

Stochastic Modeling and the Theory of Queues
Prof. Krishna Jagannathan
Department of Electrical Engineering
Indian Institute of Technology – Madras

Lecture –58

Stationary Distribution of a Countable State Space DTMC and Renewal Theory

(Refer Slide Time: 00:14)

Stationary Distribution & Renewal Theory

PMF $\{\pi_i, i \in S\}$ $\sum_{i \in S} \pi_i = 1$

such that $\pi_i = \sum_{j \in S} \pi_j P_{ji}$ } (*)

Then we can show that if $P(X_0=i) = \pi_i$, then $P(X_n=i) = \pi_i \forall i \in S, n \geq 1$.

Such a PMF $\{\pi_i, i \in S\}$ is known as a stationary distribution.

Theorem 1 Consider an irreducible countable state DTMC, provided $\{\pi_i, i \in S\}$ satisfies (*). Then

(i) the chain is +ve recurrent.

Welcome back. So, far we have look we have been looking at the connections between renewable theory and DTMC's. In particular we have said that consecutive returns to a particular state recurrent state j constitute renewal instances. Today we are going to use renewal theory renewal reward theory in particular to discover the stationary distribution of accountable state DTMC.

Recall that we say that let us say PMF π_i is given to you. So, this sum over π_i equals 1 such that it satisfies π_i equals sum over $\pi_j P_{ji}$. Suppose a PMF on the states satisfies π_i is equal to sum over j by $j P_{ji}$ then we can show that you can show that if the chain is started in the distribution π_i then for all i belongs to S then probability that X_n equals i will also be equal to π_i for all i belongs to S and for all n greater than or equal to 1 what does this mean?

We already saw this in the finite state context. So, essentially we are saying that π_i is equal to $\pi_i P$ except that I am not writing it as π_i is equal to $\pi_i P$ because the matrix P is not really a matrix now it is an infinite dimensional object but the fact remains the same that if you start

off the Markov chain at time equal to 0 in state i with probability π_i then the probability of being in state i for any time n remains π_i for every state i .

This is easy this is just by applying just applying total probability and Markov property. Now the thing is we are going to prove that this stationary distribution if is such a such a π_i PMF π_i is known as a stationary distribution. This we already seen of the DTMC such a PMF such a distribution is known as the stationary distribution of the DTMC. Now we are going to relate this π_i if at all there exists such a π_i we are going to relate it to recurrence times mean recurrence times of state i using renewal theory renewal regard theory.

So, this is the key result is the following theorem. So, theorem one consider an irreducible countable state DTMC for which π_i satisfies the star. Then let me write the following then number one the chain is positive recurrent.

(Refer Slide Time: 04:56)

(ii) The solution $\{\pi_i, i \in S\}$ is a unique solution to (*) & further $\pi_i = \frac{1}{E[T_i]}$ for all $i \in S$.

Theorem 2 Suppose an irreducible countable state DTMC is positive recurrent then (*) has a solution with $\pi_i = \frac{1}{E[T_i]}$ for all $i \in S$.

Pf of Thm 2 Suppose we start the chain @ time 0 in distribution $\{\pi_i\}$

Let $N_j(t)$ be the counting process of state j given we start in dist. $\{\pi_i\}$

The solution π_i is a unique solution to star and further π_i is equal to 1 over expected T_{ii} for all i in S . So, this theorem is a very important theorem suppose. So, it says that suppose you manage to find a solution π_i suppose if somehow by guess work or asking somebody or whatever means you manage to find a solution a distribution π_i such that some over π_i is equal to 1 and some over j $\pi_j P_{ji}$ is equal to π_i .

So, these are the balance equations this is what we call the global balance equations. So, if you manage to find a solution to the global balance equation by whatever means then we can come to some important conclusions the conclusions are that the solution you manage to find

is unique and the chain is positive recurrent. And furthermore the π_i you found out are the reciprocals of the mean recurrence time of each of the states.

So, there are three conclusions in this theorem. So, you somehow managed to find a π_i . So this theorem says that a this chain is positive recurrent P . The solution you manage to find somehow is the only solution there cannot be any other solutions to π_i is equal to $\pi_i P$ and this π_i has the interpretation as the reciprocal of the mean recurrence time of that particular state. So, this is a very important theorem.

Let me also state theorem two which is like the converse of the theorem one. So, these two theorems are central and then we will see how these theorems are proved using renewal reward theory. Suppose an irreducible countable state DTMC is positive recurrent then star has a solution with π_i is equal to 1 over expected T_{ii} for all i in S . So, this is like the converse theorem two is like the converse theorem one says that if you somehow manage to find a solution to this π_i is equal to $\pi_i P$ to these balance equations.

Then you get positive recurrence for free and the solution is unique and π_i will be equal to 1 over T_{ii} bar when expected T_{ii} for each state. Now we are saying the opposite. So, now I do not know π_i but I know that the chain is positive recurrent. So, what I know is that. So, it is an irreducible DTMC it is positive recurrent. So, all states are positive recurrent. So, 1 over expected T_{ii} a bar is something non-negative.

Now it is actually strictly positive if you think about it because expected T_{ii} is finite because it is positive recurrent. So, 1 over expected t_i is some strictly positive number we are saying that 1 over expected T_{ii} is the solution to π_i is equal to $\pi_i P$ that is what theorem two is saying. So, these two theorems can be viewed as I mean theorem two is a converse to theorem.

So positive recurrence implies there exists the this 1 over T_{ii} bar the solution to π_i is equal to $\pi_i P$ and theorem 1 says if π_i is equal to $\pi_i P$ has a solution π_i then its unique positive recurrence follows and this π_i has the interpretation of 1 over expected T_{ii} . So, bottom line is that the stationary distribution π_i for any positive record Markov chain is simply 1 over expected mean recurrence time.

We can prove all this using renewal reward proof of theorem one suppose we start the chain at time zero in distribution π let $N_j(t)$ be the counting process of state j given we start in distribution π $N_j(t)$ simply a counting process.

(Refer Slide Time: 11:23)

The slide contains the following content:

- NPTEL logo
- Equation: $P(X_n = j) = \pi_j \quad \forall n \geq 1, \forall j \in S.$
- Derivation: Consider $\frac{E[N_j(t)]}{t} = \frac{1}{t} E \left[\sum_{n=1}^t I(X_n = j) \right] = \frac{1}{t} \sum_{n=1}^t P(X_n = j) = \pi_j \quad \forall t \geq 1$
- Text: Condition on starting state i ,
- Equation: $E[N_j(t)] = \sum_{i \in S} \pi_i E[N_{ij}(t)]$
- Boxed result: $\Rightarrow \frac{1}{t} E[N_j(t)] = \frac{1}{t} \sum_{i \in S} \pi_i E[N_{ij}(t)] \quad \forall t \geq 1$

So, what am I doing I am starting in distribution π at time equal to zero what does that mean. So, at time equal to 0 the probability that I am in state i is just π_i and then I let the chain run and I am going to count the number of times state j occurs until time t and I am going to call that $N_j(t)$. Now please note that probability that X_n is equal to j is equal to π_j for all n greater than or equal to 1 and for all j and S .

So, if you look at this guy now consider expected $N_j(t)$ over t that is just one over t sum over n is equal to 1 to t probability that X_n equals j . So, you can look at this as you know maybe I should write one over one more step maybe I should write it as expectation of the indicator what is the what is $N_j(t)$? $N_j(t)$ is simply the indicator that X_n is equal to j summed over n is equal to 1 to t .

Now I can take the expectation inside because inside the summation because this is a finite sum. So, and if I take the expectation inside the summation I have the expectation of an indicator and the expectation of indicator is simply probability. So, this becomes 1 over t sum over n is equal to 1 to t probability that X_n equals j but of course probability that X_n equals j is simply π_j . So, this is π_j correct.

So this is true for all t correct. Now you can condition on starting state i and use total expectation to write expected n_j tilde t is simply equal to sum over π_i sum over $i \in S$ expected N_{ij} of t . So, why is this? So, N_j tilde t is the number of occurrences of state j until time t but you could have started in any of the states i and the probability of starting in state i is by assumption π_i am starting in the stationary distribution.

So, this is my total probability that the total expectations the law of total expectations. So, this implies what. So, this implies an important equation. So, one over t this implies expected N_j tilde t is equal to $1/t$ sum over $i \in S$ π_i expected N_{ij} of t for all t not equal to 1. So, this is an important equation let me call this dagger.

(Refer Slide Time: 15:27)

The slide contains the following content:

- NPTEL logo in the top left corner.
- A handwritten equation:
$$\Rightarrow \left[\mathbb{E}[N_j(t)] = \frac{1}{t} \sum_{i \in S} \pi_i \mathbb{E}[N_{ij}(t)] \right] \quad \forall t \geq 0 \quad \dagger$$
- A diagram showing a horizontal timeline from state i to state j at time t . A double-headed arrow labeled T_{ij} indicates the first passage time from i to j . A single-headed arrow labeled $t - t_i$ indicates the time interval from t_i to t .
- Two handwritten equations below the diagram:

$$\mathbb{E}[N_j(t) | T_{ij} = t_i] = 1 + \mathbb{E}[N_j(t - t_i)] \leq 1 + \mathbb{E}[N_j(t)] \quad \forall t \geq 0$$

$$\mathbb{E}[N_j(t)] \leq 1 + \mathbb{E}[N_j(t)] \quad \forall t \geq 0$$

Now I have to bound. So, but I know that. So, the limit t tending to infinity of the left hand side is equal to π_j . So, that also I know. So, but let me let me do the following ah. So, this i this i know to be equal to π_j but I will use that later maybe I should just write it down. So, this implies π_j is equal to $1/t$ sum over $i \in S$ π_i expected N_{ij} of t actually let me call this is true for all t greater than or equal to 1.

So, this let me call this dagger this is an important equation now looking at expected N_{ij} of t . So, if you are starting at i and then you are looking at the subsequent occurrences of j that is your N_{ij} process. Now if you are looking at some time t you can write expected. So, let me call this time as T_{ij} . So, it is the first time that the first passage time from i to j and then I have I am looking at a delay renewal process basically of returns to j .

Now this expected N_{ij} of t can be written as. So, let me let me condition on expected n_j of t given T_{ij} equal to some $t-1$. So, I am taking this time to be equal to some $t-1$. This will be equal to 1 plus expected. So, you will have 1 occurrence of j and after which you are looking at n_j of t minus $t-1$. So, conditioning on T_{ij} is equal to $t-1$ and then so, n_{ij} of t will simply be $1 + N_{jj}$ of $t-1$ correct.

So, this is $t-1$ that is $t-1$ and this is less than or equal to 1 plus expectation of N_j of t . So, this is true for all $t-1$ and in the hand side. So, this is two for all $t-1$ and in the hand side there is no $t-1$. So, this seems to be true for all $t-1$. So, I can write expected i can just remove the conditioning and write expected N_{ij} of t is less than or equal to $1 +$ expected N_{jj} of t for all and this is true even if this $t-1$ had to be infinite this is still true you can easily verify that this equation will still trivially hold.

(Refer Slide Time: 19:55)

The slide contains the following content:

- NPTEL logo
- Equation 1:
$$\pi_j = \frac{1}{t} \sum_{i \in S} \pi_i \mathbb{E}[N_{ij}(t)] \leq \frac{1}{t} \sum_{i \in S} \pi_i (1 + \mathbb{E}[N_{jj}(t)]) \quad \forall t \geq 1$$
- Equation 2:
$$t\pi_j \leq \sum_{i \in S} \pi_i (1 + \mathbb{E}[N_{jj}(t)]) = 1 + \mathbb{E}[N_{jj}(t)]$$
- Equation 3 (boxed):
$$\mathbb{E}[N_{ij}(t)] \geq t\pi_j - 1$$
- Text: "Since $\sum \pi_i = 1$, $\pi_j > 0$ for some j . \Rightarrow for such j ,
- Equation 4:
$$\lim_{t \rightarrow \infty} \mathbb{E}[N_{ij}(t)] = +\infty \Rightarrow j \text{ is recurrent.}$$

So, we have this equation which is nice. So, we can put this back into our dagger. So, then what happens. So then I have sub back in this dagger and then what happens I will get π_j equal to 1 over t sum over i belongs to S π_i expected N_{ij} of t which is less than or equal to 1 over t sum over i belongs to S π_i 1 plus expected n_{jj} of t this is true for all t and now what does this mean I have $t\pi_j$ is less than or equal to sum over i belongs to S π_i 1 plus expectation of yeah that is expectation of N_{jj} of t .

So, this is π_i sorry. So, this is the mistake I made I think sorry I made this mistake my apologies. So, I made this mistake. So, here this also should be π_i . So, now if you look at this guy this term is has no i in it. So, this is just sum over π_i which is one because I know

that π_j is a distribution. So, this becomes equal to 1 plus expectation of N_{jj} of t which means that i can write expectation of N_{jj} of t is greater than or equal to $t \pi_j - 1$.

This is true for all states j . Now since sum over π_i is equal to 1 π_j is strictly positive for some j otherwise they cannot sum to 1. So this implies for such j expected N_{jj} t limit t tending to infinity is greater than or equal to $t \pi_j - 1$ and π_j is strictly positive. So, this becomes infinity plus infinity and this implies from our earlier theorem if this limit of N_{jj} t is going up to infinity then this process is a legitimate renewal process the place that state j is recurrent correct and if you go back to the above equation you get $\sum \pi_j$.

(Refer Slide Time: 23:56)

NPTEL

$$0 < \pi_j \leq \frac{1}{t} + \frac{E[N_{jj}(t)]}{t}$$

$$\Rightarrow \lim_{t \rightarrow \infty} \frac{E[N_{jj}(t)]}{t} > 0$$

$$\parallel$$

$$\frac{1}{E[T_{jj}]} \Rightarrow E[T_{jj}] < \infty \Rightarrow j \text{ is recurrent.}$$

The DTMC is recurrent

$$\frac{1}{E[T_{jj}]} = \lim_{t \rightarrow \infty} \frac{E[N_{jj}(t)]}{t} \geq \pi_j.$$

Back to \dagger
$$\pi_j = \frac{1}{t} \sum_i \pi_i E[N_{ij}(t)]$$

So, now you have you proven that j is recurrent you have to prove that j is positive or current now if you go back to this dagger you have π_j is less than or equal to $1/t$ plus expected N_{jj} of t this is just an upper bound on the dagger equation this is this i know to be true and this guy is strictly positive. So, this implies limit t tending to infinity expected N_{jj} of t over t is strictly positive.

Now N_{jj} t we already proved that N_{jj} t is a renewal process that is because state j is recurrent we already proved the state j is recurrent out here. So, N_{jj} t is a renewal process. Now we are saying that limit t tending to infinity expected N_{jj} t by t is strictly positive but by elementary renewal theorem this is equal to the reciprocal of T_{jj} . So, this implies that expected T_{jj} is finite implies j is positive recurrent. So, not only is j recurrent j is also positive recurrent.

So, in fact we can prove that $\frac{1}{T_{jj}}$ is expected $\frac{1}{T_{jj}}$ which is just limit t tending to infinity expected N_{jj} of t over t which is by elementary inverse theorem is greater than or equal to π_j which I know is strictly positive in fact this greater than or equal to we will show holds with an equality that will be our next result in proving this theorem. So, just to recap, so, I have proved that j is i follow the sequence of step first prove that j is recurrent by showing that expected N_{jj} goes off to infinity as t goes to infinity.

Then I prove that limit t tending to infinity expected N_{jj} by t is positively positive which means j is positive recurrent. Now the entire chain must be positive recurrent because you have one irreducible class. So, you know that all the states in the same class must be of the same type. So, if j is positive or current then the entire class must be positive recurrent. So, the DTMC itself is positive or correct.

So, from this we can prove that the DTMC is positive recurrent. Since the DTMC is reducible. So, we got that the DTMC is positive current and we got this lower bound on expect one over expected T_{jj} is greater than or equal to π_j now we will show that there is also an upper bound. So, that π_j will be equal to $\frac{1}{\text{expected } T_{jj}}$. Now you go back to dagger which says that π_j is equal to sum over.

So, $\frac{1}{t} \sum_{i=1}^m \pi_i$ sorry I keep making this mistake π_i expected N_{ij} of t have written it correctly correct.

(Refer Slide Time: 28:24)

Fix some state m

$$\pi_j \geq \sum_{i=1}^m \pi_i \frac{E[N_{ij}(t)]}{t} \quad \forall m \geq 1$$

Take $\lim_{t \rightarrow \infty}$

$$\pi_j \geq \sum_{i=1}^m \pi_i \lim_{t \rightarrow \infty} \frac{E[N_{ij}(t)]}{t} = \frac{1}{E[T_{jj}]}$$

π_i 's are unique!

$$\pi_j \geq \frac{1}{E[T_{jj}]} \quad \forall m \geq 1$$

$$\boxed{\pi_j \geq \frac{1}{E[T_{jj}]} - (2)}$$

$$\textcircled{1} \& \textcircled{2} \Rightarrow \boxed{\pi_j = \frac{1}{E[T_{jj}]}}$$

Now I am going to just go from fix some state m and then I am going to go the summation I am going to go to only till m . So, I have π_j is greater than or equal to $\frac{1}{t} \sum_{i=1}^m \pi_i$ expected N_{jj} of t this is true for all m greater than or equal to 1. So, I am just fixing some. So, I want to take some limit inside. So, I am fixing some finite sum. The reason I am doing this is, so, that I can take.

So, take limit and turning to infinity. So, I can simply just take t inside. So, that is perfectly legitimate. Now take limit t tending to infinity then I will get π_j is greater than or equal to $\sum_{i=1}^m \pi_i$ equals $\frac{1}{m}$. See I am taking the limit inside this finite sum which is why I took only finite m . So, π_j limit t tending to infinity N_{ij} of t over t expectation expected N_{ij} of t over t . So, now we know that the state j is positive recurrent.

So, this guy this limit by elementary renewal theorem for delayed renewal processes this will be equal to $\frac{1}{\text{expected } T_{jj}}$. So, this is equal to. So, therefore π_j is greater than or equal to $\frac{1}{\text{expected } T_{jj}}$ times $\sum_{i=1}^m \pi_i$ maybe I should say again I made the mistake. So, I have to write i here I keep making this mistake sorry $\sum_{i=1}^m \pi_i$ this is true for all m greater than or equal to 1.

Since this is true for all m greater than or equal to 1 and so, I can take m as large as I want. So, and then this guy will get as close to 1 as I want. So, since this is true for all m greater than or equal to 1 I have that π_j is greater than or equal to $\frac{1}{\text{expected } T_{jj}}$ because if π_j were less than $\frac{1}{\text{expected } T_{jj}}$ I should be able to take little m large enough. So, that the scenic with this inequality above is not true.

So, I got this bound. So, I have this guy let me call this let me call this 2 and I have already upper bound is not it where do I have that this I have 1. So, I have $\frac{1}{\text{expected } T_{jj}}$ is greater than or equal to π_j and here I have $\frac{1}{x^2 T_{jj}}$ less than or equal to π_j . So, 1 and 2 imply that π_j is equal to $\frac{1}{\text{expected } T_{jj}}$ and this is true for all states j there is nothing special about this j that I fixed.

So, what have I proved? If I proved that if there is a PMF π_i if there is a distribution π_i which satisfies the balance equations $\pi_i = \pi_i P$ and $\sum \pi_i = 1$ then that π_i must be equal to $\frac{1}{\text{expected } T_{jj}}$ therefore π_j 's are unique. So, whenever there

is a solution π_j that π_j must be equal to 1 over expected T_{jj} and the chain is positive recurrent. So, if you go back to theorem one I have proved theorem one.

So, I have proved this I have proved uniqueness and I have proved positive recurrence. So, I am done proving theorem 1.