


**Image Signal Processing**  
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**Lecture 79**  
**Linear Estimator**

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Linear MS estimation

The linear MS estimate of  $Y$  in terms of r.v.s  $X_i$ ,  
 $i=1, 2, \dots, n$  is given by


$$\hat{Y} = \sum_{i=1}^n a_i X_i = a_1 X_1 + a_2 X_2 + \dots + a_n X_n$$

where  $a_i$ 's are all constants.

that must be determined such that  
 $E[(Y - \hat{Y})^2]$  is minimum

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We are looking at linear mean square estimator or estimation of a random variable  $y$  from a number of random variables  $x$ .  $N$  number of variables let us say. The linear mean square estimates of  $y$  in terms of random variables. Let us say  $x_i$ ,  $i$  equal to 1 to  $n$  is given by, now in this case we are, this is a constrained situation. We are constraining the estimator to be linear is given by, let us say  $y$  hat is equal to summation  $a_i x_i$ . In general, of course, you can also have a plus  $b$  in order to make it an assigned estimator, but then I will take a simpler case where we simply have, anyway  $b$  is also very straightforward to solve.

$i$  equal to 1 to  $n$  or in other words, this  $a_1 x_1$  plus  $a_2 x_2$  plus  $x_n$  where  $a_i$ 's are all constants which is why this is a linear estimator and the idea is that we would like to solve for  $a_1$   $a_2$  right up to an such that expectations so idea is constants that must be determined such that, we would like to find out these constants, such that the expectation of  $y$  minus  $y$  hat where  $y$  hat is this linear estimator,  $y$  is as small as possible, is minimum, under the constraint that  $y$  hat is linear. Correct? This is why we are looking at these constants.

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To determine the coeff  $a_1, a_2, \dots, a_n$

$$e = E \left[ (y - (a_1 x_1 + a_2 x_2 + \dots + a_n x_n))^2 \right]$$
$$\frac{\partial L}{\partial a_i} = E \left[ 2(y - (a_1 x_1 + a_2 x_2 + \dots + a_n x_n)) (-x_i) \right] = 0$$
$$E \left[ (y - (a_1 x_1 + a_2 x_2 + \dots + a_n x_n)) x_i \right] = 0, \quad i = 1, 2, \dots, n$$

Error in the estimate of  $y$   
The orthogonality principle for Linear

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What this means is that now, let us again so to determine the coefficients  $a_i$ 's  $a_1$  to  $a_n$ , let us define the error to be  $e$  is equal to expectation  $y$  minus  $\hat{y}$  is  $a_1 x_1$  plus  $a_2 x_2$  plus plus plus  $a_n x_n$ . The whole square. This we want to be minimum. Now suppose, we do a partial differentiation of  $e$  with respect to some  $a$ , let us say  $a_j$  or  $a_i$  let us say. Let us put it as  $a_i$ . Then what we find is, this will be equal to expectation.

Then we will have 2 times  $y$  minus  $a_1 x_1$  plus  $a_2 x_2$  plus plus plus all the way up to  $a_n x_n$  and the whole thing. If you were to examine this equation more carefully, this simplifies to expectation  $y$  minus  $a_1 x_1$  plus  $a_2 x_2$  plus plus plus  $a_n x_n$  into  $x_i$  is equal to 0.  $i$  equal to 1 to  $n$ . Now, what this means is, this is the error in the estimator  $y$ . Correct? The error in the estimator  $y$  because this is  $\hat{y}$ . So,  $y$  minus  $\hat{y}$  is the error, the estimator and what it is saying is this error is orthogonal to  $e$  to the observe random variables  $x_1$  to  $x_n$  and this is called the orthogonality principle for linear MS, mean square estimator.

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Special cases: If  $y$  is in the span of  $x_1, x_2, \dots, x_n$  then  $\hat{y} = y$

If  $y$  is orthogonal to the  $x_i$ , then  $\hat{y} = 0$

$$E[(y - (a_1x_1 + \dots + a_nx_n))x_i] = 0$$

$$E[(y - (a_1x_1 + \dots + a_nx_n))x_n] = 0$$

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So, in general, so this orthogonality principle, in a way, if you want to look at the geometric interpretation, what it would mean is that, if you are trying to, if you have your  $x_1$  to  $x_n$  like that exceed it all the way up to  $x_n$ . You have a  $y$ , that is sitting somewhere up there and what you are trying to do is, you are going to drop this orthogonally onto the space spanned by  $x_1$  to  $x_n$  and this is your  $\hat{y}$  and this error which you have here, this error is actually orthogonal to each of the  $x_i$ 's and some special cases, what will be the special cases that you might have here.

Special cases will be like if  $y$  is a vector in the span of  $x_1$  to  $x_n$ ,  $x_1, x_2$  up to  $x_n$ . Then your  $\hat{y}$  will be equal to  $y$  itself. You cannot make any error because of the fact that it lies in the span of the axis. Now if  $y$  is orthogonal on the other hand the other extreme, the  $y$  is orthogonal to the  $x_i$ 's then your  $\hat{y}$  is 0, if it is orthogonal. Now, in this case what you find is that through this principle, orthogonality principle says the error is orthogonal to all the  $x_i$ 's, each of the  $x_i$ 's and therefore this itself leads you to  $n$  number of equations if you observe.

This equation is valid for every  $i$  and therefore you can write down a set of equations. So we can go ahead and write this down as from the earlier page. What we have is expectation  $y$  minus  $a_1x_1$  plus plus plus  $a_nx_n$  into  $x_i$ . It is equal to 0 then all the way up to expectation  $y$  minus  $a_1x_1$  plus plus plus  $a_nx_n$ ,  $x_n$  is equal to 0 because the other equation is valid for  $i$  equal to 1 to  $n$ . Therefore, if you write down a bunch of equations, you will get  $n$  equations like that.

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$$E[a_1 x_1 x_i + a_2 x_2 x_i + \dots + a_n x_n x_i] = E[y x_i], \quad i=1, 2, \dots, n$$

Let  $r_{ji} = E[x_j x_i]$  and  $r_{0i} = E[y x_i]$

$$\begin{bmatrix} r_{11} & r_{12} & \dots & r_{1n} \\ r_{12} & r_{22} & \dots & r_{2n} \\ \vdots & \vdots & \ddots & \vdots \\ r_{1n} & r_{2n} & \dots & r_{nn} \end{bmatrix} \begin{bmatrix} a_1 \\ a_2 \\ \vdots \\ a_n \end{bmatrix} = \begin{bmatrix} r_{01} \\ r_{02} \\ \vdots \\ r_{0n} \end{bmatrix}$$

$$R \underline{a} = \underline{r}$$

$$\underline{a} = R^{-1} \underline{r}$$

*Handwritten notes:*  
 Find linear MS estimate of y given only x  
 $y = ax + b$   
 Solve for b and a  
 $E[(y - (ax+b))^2]$  is minimum

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Therefore, for each of these equations, it will get something like expectation  $a_1 x_1 x_i$  plus  $a_2 x_2 x_i$  plus plus plus  $a_n x_n x_i$  is equal to expectation  $y x_i$  where  $i$  goes from 1 to  $n$ . Now let us suppose  $r_{ji}$  is, suppose this indicates expectation  $x_j x_i$  and  $r_{0i}$  indicates expectation of  $y x_i$ . Then what we have is, we can write this entire thing in matrix vector form as where the unknowns are your  $a_1 a_2$  all the way up to  $a_n$ .

These are your unknowns and we can write this as a matrix multiplying a vector is giving you a vector and this will be like, so we start with  $i$  equal 1. So what you will have is  $r_{11}$  and then you have  $x_2$ . So now for  $i$  equal 1, you have got the what  $r_{11}$  which is expectation  $x_1$  square. Then you have got  $r_{21}$  which would be the expectation  $x_2 x_1$  and then all the way up to expectation  $x_n x_1$ . Therefore, you have got like  $r_{11} r_{21}$  all the way up to  $r_{n1}$ .

Then similarly, you can have  $r_{12}$  because when you put  $i$  equal to 2 then you will get  $x_1 x_2$ ,  $x_1 x_2$  is, so when you put  $x_1 x_2$  that is  $r_{12}$ .  $x_1 x_2$  is  $r_{12}$ . Then  $r_{22}$ , you got like  $r_{22}$  and then all the way down to  $r_{1n}$  because you going to put  $x_n$  here, if you are going to put  $x_1 x_n$  and you know  $x_1 x_n$  is like  $r_{1n}$  and then you have got  $r_{1n}$  and then you got  $r_{2n}$  all the way up to  $r_{nn}$ . This is equal to  $r_{01}$  because expectation  $y x_i$ . So  $y x_1$  is  $r_{01}$ ,  $y x_2$  is  $r_{02}$  and then all the way up to  $r_{0n}$  or in other words, if you call this as  $\underline{a}$  and if you call this is as  $\underline{r}$  matrix and then this we call as  $\underline{r}$ , then we got like  $\underline{r} \underline{a} = \underline{r}$  or in other words  $\underline{a}$  is nothing but  $\underline{r}^{-1} \underline{r}$ .

So, your  $a$  is which is your  $a_1$   $a_2$  right up to  $a_n$ , those can be found out from this kind of a statistical information. Try to kind of solve this and leave it as a small exercise. Show or find linear MS, mean square estimates of  $y$ , given only  $x$ . Now, what we have solved is a more general case, solved for the general case  $x_1$  to  $x_n$ , to solve it for the case, when let us say you have  $y$  and then you observe only another random variable which is  $x$  and so you need to be able to show us to what should be then your  $a$ .

So, in this case your  $a$  will be, so in general try to solve for  $y$  is equal to  $ax$  plus  $b$ . So, solve for  $a$  and  $b$ . Now in our case we chose  $b$  to be 0 but then let me just give you a simple example. Try to estimate  $y$  from  $x$  as  $ax$  plus  $b$ , which is like an affine estimator, purely linear due to  $y$  equal to  $x$ . Now, we are looking at some kind of an affine estimator so solve for  $b$  and  $a$  and  $a$  such that  $b$   $n$  such that expectation  $y$  minus  $ax$  plus  $b$  the whole square is minimum. So, it is a simpler example of the one which we have already seen. Now, this linear mean square estimator is actually nice to have.

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The General Orthogonality Principle for  
Non-Linear Estimator

If  $g(x)$  is a non-linear MS estimator of  $Y$   
 i.e. if  $g(x) = E[Y|x]$  then the estimation  
 error  $Y - g(x)$  is orthogonal to any function  $w(x)$   
 (linear or non-linear) of the data  $x$ .

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Now, we saw that it obeys the, what is called, orthogonality principle. We said that error is orthogonal to the observed random variables  $x_1$  to  $x_n$ . Moving forward, we are going to look at something called the general orthography principle and this is with respect to the conditional mean. The general orthogonality principle. See, as far as linear estimator is concerned, we said that the error that you make in the estimate of  $y$  is orthogonal through  $x_i$ 's. Now the general

orthogonality principle goes further beyond and states the following. This is for non-linear estimator that means the conditional mean.

What it says is if  $g$  of  $x$  is a non-linear mean square estimator of  $y$  that is if  $g$  of  $x$  equal to the conditional mean of  $y$  given  $x$  then, this estimation error which is  $y$  minus  $y$  hat or  $y$  minus  $g$  of  $x$  in this case, is orthogonal to any function. Note this, to any function. There we only had that it is orthogonal to  $x$ 's. We are saying to any function  $w$  linear or non-linear. Now, let me just remove the comma of the data  $x$ . Hence, so goes further beyond.

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Proof: Let  $g(x) = E[y|x]$   
 we need to show that  
 $E[(y-g(x))w(x)] = 0$

Examine  $E[g(x)w(x)] = E_x[w(x)E[y|x]]$   
 $= E[E[w(x)y|x]] = \int f_x(x) \left( \int w(x)y f_{y|x}(y|x) dy \right) dx$   
 $= \int \int y \cdot w(x) \cdot f_{y|x}(y|x) f_x(x) dx dy$

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So, we will try to prove this. What do we mean by a general orthogonality principle? So the idea is that so we know that proof would go as follows. The proof goes as follows. Let  $g$  of  $x$ , as has been already stated, is the conditional mean of  $y$  given  $x$ . Let us put it like that. So, we need to show, what do we have to show, we need to show that expectation  $y$  minus  $g$  of  $x$  into any guy, any this one  $w$  of  $x$ , any function of  $x$  is equal to 0, for any this one  $w$  of  $x$ . We do not want to really constrain  $w$  of  $x$ .

It could be any function of  $x$  linear or non-linear. Now examine expectation  $g$  of  $x$  into  $w$  of  $x$ . This we know is nothing but expectation of  $w$  of  $x$  into  $g$  of  $x$  is conditional mean of  $y$  given  $x$  and this we know is a function  $x$ . This one, the conditional mean of  $y$ , let me just write this, and this we know is simply a function of  $x$  and therefore this outer expectation is with respect to  $x$ .

So, this equal to, if not exclusively say that, we should automatically infer or we should automatically know that the outer expectation is with respect to  $x$ . This we can write because of the fact that the inner integral here is with respect to  $y$ , we can write this as expectation of expectation of  $w$  of  $x$  into  $y$  given  $x$  and this furthermore is nothing but integral  $f_x$  of  $x$  and then followed by an integral which is  $w$  of  $x$  into  $y$   $f_y$  given  $x$ ,  $dy$  and outside you will have  $dx$  or this in other words is, double integral  $y$  into  $w$  of  $x$  into  $f_y$  given  $x$   $f_x$  of  $x$   $dx$   $dy$ .

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$$= \int \int y \cdot w(x) \cdot f_{x,y}(x,y) dx dy$$

$$= E[w(x) | y] \quad \text{Hence proved.}$$

This is equal to, so if you see observe here, this we can write as  $y$  double integral  $y$  into  $w$  of  $x$  into joint because you are multiplying, if  $y$  given  $x$  with  $f_{x,y}$   $dx$   $dy$  or this in turn is expectation of  $w$  of  $x$  into  $y$  and hence we have the proof which is equal to this  $y$  into  $w$  of  $x$ . So, we wanted to show that this is equal to  $E[w(x) | y]$  that is equal to showing that expectation  $w$  of  $x$  into  $y$  is equal to expectation  $y$  into  $w$  of  $x$ , which is what we have proved here. So hence proved.

Now, at this point, what we would like to do next really is what is called a normality theorem, what is called the normality theorem and this theorem is the one that actually, you know, reveals the equivalence between a linear estimator and a non-linear estimator, the conditions under which the conditional mean becomes linear. Typically, we know that it is a nonlinear function of  $x$  but then under certain condition, it becomes a linear function of  $x$ , which means that the linear estimator is indeed the optimal one. It is no longer sub-optimal. So, we so we will see next, what is the normality theorem.