

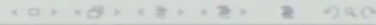
Stochastic Processes

Module 8: Renewal Processes

Lecture 5: Non Markovian Queues

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Video Course on
Stochastic Processes -1

By

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Lecture# 5
Non Markovian Queues

Outline

$GI/M/1$ Queue

$GI/M/c$ Queue

$GI/M/1/N$ Queue

$GI/G/1$ Queue

References



This is stochastic processes module 8 renewal processes. In the lecture 1 we have discussed the renewal function and renewal equation. In the lecture 2 we have discussed generalized renewal processes and renewal limit theorems. In lecture 3 we have covered Markov renewal and regenerative processes. In lecture 4, we have discussed non Markovian queues such as MG1 queue, MG1N queue, MGCC loss systems. This is lecture 5, non-Markovian queues. In this lecture we are going to discuss $GI/M/1$ Queue, $GI/M/c$ Queue, then $GI/M/1/N$ Queue, and finally $GI/G/1$ Queue.

GI/M/1 Queue

- ▶ Consider the queueing model $GI/M/1$.
- ▶ Customers arrive at time points $0 = t_0, t_1, t_2, \dots$
- ▶ Let $Z_n = t_{n+1} - t_n, n = 1, 2, \dots$ be i.i.d. random variables with distribution function $A(\cdot)$ with mean $1/\lambda$.
- ▶ Let the service time distribution be exponential with mean $1/\mu$.
- ▶ Let $Q(t)$ be the number of customers in the system at time t



What is GI/M/1 queue? It means that the inter arrival time follows a non exponential distribution which are independent. Therefore the GI some books they use on G as a notation. M stands for the service time is exponential distribution. Only one server in the system with the infinite capacity. So consider the customers arrive at a time point say t_0, t_1, t_2 , and so on. Let Z_n is equal to $T_{n+1} - T_n$ with iid random variables with the distribution function with the CDF of a with the mean $1/\lambda$. Therefore, as a special case if you have seen that the inter-arrival time is exponential distribution with the mean $1/\lambda$ then the arrival follows Poisson process with the parameter λ but in this GI/M/1 model the inter arrival time is a non exponential distribution with the CDF function $A(\cdot)$ with the mean $1/\lambda$. Let the service time distribution be exponential with the mean $1/\mu$. Let $Q(t)$ be the number of customers in the system at time t . So $Q(t)$ for $t \geq 0$ is a stochastic process. Since the $Q(t)$ is a number of customers in the system at any time t therefore the corresponding stochastic process is a discrete state continuous time stochastic process. So the underlying stochastic process in the GI/M/1 queue is a $Q(t), t \geq 0$.

GI/M/1 Queue ...

- ▶ Define $Q(t_n - 0) = Q_n, n = 1, 2, \dots$
- ▶ Thus Q_n is the number in the system just before the n th arrival.
- ▶ Then $\{(Q_n, t_n), n = 0, 1, \dots\}$ is a Markov renewal process, where t_n is the instant when the n^{th} customer arrives and $Q_n = Q(t_n - 0)$.
- ▶ Define X_n as the number of potential service completions during the inter-arrival period Z_n .
- ▶ Let $\{b_j, j = 0, 1, \dots\}$ be the distribution of X_n .
- ▶ Then,

$$b_j = P(X_n = j) = \int_0^{\infty} \frac{e^{-\mu t} (\mu t)^j}{j!} dA(t)$$



Now define Q of t_n minus 0 as Q_n . Thus Q_n is a number of customers in the system just before the n th arrival. Q_n is the number of customers in the system just before the n th arrival therefore the Q_n for n is equal to 1, 2 and so on this follows a discrete state discrete time stochastic process. So this is an embedded stochastic process from Q of t , the Q of t is a discrete state continuous time stochastic process whereas Q of n is a discrete state discrete time stochastic process because the Q_n is the number of customers in the system just before the n th arrival. The bivariate random variables Q_n, t_n for different values of n forms a Markov renewal process where t_n is the instant when the n th customer arrives and the queuing is defined Q of t_n minus 0. Since inter arrival time is a non exponential distribution with the mean $1/\lambda$ and the service time is exponential distribution with the mean $1/\mu$ single server in the system and infinite capacity. Therefore the Q_n, t_n form a Markov renewal process. And the T_n is a time instant in which the arrival occurs. Now define the random variable X_n as the number of potential service completions during the inter arrival periods at n . Z_n is nothing but t_n plus 1 minus t_n that is an inter arrival time. The X_n is the number of potential service completions during the inter arrival period is added and P_j be the distribution of X_n . Obviously X_n is their discrete type random variable for fixed n and the P_j is a probability mass function for the random variable X_n . Since the number of potential service completion could be 0, 1 and so on so the probability mass function for different values of j , P_j 's is nothing but the probability of X_n is equal to j that is nothing but the integration 0 to infinity $e^{-\mu t} (\mu t)^j$ divided by $j!$ factorial and the integration with respect to A of t where A of t is the distribution function of inter arrival time with the mean $1/\lambda$ and μ is the parameter for exponential distribution of service time.

GI/M/1 Queue ...

- ▶ The relationship between Q_n and Q_{n+1} is given by

$$Q_{n+1} = \begin{cases} Q_n + 1 - X_{n+1}, & Q_n + 1 - X_{n+1} > 0 \\ 0, & Q_n + 1 - X_{n+1} \leq 0 \end{cases}$$

- ▶ Q_{n+1} is independent of X_{n+1} and Q_{n+1} does not depend on any random variable with an earlier index parameter than n .
- ▶ $\{Q_n, n = 0, 1, \dots\}$ is a time homogeneous discrete time Markov chain.



The way we define the Q_n is nothing but the number of customers in the system just before the n th arrival one can relate the Q_n plus 1 in terms of Q_n . We can find the relationship between Q_n and a Q_n plus 1 that is Q_n plus 1 will be Q_n plus 1 minus X_n plus 1 whenever Q_n plus 1 minus X_n plus 1 is greater than 0; otherwise it is 0 for Q_n plus 1 minus X_n plus 1 is less than 0. The reason is the number of customers in the system just before the n th arrival n plus 1-th arrival the same as the number of customers in the system just before the n th arrival plus the n plus 1-th customer who arrive that is plus 1 minus how many customers are served during the inter arrival period that is X_n plus 1. So if you subtract that Q_n plus 1 by X_n plus 1 you will get Q_n plus 1. Whenever Q_n plus 1 minus X_n plus 1 is greater than 0. If it is less than or equal to 0 then the number of customers will be again 0 just before the n plus 1-th arrival also will be 0. We know that the X_n plus 1 is independent of Q_n plus 1. Hence the Q_n plus 1 depends only on Q_n and Q_n plus 1 is independent of X_n plus 1 therefore the Q_n forms a time homogeneous discrete time Markov chain. The Q_t is a discrete state continuous time stochastic process and the Q_n is the discrete time discrete state stochastic process. Since Q_n plus 1 is equal to Q_n plus 1 minus X_n plus 1 for 0 and Q_n plus 1 is independent of X_n plus 1 as well as Q_n plus 1 depends only on Q_n therefore the Q_n for the n is equal to 0, 1, 2, and so on form a time homogeneous means time invariant and also it satisfies the Markov property. Therefore this discrete time discrete state stochastic process is called a discrete time Markov chain satisfying the time homogeneous property therefore it is called the time homogeneous discrete time Markov chain.

GI/M/1 Queue ...

- ▶ The transition probability is given by

$$\begin{aligned} p_{ij} &= P(Q_{n+1} = j \mid Q_n = i) \\ &= \begin{cases} P(X_{n+1} = i - j + 1), & j > 0 \\ P(X_{n+1} \geq i + 1), & j = 0. \end{cases} \end{aligned}$$

- ▶ In matrix form,

$$P = \begin{bmatrix} \sum_{i=1}^{\infty} b_i & b_0 & 0 & 0 & \dots \\ \sum_{i=2}^{\infty} b_i & b_1 & b_0 & 0 & \dots \\ \sum_{i=3}^{\infty} b_i & b_2 & b_1 & b_0 & \dots \\ \vdots & \vdots & \vdots & \vdots & \ddots \end{bmatrix}$$



Once we know that Q_n is a discrete time Markov chain you can find the one-step transition probability matrix whose elements are P_{ij} 's. So the P_{ij} 's are nothing but what is the probability that Q_n plus 1 will be j given that Q_n was i that is same as what is the probability that i plus j minus 1 customers are served during the inter arrival time period. That is a probability of X_n plus 1 is equal to i minus j plus 1 whenever j is greater than 0. If j is equal to 0 then it is nothing but a probability that X_n plus 1 is equal to i plus 1.

GI/M/1 Queue ...

- ▶ Define $Q(t_n - 0) = Q_n, n = 1, 2, \dots$
- ▶ Thus Q_n is the number in the system just before the n th arrival.
- ▶ Then $\{(Q_n, t_n), n = 0, 1, \dots\}$ is a Markov renewal process, where t_n is the instant when the n^{th} customer arrives and $Q_n = Q(t_n - 0)$.
- ▶ Define X_n as the number of potential service completions during the inter-arrival period Z_n .
- ▶ Let $\{b_j, j = 0, 1, \dots\}$ be the distribution of X_n .
- ▶ Then,

$$b_j = P(X_n = j) = \int_0^\infty \frac{e^{-\mu t} (\mu t)^j}{j!} dA(t)$$



In matrix form you can write it P as the matrix whose elements are P_{ij} so the first element will be P_1 plus P_2 and so on and the second element in the first row will be b_0 . Since P_j 's are nothing but the probability mass function for the random variable X_n that is for all n , for all n it is identically distributed. Therefore the probability mass function of X_n is P_j 's and the running index for j is 0, 1, and so on therefore if you make the row sum b_0 plus b_1 plus b_2 and so on that will be one.

Whereas in the second row the first element will be b_2 plus b_3 and so on. The second row second element will be b_1 . Second row third element will be b_0 and so on. substitute i and j in the above equation. You substitute i and j accordingly you will get this values summation of b starting from 2 and b_1, b_0 and so on. Similarly you can get the third row. You can verify that each element of B capital matrix will be lies between 0 to 1 and the row sum will be 1. It's basically a stochastic matrix. This is a one-step transition probability matrix capital P.

Our interest is to find out the steady state or limiting distribution. For that we need irreducible and positive recurrent and a periodicals also.

GI/M/1 Queue ...

- ▶ For the DTMC to be irreducible, $b_0 > 0$ and $b_0 + b_1 < 1$. We can easily determine that the DTMC is aperiodic.

- ▶ Let

$$\phi(\theta) = \int_0^{\infty} e^{-\theta t} dA(t), \quad \text{Re}(\theta) > 0$$

- ▶ The probability generating function of $\{b_j\}$ is obtained as

$$\beta(z) = \sum_{j=0}^{\infty} b_j z^j = \phi(\mu - \mu z), \quad |z| \leq 1$$

- ▶ Hence,

$$E(Z) = \beta'(1) = -\mu \phi'(0) = \frac{\mu}{\lambda}$$

- ▶ We define the traffic intensity

$$\rho = (\text{arrival rate}) / (\text{service rate})$$



So for the irreducible you need a b_0 has to be greater than 0 as well as $b_0 + b_1$ has to be less than 1. If this is satisfied then you will get the conclusion the given time homogeneous discrete time Markov chain will be irreducible. That means that each state is communicating with each other states with the condition of b_0 is greater than 0 $b_0 + b_1$ is less than 1.

We can easily determine that the DTMC is aperiodic. We have discussed the aperiodic in the discrete-time Markov chain so we can verify that this is a – this discrete time Markov chain is aperiodic also that means with the period 1. Now we find out the Laplace transform of the CDF of inter arrival time distribution that is a P_i . So P_i is equal to integration 0 to infinity $e^{-\theta t} dA(t)$ where $A(t)$ is a CDF of inter arrival time distribution with the real of θ has to be greater than 0. Now we are finding the probability generating function for P_j 's that is a distribution of X_n . So that is $\beta(z)$ that's the summation j is equal to 0 to infinity $P_j z^j$. That you can write down in terms of Laplace transform. So this is Laplace transform. So Laplace transform of $A(t)$ so it's a $\phi(\mu - \mu z)$.

Now we can find out the expectation of Z that is nothing but the $\beta'(1)$ if you differentiate probability generating function then substitute Z is equal to 1 will be the mean of – mean number of arrivals that is Z .