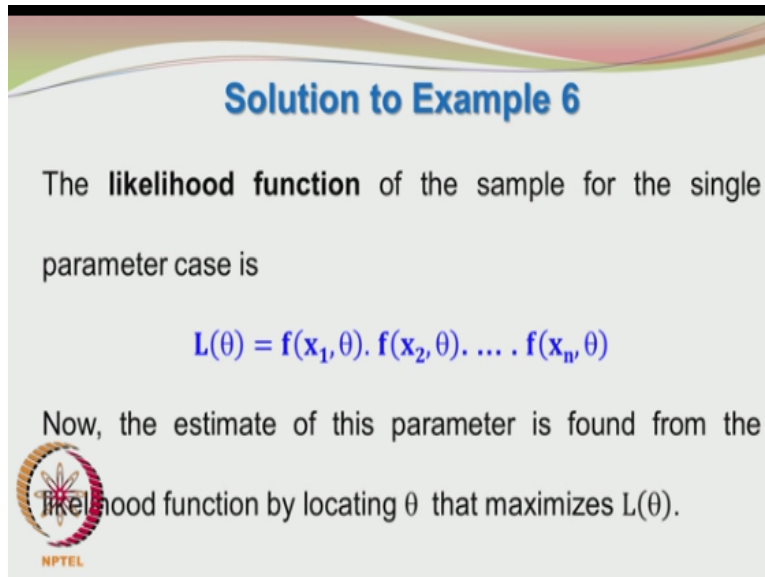
Let us take the random sample and once their values are known, denote them as usual by  $x_1, x_2, \dots, x_n$ .

 NPTEL

We have to define the maximum likelihood function. We are using the method of maximum likelihood parameter estimation method to find out what  $a$  is? The Pareto probability density function is expressed only in terms of a single parameter  $\theta$ . It is represented as  $f$  of  $x$ ,  $\theta$ . Let us take a random sample and once their values are known, will denote them by  $X_1, X_2$ , so onto  $X_n$ .

So, the likelihood function of the sample for the single parameter case is  $L$  of  $\theta = f$  of  $X_1$ ,  $\theta * f$  of  $X_2$ ,  $\theta$  so on to  $f$  of  $X_n$ ,  $\theta$ . So, we have to estimate this parameter by maximizing this relationship. So, first let us get the relationship,  $L$  of  $\theta = f$  of  $X_1$ ,  $\theta * f$  of  $X_2$ ,  $\theta$  so on to  $f$  of  $X_n$ ,  $\theta$  and that would be  $a/X_1$  to the power of  $a+1 * a/X_2$  to the power of  $a+1$  so on to  $a/X_n$  to the power  $a+1$ .

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**Solution to Example 6**

The **likelihood function** of the sample for the single parameter case is

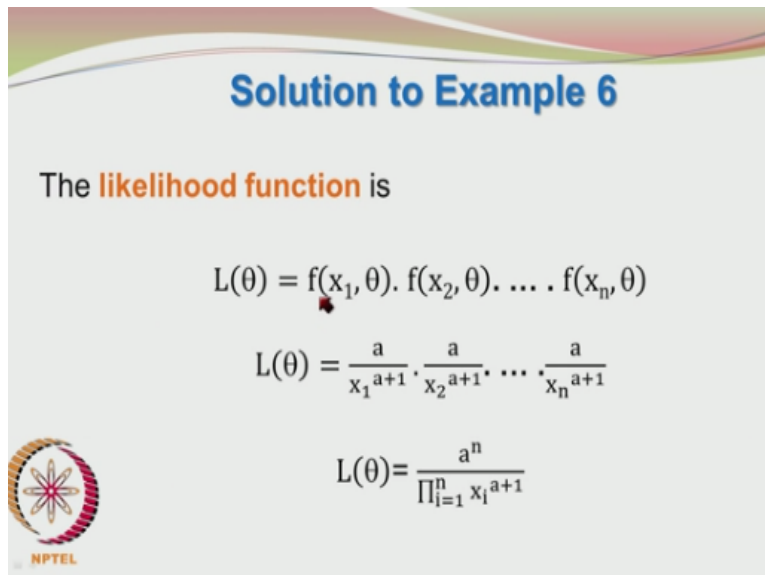
$$L(\theta) = f(x_1, \theta) \cdot f(x_2, \theta) \cdot \dots \cdot f(x_n, \theta)$$

Now, the estimate of this parameter is found from the likelihood function by locating  $\theta$  that maximizes  $L(\theta)$ .




So,  $L$  of  $\theta = a$  power  $n$ , because I am doing it in  $n$  times and this is the product of all the  $X$  values to the power of  $a+1$  and when we take natural logarithm on both sides, we get  $\ln$  of  $L = \ln$  of  $f$  of  $X_1, \theta * f$  of  $X_2, \theta$ , so onto  $f$  of  $X_n, \theta$ . So,  $\ln L = \ln$  of a power  $n$  / the product of the entities  $X_i$  to the power of  $a+1$ ,  $i$  running from 1 to  $n$ . So we take  $\ln L$ , we have this we can split into 2 parts,  $\ln$  of a power  $n$  becomes  $n \ln a$ .

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**Solution to Example 6**

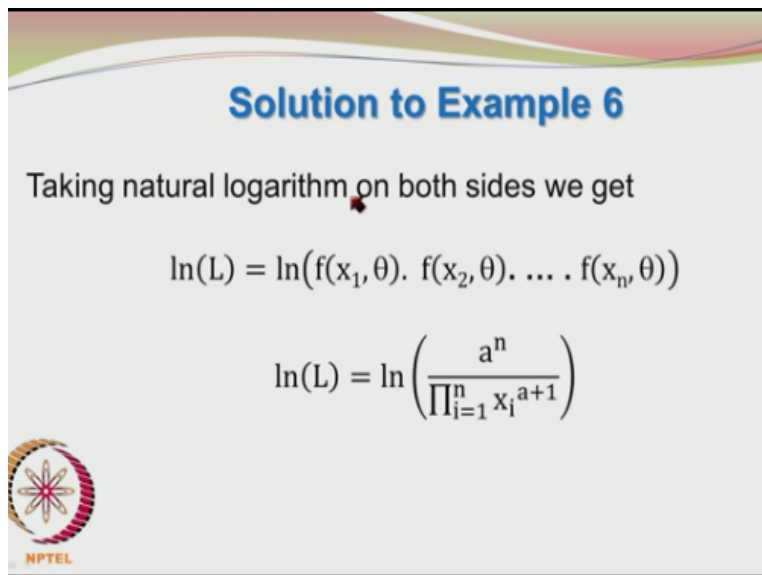
The **likelihood function** is

$$L(\theta) = f(x_1, \theta) \cdot f(x_2, \theta) \cdot \dots \cdot f(x_n, \theta)$$
$$L(\theta) = \frac{a}{x_1^{a+1}} \cdot \frac{a}{x_2^{a+1}} \cdot \dots \cdot \frac{a}{x_n^{a+1}}$$
$$L(\theta) = \frac{a^n}{\prod_{i=1}^n x_i^{a+1}}$$


And then this becomes  $\ln$  of product of  $X_i$ 's to the power of  $a+1$ . So again this is quite simple, you will get  $\ln$  of  $L = n \ln a$ , we saw this earlier. How did this get simplified? You know that the log of product of entities, the sum of  $\ln$  of those entities, so the  $a+1$  is common here and you can

put  $a+1$  here and then you get  $\sum_{i=1}^n \ln$  of  $X_i$ . The next step is to differentiate this function with respect to  $a$  and then equate it to 0.


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**Solution to Example 6**

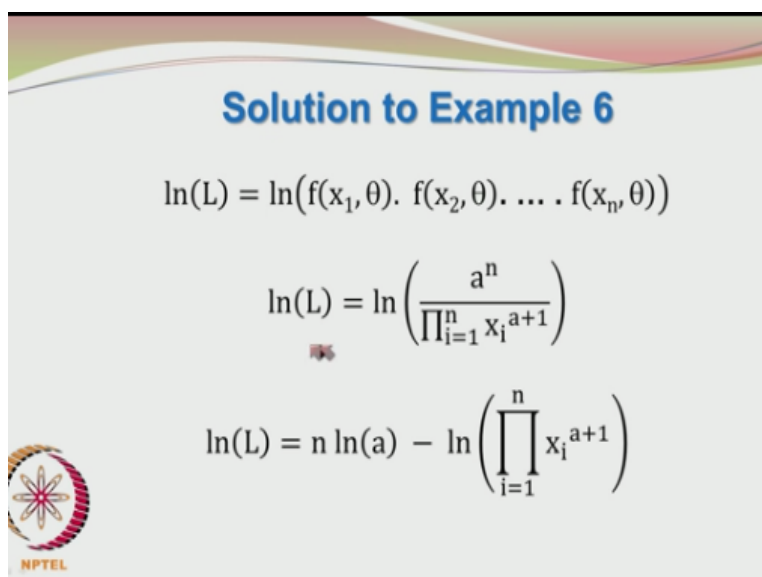
Taking natural logarithm on both sides we get

$$\ln(L) = \ln(f(x_1, \theta) \cdot f(x_2, \theta) \cdot \dots \cdot f(x_n, \theta))$$

$$\ln(L) = \ln\left(\frac{a^n}{\prod_{i=1}^n x_i^{a+1}}\right)$$


And when you differentiate with respect to  $a$ , this becomes  $n/a$  and here, we had  $a+1$ , there was no  $a$  inside, so that became quite simple,  $-1 \cdot \sum \ln X_i$ . So, the estimated parameter  $a$  is given by  $n$  divided by  $\sum_{i=1}^n \ln$  of  $X_i$ , so quite simple. Let us move onto the next problem. Use the method of moments to find the parameter estimators of the following probability distribution function.


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**Solution to Example 6**

$$\ln(L) = \ln(f(x_1, \theta) \cdot f(x_2, \theta) \cdot \dots \cdot f(x_n, \theta))$$

$$\ln(L) = \ln\left(\frac{a^n}{\prod_{i=1}^n x_i^{a+1}}\right)$$

$$\ln(L) = n \ln(a) - \ln\left(\prod_{i=1}^n x_i^{a+1}\right)$$


F of  $x = 1/B-A = 0$  otherwise. So, we have to estimate both A and B. We are going to use the method of moments, so f of  $x = 1/B-A$  and the first moment E of X is obtained from the distribution in the following manner, expected value of  $X = A$  to  $B$ ,  $x dx/B-A$  which is  $x^2/2$ , so  $B^2 - A^2/2$ ,  $B+A * B-A/B-A$ , so B-A will cancel out. So, we have  $B+A/2$  and expected value of  $X^2$ , the second moment is given by  $x^2 dx/B-A = x^3/3$ .

**(Refer Slide Time: 41:52)**

**Solution to Example 6**

$$\ln(L) = n \ln(a) - \ln\left(\prod_{i=1}^n x_i^{a+1}\right)$$

$$\ln(L) = n \ln(a) - (a+1) \sum_{i=1}^n \ln(x_i)$$

Next differentiate this function with respect to "a" and equate it to zero

$X^3/3$  will become  $B^3 - A^3$  and so, you are having  $B^3 - A^3$  divided by  $B - A$  which is  $B^2 + BA + A^2$  and that is what we have here. So, these distribution moments may be equated with the first and second sample moments and when we do that we get  $m_1$  as  $1/n \sum_{i=1}^n X_1 + X_2 + \dots + X_n$ . We will just correct the typo. So  $m_1 = 1/n \sum_{i=1}^n X_i = A + B/2$ .

**(Refer Slide Time: 42:23)**

## Solution to Example 6

$$\ln(L) = n \ln(a) - \ln\left(\prod_{i=1}^n x_i^{a+1}\right)$$

$$\frac{d(\ln(L))}{da} = \frac{n}{a} - (1) \sum_{i=1}^n \ln(x_i) = 0$$

 This simplifies to

$$\hat{a} = \frac{n}{\sum_{i=1}^n \ln(x_i)}$$

$M_2 = 1/n \sum_{i=1}^n X_i^2$ ,  $X_1^2 + X_2^2 + \dots + X_n^2$  that is  $A^2 + AB + B^2/3$  which is same as what we had and so we have 2 equations and 2 unknowns. The unknowns are A and B, the moments are  $m_1$  and  $m_2$ , those are not unknowns, okay. So, we can write  $m_1 = A + B/2$  and  $m_2 = A + B^2 - AB$ , where this can be written as  $A^2 + 2AB + B^2 - AB$  that would be  $A^2 + AB + B^2$ .

**(Refer Slide Time: 42:48)**

## EXAMPLE 7

Use the Method of Moments to find the parameter estimators of the following probability density function

$$f(x) = \frac{1}{B-A} \quad A \leq x \leq B$$

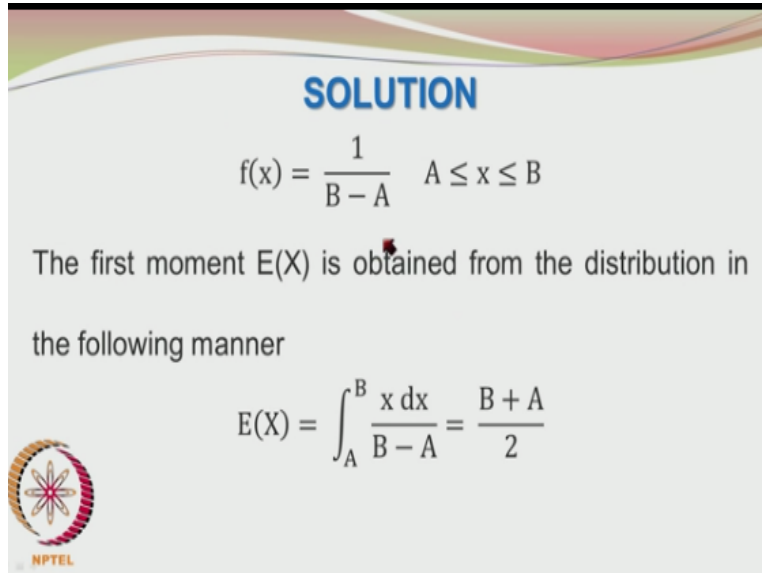
$$= 0 \text{ otherwise}$$

And when you have these and you solve for A and B, you get these 2 relations. I leave the quadratic equations solving to you, I hope you get the same answers as I did. So, thanks for your attention and we were doing some illustrative problems. There are lots of books on statistics and



probability which have many interesting problems. I requested you to not only solve these problems independently, but also look up the problems in various books.


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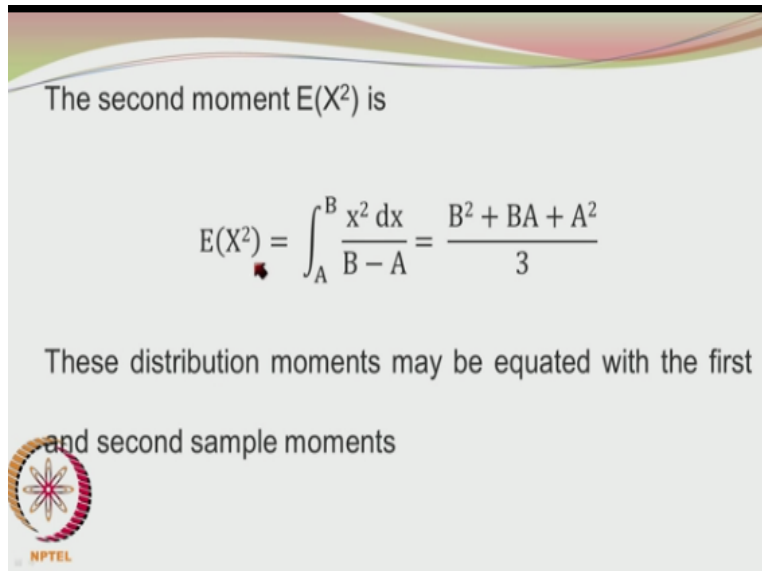
**SOLUTION**

$$f(x) = \frac{1}{B-A} \quad A \leq x \leq B$$

The first moment  $E(X)$  is obtained from the distribution in the following manner

$$E(X) = \int_A^B \frac{x \, dx}{B-A} = \frac{B+A}{2}$$



**(Refer Slide Time: 43:34)**



The second moment  $E(X^2)$  is

$$E(X^2) = \int_A^B \frac{x^2 \, dx}{B-A} = \frac{B^2 + BA + A^2}{3}$$

These distribution moments may be equated with the first and second sample moments




And try to solve them without any assistance either from these lectures or from the worked out examples in those book. Try to solve them on your own and if you are getting the correct answer well and good, nothing more has to be said, but if you are finding some difficulties and you are not able to get the correct answer, go through the lecture material again. See where exactly you have not understood correct your concepts.

**(Refer Slide Time: 44:43)**

## SOLUTION

These distribution moments may be equated with the first and second sample moments

$$m_1 = \frac{1}{n} \sum_{i=1}^n (X_1 + X_2 + \dots + X_n) = \frac{A + B}{2}$$


$$m_2 = \frac{1}{n} \sum_{i=1}^n (X_1^2 + X_2^2 + \dots + X_n^2) = \frac{A^2 + AB + B^2}{3}$$

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## SOLUTION

$$m_1 = \frac{A + B}{2}$$

$$m_2 = \frac{A^2 + AB + B^2}{3} = \frac{(A + B)^2 - AB}{3}$$

From these two equations we may solve for A and B to get

$$A = m_1 + \sqrt{3(m_2 - m_1^2)}$$

$$B = m_1 - \sqrt{3(m_2 - m_1^2)}$$

And then hopefully you will be able to work out these kinds of problems in the correct manner. The important thing is not the actual numerical solving, but the interpretation, the assumptions made and the concepts being applied with these kinds of problems. So thanks for your attention, will see you in the next class.