

Stochastic Modeling and the Theory of Queues
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Lecture –60

Stationary Distribution and the Steady State Behaviour of a Countable-state DTMC - P1

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The slide content is as follows:

Stationary Distribution & Steady State
(A direct approach without Renewal Theory)

Ref: Grimmett & Stirzaker (Ed 3) - Section 6.4

Given a DTMC irreducible, recurrent

(i) Existence of a stationary "measure"
+ve recurrence - stationary distribution

(ii) If the DTMC is also aperiodic & (+ve recurrent) gives
Want to show convergence to steady-state: $\lim_{n \rightarrow \infty} P_{ij}^{(n)} = \pi_j \cdot \mathbb{1}_{j \in E}$

Welcome back till the last module we discussed stationary distribution and steady state behaviour of countable state Markov chains using renewal theory. So, we used renewal reward theory to interpret the steady state probability with π_i as 1 over T_i which is the expected recurrence time. We also use Blackwell's theorem to prove that $P_{ij}^{(n)}$ converges to π_j if you have an irreducible a periodic positive recurrent Markov chain. Now all this is perfectly valid what we have done.

So, far is perfectly valid however it makes heavy use of renewal theory. Now one might argue that is renewal theory really required if it is valid to use it we have studied renewal theory but you may argue that hey so, this I do not know any renewal theory suppose can you help me learn Markov chains. It turns out that you can derive some of these results some of the key results about stationary distribution and steady state probability and all that without using renewable theory at all.

We can just use a direct elementary approach and derive some of these fundamental results about stationary distribution and steady state behaviour of Markov chains. So, that is


something I want to do today not in great detail. But I will just point out how the arguments go this part is from the book by Gramat and Sterzeker's book on probability and random processes third edition section 6.4 which is this is a lovely book.

And this section is a good exposition on the stationary distribution and renewal without renewal theory all. So, my agenda for today is to roughly tell you how the stationary distribution can be constructed explicitly without using any renewal theory and how the convergence to steady state can be proved. So, I am given let us say given DTMC which is irreducible and recurrent.

I want to prove that I want to prove the existence of a stationary distribution I want to prove the existence of a stationary measure. And in the case of positive recurrence the stationary measure will be a stationary distribution I will tell you what I mean by a stationary measure it is a counting measure that we will introduce. And also for an aperiodic if the DTMC is also aperiodic. So what I mean is that it is irreducible recurrent and aperiodic positive recurrent any periodic.

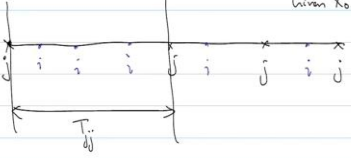
Then I want to show that there is a convergence to steady state that is what i want to do is I want to show that $\lim_{n \rightarrow \infty} P_{ij}^{(n)}$ is equal to π_j all this is for all j in S for all i in S and for all j in S . So, this is convergence to steady state. So, remember the for the first for the existence and the uniqueness of stationary measure and all that we have used for the stationary distribution we used renewable reward and for the second we use Blackwell's theorem. I want to do it without using any of these renewable theory arguments.

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Recurrent, irreducible DTMC

Fix $j \in S$ recurrent.
Given $X_0 = j$



$R_i = \# \text{ of visits to state } i \text{ before return to } j \text{ given } X_0 = j$

$\rho_i(j) \triangleq \mathbb{E}[R_i | X_0 = j] \quad i \in S$

Suppose so, you are given a DTMC are given a recurrent irreducible DTMC. So, I am going to say let us say I fix j recurrent. So, I am just going to fix a recurrent state j . So that is my Markov chain running in time let us say I started j let us say given $X_0 = j$ let us say I have a j which is a recurrent state and I start the Markov chain at state j then this is a recurrent state. So, I have to return to state j at some point and I will in fact with probability 1 I will keep returning to state j infinitely many times.

Earlier said that we said that these subsequent consecutive return times constitute a renewal process but I do not want to. I am going to pretend that I know nothing about renewal theory at all in this lecture. So, what I am going to do is you take any other state i let me mark it differently. So, let me say that I am hitting state i in these places this at these times I am going to say R_i .

So, I am going to just consider this interval between two consecutive visits to j . So, I started j and I look at the first return to j and I used to call this T_{jj} . So, I will keep the notation but I will not use any renewal terminology I am going to define R_i as the number of visits to state i . So, state i is in the same class, state i is in the same class as j . Number of visits to state i before return to j given $X_0 = j$.

So, this is some this is some number. So, earlier we used to say R_i is some kind of a random variable we said this is some expected reward inside a renewal interval and all that but forget all that it is just a random variable at this point and I am going to say I am going to call $\rho_i(j)$

of j defined as expected R_i which is the expected number of visits to i given that i started at j . So, it is the expected number of visits to i before I come to j given that i started with j .

So, this I will define for I am fixing a state j and this is defined for the other states S including state j . Now I am going to use this ρ_i of j as some kind of a it is some kind of a counting measure. So, I am going to argue that there is a stationary measure can be con stationary measure and eventually a stationary distribution can be constructed using these ρ_{ij} 's.

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The slide contains the following mathematical derivations:

$$R_i = \sum_{n=1}^{T_{jj}} \mathbb{I}_{\{X_n=i\}} = \sum_{n=1}^{\infty} \mathbb{I}_{\{X_n=i\} \cap \{n \leq T_{jj}\}}$$

$$\rho_i(j) = \mathbb{E} \left[\sum_{n=1}^{\infty} \mathbb{I}_{\{X_n=i\} \cap \{n \leq T_{jj}\}} \mid X_0=j \right]$$

$$\rho_i(j) = \sum_{n=1}^{\infty} \mathbb{P}(X_n=i, n \leq T_{jj} \mid X_0=j)$$

Define $\underline{\rho}(j) = [\rho_1(j), \rho_2(j), \dots]$

So, to be little more precise maybe I should define R_i as sum over n is equal to 1 to T_{jj} indicator $x_n=i$ maybe I should say n is equal to 1 to T_{jj} indicator x_n is equal to i . So, it is the number of times I am in state i before this before this stopping time T_{jj} or I can also write this as sum over n equals 1 to infinity indicator x_n equals i intersection n less than or equal to T_{jj} this is my expected this is R_i .

I can write then ρ_{ij} as the expectation of all this expected that given $x_0=j$ which is just sum over n equals 1 to infinity expectation of sum over n equals 1 to infinity indicator $x_n=i$ indicator sorry intersection n less than or equal to T_{jj} . This I can take the expectation inside because all the terms are non-negative. So, this just becomes n equals 1 to infinity probability that $x_n=i$ and n less than or equal to T_{jj} given $x_0=j$.

So, this is my ρ_{ij} let me go ρ_i of j . So, now my claim is that this ρ_i of j which is the expected number of visits to state i between two successive returns to j is some kind of a

stationary measure in the sense that it satisfies the balance equations see this ρ_{ij} will turn out to be a number I am going to argue that this is a stationary measure. So let me also define the vector ρ vector ρ of j as ρ_1 of j ρ_2 of j and all that over all the states.

So, this is a vector which contains all the ρ_i of j as you run over i . So, I am going to make a few claims.

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Claims

(i) $\rho_j(j) = 1$

(ii) $0 < \rho_j(j) < \infty$

(iii) $\sum_{i \in S} \rho_i(j) = E[T_{jj}] \leftarrow \text{True even if } R_{jj} \text{ is infinite.}$

(iv) $\rho_j(j)$ is a "stationary measure" i.e.,

$\rho_j(j) P = \rho_j(j)$ i.e. $\rho_j(j) = \sum_{k \in S} \rho_k(j) P_{kj} \quad \forall i \in S$

Now claims first claim is that row j of j equals 1. So, this is the expected number of visits to state j between the two visits to j and one of the final return to j you count the last return. So, in fact you he you return to j when you turn to j . So, that is the only time you count this particular reward. So, if you put I am saying reward but it is not I mean I am not really in the I mean in the renewable world we used to think of it as the reward.

You used to you basically collect one unit of reward when you return to j but it is here we are just counting and that is. So, we get one expected unit of reward. So, if you look at in fact r_j will be equal to 1 in this the number of times you visit j between successive returns to j . So, that will be equal to 1 this is claim 1. Claim 2 is that these ρ_{ij} 's ρ_i of j s are strictly between zero and infinity.

Claim 3 is that if you sum over all the states is ρ_i of j you get expected T_{jj} . So, the expected time between successive returns to j remember that j is recurrent. So, you do have T_{jj} is a random variable what we are saying is that the expected T_{jj} the expectation of this random variable is equal to sum over all the states i of ρ_{ij} and this is true even if RHS is

infinite of course if T_{jj} is a random variable expected T_{jj} could be finite or infinite expected T_{jj} is finite then we know that the state the chain is positive recurrent expected data is infinite the chain is null recurrent.


But fact is that if you sum over the ρ_{ij} 's you get expected T_{jj} this is true for both positive recurrent and different chains. And most important of all this vector ρ , ρ of j is a stationary measure ie what we can show is that $\rho_j P$ equals ρ_j whereas I am writing this πP is equal to π kind of a thing except this π I am writing instead of π I am writing row j and this row j is not a probability distribution.

It is a stationary measure in the sense it is stationary you put ρ_j into p you get another you get another ρ_j but it is not a distribution because it does not have to sum to one it does not sum to one in general it does not come to one at all actually. Now so, this is a shorthand for saying that maybe I should write this see this P need not be a matrix P is infinite dimensional object here. So, I should really write ρ_i of j is equal to sum over k ρ_k of $j P_{kj}$.

Sorry P_{kj} I am sorry P_{ki} ρ_k of $j P_{ki}$ this is true for any this is true for this stationary measure. Let us say this ρ_j vector this is true for all i in S . So, what I am claiming is that this ρ_{ij} which i defined as the expected number of visits to state i between two successive which is to j . This ρ_{ij} satisfies ρ_{jj} equal to 1 which is reasonably clear because you have you basically return to j and that is the only time you count that.

And the second claim is saying that this ρ_{ij} is strictly positive and strictly finite and third is saying that the sum over all i ρ_{ij} is equal to expected time of return between j to j expected T_{jj} and finally it is the most important it is saying that this ρ_j vector satisfies ρ_j of P ρ_j is equal to ρ_j . We have to prove this I have indicate you how this is proved.

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Remark: If the DTMC is +ve recurrent, i.e., $E[T_{ii}] < \infty$,
 then $\frac{\rho_i(j)}{E[T_{ii}]}$ can be interpreted as the stationary dist of i .

$$\pi_i = \frac{\rho_i(j)}{E[T_{ii}]} \quad i \in S \quad \text{clearly } \sum \pi_i = 1 \quad (\text{from (iii)})$$

$$\text{and } \pi P = \pi \quad (\text{from (iv)})$$

$$\pi_i = \frac{\text{Expected \# visits to } i \text{ when starting returns to } j}{\text{Expected recurrence time of } j}$$

So, now let us look at what the consequences if the DTMC is positive recurrent i.e. if expected T_{jj} is finite then if you take $\rho_i(j)$ over expected T_{jj} can be interpreted as the stationary distribution of i . So, that what I am saying is that if you call π_i as $\rho_i(j)$ over expected T_{jj} . So, I am saying. So, this is finite the denominator is finite. So, this π_i will turn out to be something strictly positive because $\rho_j(j)$ is strictly non-zero and expected T_{jj} because i have positive recurrence is finite.

So, ρ_{ij} over expected T_{jj} will be some positive number now this if you take this π_i and you calculated for all i and S then clearly sum over π_i is equal to one that is because I am normalizing by expected T_{jj} and that is because sum over ρ_{ij} over x sum over ρ_{ij} is equal to expected T_{jj} and I am normalizing by expected T_{jj} . So, sum over equal to sum over π_i equal to one this is by due to 2 sorry due to 3 which we have not proven and πP is equal to π from 4, property 4.

So, further if the DTMC is positive recurrent I can take this ρ_{ij} and divide it by expected T_{jj} and interpret this as π_i call this π_i and this π_i will be a distribution over all the states in the sense that it sums to 1 and πP will be equal to π it will be a stationary distribution of course if the state if the Markov chain is null recurrent you have the stationary measure ρ_{ij} if it still satisfies ρ_{ij} is equal to ρ_j .

But you know if you divide it by expected T_{jj} you get all zeros. So, it does not sum to one. So, this π_i will all be 0 and it is not really a distribution but so, even for a null recurrent chain everything I have said before about ρ_{ij} still holds except that you do not get a

probability distribution over the states in a null recurrent case. In a positive recurrent case you do get a distribution over the states.

So, what we are saying is that this is the steady state. So, π_i is the expected number of visits. So the steady state probability of steady state distribution of state i has the interpretation as expected number of visits to i between successive returns to j divided by expected say expected recurrence time of j which we are taking to be finite. So this we already knew from Renewal Reward theory.

In fact we considered a reward process which gives a reward of 1 when you visit i but the renewal process we considered was between with us to j and we used renewal reward theory to come up with this. So, this is not surprising to us at all because we have seen this using Renewal theory. But all I am saying is that you can explicitly construct this π_i like. So, and from elementary principles not using any renewal theory or anything we can use we can prove these claims one two three and four which I have not done which I will indicate how to do.

So, what I am saying is that you define the ρ_{ij} and if you are expected T_{jj} is finite you divide ρ_{ij} by expected T_{jj} . And you call it that turns out to be π_i which is your stationary distribution which we already knew from because we know renewal theory but we are constructing from first principles now.