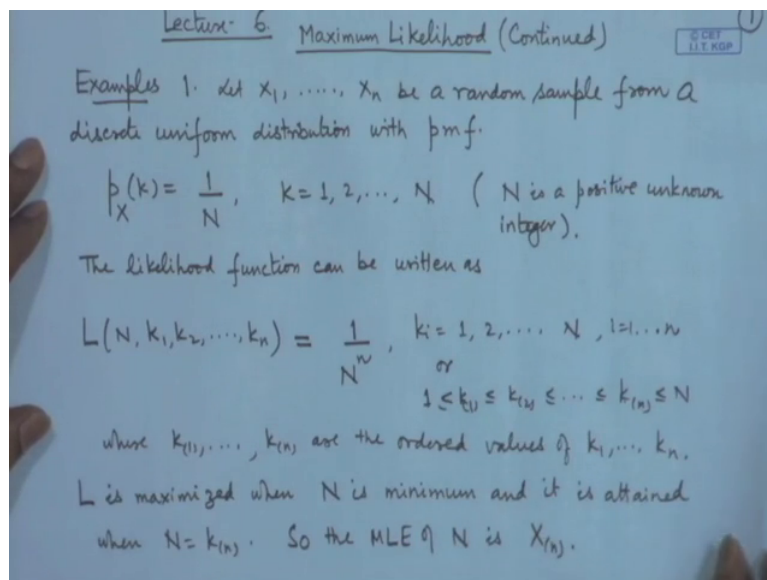


Statistical Inference
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Module No. # 01
Lecture No. # 06
Finding Estimators – III

Yesterday we have discussed in detail various probability models, and how to find out the maximum likelihood estimators for that. We have seen here that the effect of changing the parameter space, or effect of the prior information on the parameter space plays an important role in the maximum likelihood estimation, which makes it different from the other methods, such as unbiased estimation or the method of moment's estimation. So, today I will explain this method with the help of several other examples, and we will discuss certain important large sample properties of the maximum likelihood estimators.

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Let me start with the, a couple of examples on discrete distributions. So, let us consider, say a discrete uniform distribution. Let $x_1 \times 2 \times n$ be a random sample, from a discrete uniform distribution. So, a discrete uniform distribution is usually concentrated on n points, and

normally we take the points from 1 to n , and each one will be equal probability. So, we can consider the probability mass function as follows, with probability mass function given by. So, we write $p(x = k) = \frac{1}{N}$, where k can take values $1, 2, \dots, N$. Now, in this case there may not be any inference problem, if we know on how many points the distribution is concentrated. The inference problem arises if we do not know how many points are there. So, this type of situation may arise, where we know that each possibilities with equal probability, but how many possibilities are there that may not be known. So in that case, we may be interested in estimating that number. So, we are assuming here that n is a positive unknown integer. So, we proceed as before, we write down the likelihood function, which is the joint distribution of x_1, x_2, \dots, x_n .

So, we consider points x_1 is equal to k_1 , x_2 is equal to k_2 , x_n is equal to k_n . So, we can write it in the following fashion. The likelihood function can be written as $L(N)$, and as I mentioned, we are considering the points k_1, k_2, \dots, k_n , which are the observed values of the random variables x_1, x_2, \dots, x_n respectively. So, that is equal to $\frac{1}{N^n}$, where each of the k_i 's can take values $1, 2, \dots, N$ or i is equal to 1 to N . Now, the problem here is to, maximize this function with respect to n . As n is appearing at the denominator, it will be the minimum value of n . So, this will be maximized when n is taking the minimum value. Now, what is the minimum value of n that is possible here. So, this reason we can write it in a more appropriate fashion; that k_1, k_2, \dots, k_n , suppose i order then, $k_1 \leq k_2 \leq \dots \leq k_n$, then the reason can be written as; $k_1 \leq k_2 \leq \dots \leq k_n \leq N$.

So, from here it is clear that the minimum value of n , that is possible is the maximum value of k_1, k_2, \dots, k_n , where k_1, k_2, \dots, k_n are the ordered values of k_1, k_2, \dots, k_n . So, L is maximized, when N is minimum and it is attained, when N is equal to k_n . Now k_n corresponds to the largest order statistics here. So, we conclude that the maximum likelihood estimator of N is x_n . You can notice the analogy with the continuous uniform distribution, which we discussed in the previous class. In the continuous uniform distribution on the interval zero to θ , the maximum likelihood estimator for θ was also the largest order statistics, that is x_n . So, in the discrete uniform case also the same thing is happening. The only difference here is that, here x_i 's are taking positive integral values here.

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Hypergeometric Distribution

$$N \rightarrow \text{Size of the population}$$

$$\swarrow \quad \searrow$$

Cat A Cat B

$$M \quad \quad N-M$$

Suppose a random sample of size n is taken from the population and let X denote the number of items of type (Cat A) in the sample. Then

$$P(X=x) = \frac{\binom{M}{x} \binom{N-M}{n-x}}{\binom{N}{n}}, \quad \max(0, n-N+M) \leq x \leq \min(M, n)$$

Case I: M is known, but N is unknown.
We have to find MLE of N .

Let us take another important discrete distribution; that is hypergeometric distribution. Now a hypergeometric distribution is usually considered in the following fashion, that there is a large population of size n . This is the size of the population. Now this population is divided into two parts. Let us say category A and category B. The entire population, for example we may divide, employees of an organization by two categories that is those who are in the supervisory position and those who are in the working conditions, and that is they are the lower level employees and the higher level employees.

We may divide the patients into two groups; say those who are having communicable diseases, those who do not have communicable diseases. We may divide a section of a student into the students, who are following engineering discipline and the others who are studying say medical discipline. So, we have a large population, and the population size of one category is M , and therefore the other category population has N minus M numbers. Suppose, we take a random sample **a random sample** of size small n , is taken from the population. And let x denote the number of items, items means it could be persons or anything, of type category A in the sample. Then the probability distribution of x is given by $\frac{M \text{ C } x, N \text{ minus } M \text{ C } n \text{ minus } x}{N \text{ C } n}$.

Now obviously, this random variable x , it can take values from 0 to n , because in a random sample of size n , you may have none of this category and all of the other category, 1 of 1 category, n minus 1 of another category and so on. However, this is also subject to the

restrictions of the total elements of each type, and therefore we may write the restrictions in a more strict sense as; that is x is a integer between maximum of 0 and n minus N plus M to minimum of n M . Now, when we look at this probability model, there can be two different cases; one case could be, that the total population size is unknown. Now this type of situation arises for example, in estimating say, we have a lake and a company which is involved in the fishing. It may like to estimate that how much of fish amount will be available in the lake, if they start the fishing operations. Now; obviously, one cannot take out the water from the lake, and count the how many fish will be there.

So, we assume that the size of the population that is capital N is unknown. Now one may conduct the following experiment, which is known as capture recapture technique. We take a random sample of size capital N from the lake. The fish that are taken out they are tagged; that means, they are marked with something, then they are shifted back to the lake. So, that they get mixed up with the entire population of the fish. Later on we consider a random sample of size n , from the fish once again; from the lake we again take a random sample of size n . Now out of that you look at how many of them are tagged, and how many of them are untagged. So, now, this capital M is known to you and capital N is not known to us, and the problem will come how to estimate capital N .

Similarly, there can be another problem, where the total population size is known, we may like to estimate how many people are suffering from a certain diseases or a certain virus. For example, how many people are infected with H I V virus. In that case, we again take a sample of size n . And in that sample, x will denote the number of people who are actually infected with the virus, and then on the basis of that we estimate N . So, in this case capital N may be known, but capital M is unknown. So, when we consider this hyper geometric model, there are two cases. So, case one is that M is known, but N is unknown. So in this case, we have to find the maximum likelihood estimator of N . In order to do that, we write the likelihood function. Now in this case the observation is, the sample of size n has been taken, and x is the number of items of type category A . So, this is the recorded item. So, this function itself denotes the likelihood function in this particular case, because this is the probability mass function of the observation here.

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The likelihood fn. is

$$L(N, x) = \frac{\binom{M}{x} \binom{N-M}{n-x}}{\binom{N}{n}}$$

$$\frac{L(N, x)}{L(N-1, x)} = \frac{\binom{M}{x} \binom{N-M}{n-x} / \binom{N}{n}}{\binom{M}{x} \binom{N-1-M}{n-x} / \binom{N-1}{n}}$$

$$= \frac{M! (N-M)! x! (M-x)! (n-x)! (N-1-M-n+x)! n! (N-n)! (N-1)!}{x! (M-x)! (n-x)! (N-M-n+x)! M! (N-1-M)! N! n! (N-1-n)!}$$

$$= \frac{(N-n) (N-M)}{N (N-M-n+x)} > 1 \text{ if } N < \frac{nM}{x}$$

$$< 1 \text{ if } N > \frac{nM}{x}$$

So we observe that the $L(N, x)$ achieves its maximum when

So, the likelihood function is, let me call it $L(N, x)$ here. So, that is equal to $\binom{M}{x} \binom{N-M}{n-x} / \binom{N}{n}$. And we need to maximize this with respect to capital N . Now the methods that I mentioned in the previous examples, cannot be directly implemented here. The main reason is that here n is an integer, so we cannot apply differentiation procedure taking log etcetera. So, we carry out a different analysis. Let us write down, we try to see the increasing or decreasing nature of this function in a straightforward fashion. Let us consider for example, the value of the likelihood function at N , and the value of the likelihood function at N minus 1. So, this is $\binom{M}{x} \binom{N-M}{n-x} / \binom{N}{n}$ divided by $\binom{M}{x} \binom{N-1-M}{n-x} / \binom{N-1}{n}$. And then this whole thing is divided by $\binom{M}{x} \binom{N-M}{n-x} / \binom{N}{n}$ divided by $\binom{M}{x} \binom{N-1-M}{n-x} / \binom{N-1}{n}$. We may expand the factorial here, so we will get $M!$ factorial divided by $x!$ factorial,

$(M-x)!$ factorial, $(n-x)!$ factorial, $(N-M-n+x)!$ factorial, then we have $(N-1-M-n+x)!$ factorial, then $n!$ factorial, $(N-n)!$ factorial, and then we had this $N!$ factorial and $(N-1-n)!$ factorial. So, we write that also, $N!$ factorial, $n!$ factorial, $(N-1-n)!$ factorial. And in the similar way this will be $(N-1-M)!$ factorial, $n!$ factorial, $(N-1-n)!$ factorial. So, it is easy that one can simplify these terms, and we get it as $(N-n)!$ factorial, $(N-M)!$ factorial, $N!$ factorial, $(N-M-n+x)!$ factorial, into $N!$ factorial, $(N-1-M)!$ factorial, $n!$ factorial, $(N-1-n)!$ factorial.

minus M divided by N into N minus M plus x . Now you notice that, this is greater than 1 if N is less than nM by x , and it is less than 1 if N is greater than nM by x .

Now obviously, you can see N is taking integer values from 1 2 and so on. Now this ratio; that is $\frac{LNx}{LN-1x}$, divided by $\frac{LN-1x}{LN-2x}$. So, what we are observing here is that, if i increase N . From $N-1$ to n if I go, then this ratio is greater; that means, it is an increasing function of N , when N is less than nM by x . And when N is bigger than nM by x , then this value starts decreasing. Therefore, you can say that this function increases till this and then decreases, therefore the maximum of $L(N)$ function is achieved when N is equal to nM by x . Now, naturally nM by x need not be an integer, although x and M are integers, but this expression need not be an integer. So, we may take the integral portion of nM by x as the maximum likelihood estimator for N .

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$N = \frac{nM}{x}$. As $\frac{nM}{x}$ need not be an integer, we take $\lfloor \frac{nM}{x} \rfloor$ (the largest integer less than or equal to $\frac{nM}{x}$) as the MLE of N .

Case II: M is unknown, N is known.

We want to find the MLE of M .

The likelihood fn. is

$$L^*(M, x) = \frac{\binom{M}{x} \binom{N-M}{n-x}}{\binom{N}{n}}$$

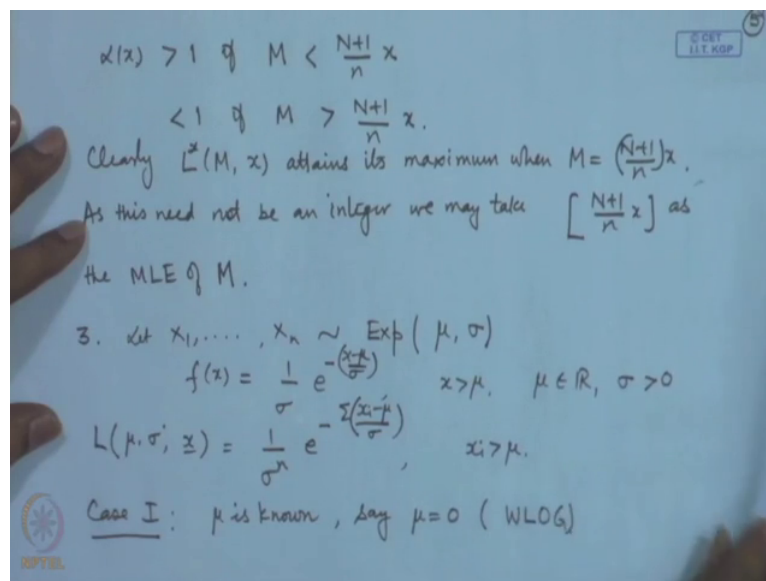
$$\alpha(x) = \frac{L^*(M, x)}{L^*(M-1, x)} = \frac{\binom{M}{x} \binom{N-M}{n-x}}{\binom{M-1}{x} \binom{N-M+1}{n-x}} = \frac{M(N-M+1-n+x)}{(N-M+1)(M-x)}$$

So, we observe that, the L function achieves its maximum, when N is equal to nM by x . As nM by x need not be an integer, we take nM by x integral portion; that is the largest integer, less than or equal to nM by x , as the maximum likelihood estimator of N . Now, let us take up the other case, when M is unknown, M is unknown and N is known. So, here we want to find out, the maximum likelihood estimator of M . Now, once again, if you consider this likelihood function here. I wrote here it as a function of N , because this is coming from the probability mass function of x . Here M and N both are involved. Now, if N is known and M is unknown, I will consider the likelihood function as a function of N . So, the likelihood function will

become. Although, it will be the same expression, it will be written as $L(M, x)$. Let me call it L^* .

So, this is $\binom{N}{M} x^M (1-x)^{N-M}$. Now as before, we have to consider the maximization of this, with respect to M . Now M is an integer and the factorials are involved here, therefore we cannot apply the usual methods of analysis; such as differentiation etcetera, rather we try to see the behavior of this in a straightforward fashion. So, once again we write $L^*(M, x)$ divided by $L^*(M-1, x)$. Now that is equal to $\frac{\binom{N}{M} x^M (1-x)^{N-M}}{\binom{N}{M-1} x^{M-1} (1-x)^{N-M+1}}$, when we write this ratio $\frac{\binom{N}{M}}{\binom{N}{M-1}}$ will be same, so that will cancel out, and we will get $\frac{M}{N-M+1} \frac{x}{1-x}$. Now as before, we can simplify this, and the term turns out to be $\frac{M}{N-M+1} \frac{x}{1-x}$. Now, once again we observed that this ratio, let me call it $\alpha(x)$.

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So, if we observe this ratio, $\alpha(x)$ is greater than 1 if M is less than $N + 1$ by $n x$, and it is less than 1 if M is greater than $N + 1$ by $n x$. So, we can easily see that, the L^* function, it is increasing for M less than $N + 1$ by $n x$, and it will start decreasing for M greater than this. Therefore, the maximum will be attained at $N + 1$ by $n x$, and therefore we can consider the integral part of this, as the maximum likelihood estimator for N . Clearly $L^*(M)$ attains its maximum, when M is equal to $N + 1$ by $n x$. As this need not be an integer, we may take the integral portion of this as the MLE of M . So, here we have seen that, in the discrete case the method of obtaining the maximum likelihood estimators differs little bit. We

have not considered another important distribution which arises quite often in statistical modeling; that is an exponential distribution. Now, the exponential distribution once again has two parameters. It may have a scale parameter; it may have a location parameter. So, I will consider a general model, and then we look at the solution here. Let x_1, x_2, \dots, x_n follow an exponential (μ, σ) distribution, when I say this we are writing down the density function as, $\frac{1}{\sigma} e^{-\frac{x-\mu}{\sigma}}$ for $x > \mu$, where x is greater than μ .

Here μ can be any real number and σ is positive. In the usual study which are related to reliability and life testing, there μ is considered as the minimum guarantee time and there μ will be positive, but in many other applications it need not be so. So I am taking the general case where μ can take any real value, and σ of course, is associated with the average, therefore σ is greater than 0. So, we consider the likelihood function here, $\frac{1}{\sigma^n} e^{-\sum \frac{x_i - \mu}{\sigma}}$. Now when we are dealing with the two parameters situation; one may have different cases. It may happen that the minimum guarantee time is fixed, and therefore we may take it to be 0. It may happen that σ is fixed, and therefore we may take it to be one. So, we consider these cases. So, case 1; let us consider say μ is known, so we may take without loss of generality, this to be 0. If that is so, then we may write the likelihood function. If we substitute μ is equal to 0, the form of this function becomes much simple.

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Then the likelihood function can be written as

$$L(\sigma, \mathbf{x}) = \frac{1}{\sigma^n} e^{-\frac{\sum x_i}{\sigma}}, \quad x_i > 0$$

$$l(\sigma) = \log L(\sigma, \mathbf{x}) = -n \log \sigma - \frac{\sum x_i}{\sigma}$$

$$\frac{dl}{d\sigma} = -\frac{n}{\sigma} + \frac{\sum x_i}{\sigma^2} = \frac{\sum x_i - n\sigma}{\sigma^2} > 0 \quad \text{if } \sigma < \bar{x}$$

$$< 0 \quad \text{if } \sigma > \bar{x}$$

So $\hat{\sigma}_{ML} = \bar{x}$

Case II: σ is known, say $\sigma = 1$ (WLOG).

And we get it as, then the likelihood function can be written as, $L(\sigma) = \prod_{i=1}^n \frac{1}{\sigma} e^{-\frac{x_i}{\sigma}}$, where each x_i will be greater than zero. So, we write down the log likelihood function that is equal to $-\ln L(\sigma) = -n \ln \sigma - \sum_{i=1}^n \frac{x_i}{\sigma}$. So, now, this is a straightforward function for σ , we can consider the derivative with respect to σ , and we get $-\frac{n}{\sigma} + \sum_{i=1}^n \frac{x_i}{\sigma^2}$. So, this will become $\frac{-n\sigma + \sum_{i=1}^n x_i}{\sigma^2}$, which gives us $\sum_{i=1}^n x_i - n\sigma$ by σ^2 ; obviously, you can study its behavior. It will be greater than 0 if σ is less than \bar{x} . It will be less than 0 if σ is greater than \bar{x} . So, if we consider the plotting of the curve as a function of σ . If we plot $L(\sigma)$, now σ is of course positive, so this is starting from zero.

So, this is increasing till \bar{x} and thereafter it is decreasing, because our derivative is positive, for $\sigma < \bar{x}$ and it is less than 0, for $\sigma > \bar{x}$. Therefore, easily you can see that the maximum occurs at \bar{x} . So, the maximum likelihood estimator of σ turns out to be the mean of the distribution. Now, here as before like we have considered in the normal distribution, one may have additional information about σ . For example, σ may be having an upper bound; such as $\sigma \leq \sigma_0$, or $\sigma \geq \sigma_0$ or σ may lie in an interval. In that case, the solutions will, for the maximum likelihood estimator will get modified accordingly, as we have discussed in the case of normal distribution. So, I will be skipping those descriptions here. Let us take up the second case, when σ is known **when sigma is known** we can take it to be one without loss of generality. Now in this case the likelihood function can be written as.

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In this case the likelihood function can be written as

$$L(\mu, \mathbf{z}) = \begin{cases} e^{n\mu - \sum z_i} & , \quad z_i \geq \mu, i=1, \dots, n \\ 0 & , \quad \text{elsewhere,} \end{cases}$$

We can see that $L(\mu)$ is maximized when μ takes its maximum value and that is $x_{(1)}$ here.

So $\hat{\mu}_{ML} = X_{(1)}$.

Case III : Both μ and σ are unknown.

$$\log L(\mu, \sigma, \mathbf{z}) = -n \log \sigma - \frac{\sum z_i}{\sigma} + \frac{n\mu}{\sigma}, \quad \mu \leq x_{(1)} \leq \dots \leq x_{(n)}$$

So $\hat{\mu}_{ML} = X_{(1)}$

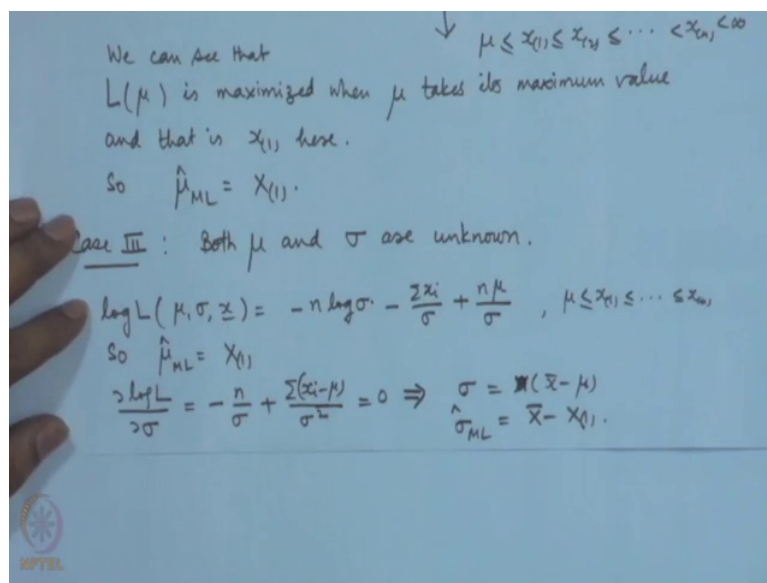
$$\frac{\partial \log L}{\partial \sigma}$$

So, this is now a function of μ , because σ is known. So, if you look at the form that I have discussed here, 1 by σ to the power n e to the power $n\mu - \sum z_i$ by σ . So, here if I put σ is equal to 1 this term vanishes, and you are left with only the exponent term, which I can simply write as e to the power $n\mu - \sum z_i$. So, e to the power $n\mu - \sum z_i$, and of course each x_i is greater than μ . And obviously, this is 0 if. Let me say elsewhere, each of x_i has to be greater than μ in this particular case. Now, if you look at this function, we have to maximize this with respect to μ here. And this $n\mu$ is occurring in the exponent without any multiplication or any other involvement of any other term. So, naturally you can easily see that, the maximization will occur for the maximum value of μ .

Now, what is the maximum possible value of μ . Now, μ is less than each of the x_i 's, therefore this reason can be written as $\mu < x_1 < x_2$ and so on. Therefore, the maximum value of μ can be only x_1 . So, the maximum likelihood estimator of μ is x_1 in this case. We can see that $L(\mu)$ is maximized, when μ takes its maximum value and that is x_1 here. Once again here this x_1, x_2, \dots, x_n denotes the order statistics of the original observations. So, $\hat{\mu}_{ML}$ is equal to the minimum of the observations. So, you have seen in the uniform distribution, we got the maximum of the observations. And in this particular case we are getting the minimum of the observations. Now, let us take the more important case, when both the parameters μ and σ are unknown.

Now let us go back to the original likelihood function, it was $1/\sigma^n \exp(-\sum x_i/\sigma)$. So, we consider the log of this, now this is having two parts; one part is involving only μ and other part is involving σ also. So, for convenience we take the log of this. So, log of likelihood function that is equal to $-\ln \sigma^n - \sum x_i/\sigma$. Now, you can see here, the role of μ is quite different, and when we consider the maximization with respect to μ , it will be attend at the maximum value of μ . So, we can easily then see that, as before the maximum value that it can take is. So, $\hat{\mu}_{ML}$ will remain to be $x_{(1)}$.

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However, for maximization with respect to σ , we can apply the usual calculus here. So, you can consider derivative with respect to σ ; that will be equal to $-n/\sigma - \sum x_i/\sigma^2$. So, we may actually put it together, because this was this term. Now this is equal to 0. If you put this you get σ is equal to $n \bar{x} / (n - 1)$, so this n gets cancelled out. So, the maximum likelihood estimator for σ will be obtained by simply replacing μ by $\hat{\mu}_{ML}$. So, $\hat{\sigma}_{ML}$ is equal to $\bar{x} - x_{(1)}$. Now, you can see here, the effect of partial information and the effect of no information. When the partial information about the parameters was there, then in the case of the estimator of σ , we got \bar{x} , but now you see it is changed to $\bar{x} - x_{(1)}$. Whereas, the effect on the estimation of μ is not there, when σ was known or σ is unknown, the estimation of μ is still the same. Now, in this case I will also consider some special cases.

Here let us consider, when sigma was known. Suppose, I have additional prior information about mu is there, in the form say mu less than or equal to zero. Basically, it means that the minimum guarantee time is upper bounded, by some number say mu naught, which we have brought down to zero. Now, in this case what will happen, if we look at the form of the likelihood function, this function is an increasing function, this function is an increasing function of mu. It starts from minus infinity; that means, it is 0, and then at 0 it will be e to the power something, and then thereafter. Now if you see, if x 1 is here then the maximum is occurring at this point. Whereas, if x 1 is here with respect to 0, then the maximum is occurring here. So, in this case mu head M L, which I will call restricted M L. This will become minimum of x 1 and 0.

So, the role of prior information is important here. You consider the second situation, suppose in place of mu less than or equal to 0 we had mu greater than 0, or greater than or equal to 0 in that case there will be no change, because x 1 is greater than or equal to mu, which will remain greater than 0. So, the maximum occurrence is at x 1, which is within the zone. So, there will not be any change in the maximum likelihood estimator when I am considering the prior information mu greater than or equal to 0. So you can actually see, that the role of the prior information is different in different situations, and this is you can say beauty of the maximum likelihood procedure, that it takes care of each situation individually. So, this is totally based on the likelihood function.

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4. Laplace or Double Exponential distⁿ. 8

Let X_1, \dots, X_n be a random sample from double exponential distribution with pdf.

$$f(x, \mu, \sigma) = \frac{1}{2\sigma} e^{-\frac{|x-\mu|}{\sigma}}, \quad x \in \mathbb{R}, \mu \in \mathbb{R}, \sigma > 0.$$

Case: μ is known, say $\mu = 0$.

$$L(\sigma, \mathbf{x}) = \frac{1}{(2\sigma)^n} e^{-\frac{\sum |x_i|}{\sigma}}$$

$$\ell(\sigma) = \log L(\sigma, \mathbf{x}) = -n \log 2 - n \log \sigma - \frac{\sum |x_i|}{\sigma}$$

$$\frac{d\ell}{d\sigma} = -\frac{n}{\sigma} + \frac{\sum |x_i|}{\sigma^2} = 0 \Rightarrow \frac{\sum |x_i| - n\sigma}{\sigma^2} > 0 \Rightarrow \sigma < \frac{1}{n} \sum |x_i|$$

$$< 0 \Rightarrow \sigma > \frac{1}{n} \sum |x_i|$$

So $\ell(\sigma)$ attains its maximum at $\frac{1}{n} \sum |x_i|$

Now, let us consider another important estimation, which is known as Laplace or Double Exponential Distribution. Laplace or Double Exponential Distribution; So, let x_1, x_2, \dots, x_n be a random sample from double exponential distribution, with the probability density function. Here x is any real number, the parameter μ is any real number and σ is a positive parameter. As before we may have different situations, like μ may be known. So we may put it to be 0, when σ may be known and we may put it to be 1 etcetera. So, let us consider the case, when μ is known, say μ is equal to 0. So, in this case the likelihood function, is $\frac{1}{2^n \sigma^n} e^{-\sum |x_i| / \sigma}$. So, the log likelihood function, that is equal to $-\ln L(\sigma, z) = -n \ln 2 - n \ln \sigma - \frac{\sum |x_i|}{\sigma}$. So, if we consider $\frac{d}{d\sigma} \ln L(\sigma, z) = -\frac{n}{\sigma} + \frac{\sum |x_i|}{\sigma^2}$ that is equal to $-\frac{n}{\sigma} + \frac{\sum |x_i|}{\sigma^2}$, plus $\frac{\sum |x_i|}{\sigma^2}$ minus $\frac{n}{\sigma}$. So, if you put this equal to 0, of course you can adjust the term this is equal to $\frac{\sum |x_i|}{\sigma^2} - \frac{n}{\sigma} = 0$. You can easily see that it is greater than 0, if σ is greater than $\frac{\sum |x_i|}{n}$, if σ is less than $\frac{\sum |x_i|}{n}$, $\frac{\sum |x_i|}{\sigma^2} - \frac{n}{\sigma}$ is less than 0 if σ is greater than $\frac{\sum |x_i|}{n}$.

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Case: μ is known, say $\mu = 0$.

$$L(\sigma, z) = \frac{1}{(2\sigma)^n} e^{-\frac{\sum |x_i|}{\sigma}}$$

$$\ln L(\sigma, z) = -n \ln 2 - n \ln \sigma - \frac{\sum |x_i|}{\sigma}$$

$$\frac{d}{d\sigma} \ln L(\sigma, z) = -\frac{n}{\sigma} + \frac{\sum |x_i|}{\sigma^2} = 0 \Rightarrow \frac{\sum |x_i| - n\sigma}{\sigma^2} = 0$$

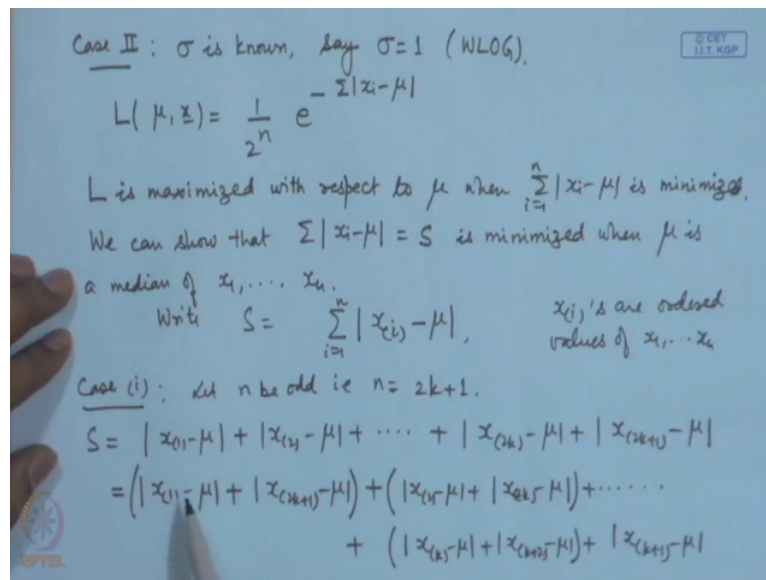
So $\ln L(\sigma)$ attains its maximum at $\frac{1}{n} \sum |x_i|$

So $\hat{\sigma}_{ML} = \frac{1}{n} \sum |x_i|$

Annotations on the right side of the slide:
 > 0 if $\sigma < \frac{1}{n} \sum |x_i|$
 < 0 if $\sigma > \frac{1}{n} \sum |x_i|$

So, the maximum occurs at $\frac{1}{n} \sum |x_i|$. So, σ attains its maximum at $\frac{1}{n} \sum |x_i|$. So, the maximum likelihood estimator of σ is equal to $\frac{1}{n} \sum |x_i|$.

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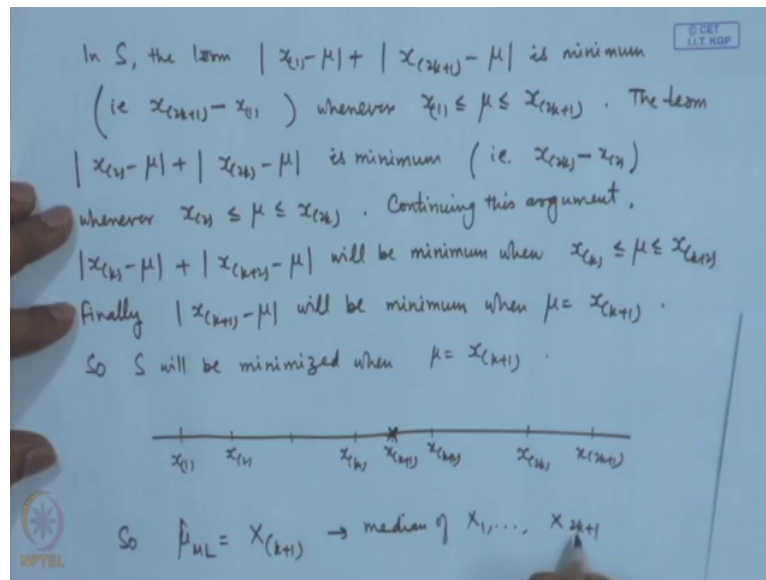


Let us take the second case, when sigma is known and once again, since sigma is a scale parameter we may take it to be 1, without loss of generality. In this case the likelihood function is equal to $\frac{1}{2^n} e^{-\sum |x_i - \mu|}$. Now you see here, this will be maximized with respect to mu if the sum of modulus of $x_i - \mu$ is minimized. L is maximized with respect to mu, when the sum of modulus of $x_i - \mu$ is minimized. Now, one can show that, this is minimized when mu is the median of the observations, because this modulus term is coming, therefore you cannot use the usual differentiation procedure here, however we can give a direct argument. We can show here that the sum of modulus of $x_i - \mu$, let me call it S is minimized, when mu is a median of x_1, \dots, x_n . Let me consider two cases. So, we write S as a sum.

And in place of the x_i 's, we can consider the ordered x_i 's, ordered values of x_1, \dots, x_n ; that means x_1 is the minimum, x_2 is the second minimum and so on as before. Now we give argument in two cases. Let us take n; that is n is equal to something like $2k + 1$. Now, this sum S we express like this, $x_1 - \mu$ plus $x_2 - \mu$ plus and so on, $x_{2k} - \mu$ plus $x_{2k+1} - \mu$. This we express as, say $(x_1 - \mu + x_{2k+1} - \mu) + (x_2 - \mu + x_{2k} - \mu) + \dots + (x_k - \mu + x_{k+1} - \mu) + |x_{k+1} - \mu|$. What we do, we look at the minimization of each of these terms which I

have clubbed together. So, if you look at these 2. Here it is the x_1 and this is x_{2k+1} . If I consider μ to be any value between these 2, then this will turn out to be $x_{2k+1} - x_1$ that will be the minimum value. So, let us write the complete argument here.

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In S the term $x_1 - \mu + x_{2k+1} - \mu$, is minimum, that is the value will be $x_{2k+1} - x_1$, whenever I choose μ to be a number between x_1 and x_{2k+1} . Similarly, the term $x_2 - \mu + x_{2k} - \mu$, this is minimum. And of course, the minimum value will be $x_{2k} - x_2$, whenever x_2 is less than or equal to μ , less than or equal to x_{2k} . So, in that way if you look at all the sums, they will be minimum, whenever μ lies between the two values, which are involved in those two terms, so if we continue this argument. The term $x_k - \mu + x_{k+2} - \mu$ will be minimum, when x_k is less than or equal to μ , less than or equal to x_{k+2} . Finally, $x_{k+1} - \mu$ will be minimum, when μ is equal to x_{k+1} . We have considered the term by term minimization of this S . So, we have taken this, this and this together then this together and so on. We have derived the condition further, minimization of each of these. Now, therefore, the overall minimum will be attained, if all the conditions are simultaneously satisfied.

Now, if you see all the conditions to be simultaneously satisfied, what will be the condition. This is the widest interval, because this is from minimum to the maximum. This interval is the second and so on. So, if I look at this scale here $x_1, x_2, \dots, x_{2k}, x_{2k+1}$, somewhere you have x_k, x_{k+1} and x_{k+2} . So, from the first one, μ should be a new value

between these two. From the second one μ should be a new value between these two, from the third one and so on. And finally, you are getting the value that is x_{k+1} . So, if μ is x_{k+1} , each of these terms that I have clubbed together, they will be the minimum. Therefore, overall S will be minimized. So, S will be minimized, when μ is equal to x_{k+1} , because this will satisfy all the conditions. So, we conclude that μ_{ML} is equal to x_{k+1} , that is actually the median of x_1, x_2, \dots, x_{k+1} , because when the number of observations is odd, the middle value will be median here.

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Case (ii) n is even say $n = 2k$.

$$S = (|x_{(1)} - \mu| + |x_{(2k)} - \mu|) + (|x_{(2)} - \mu| + |x_{(2k-1)} - \mu|) + \dots + (|x_{(k)} - \mu| + |x_{(k+1)} - \mu|)$$

Arguing as before, S will be minimum when $x_{(k)} \leq \mu \leq x_{(k+1)}$ i.e. μ is a median of x_1, \dots, x_{2k} . We may take it to be $(x_{(k)} + x_{(k+1)})/2$.

So we have $\hat{\mu}_{ML} = \text{Med}(x_1, \dots, x_n) = M$.

Case III: Both μ and σ are unknown.

$$L(\mu, \sigma; \mathcal{X}) = \frac{1}{(2\sigma)^n} e^{-\frac{\sum |x_i - \mu|}{\sigma}}$$

Now, let us consider the case when n may be even, n is even say n is equal to $2k$. Now in this case, once again we may consider the clubbing in the similar fashion, however this last term will not be there. Therefore, we will write the clubbing in this fashion $x_1 - \mu$ plus $x_{2k} - \mu$, plus $x_2 - \mu$ plus $x_{2k-1} - \mu$ and so on. In the final it will be $x_m - \mu$ plus $x_k - \mu$ and $x_{k+1} - \mu$. So, if we give the argument as before, arguing as before S will be minimum, when x_m is less than or equal to μ , less than or equal to $x_m + 1$, because now on a scale $x_1, x_2, \dots, x_m, x_{m+1}, x_{2m} - 1, x_{2m}$, this will be here. So, here if you see, the first term here will be minimum when μ lies between the largest intervals. The second one will be minimum when the μ lies between x_2 to x_{2k-1} and so on. The last sum will be minimized, when μ lies between x_k to x_{k+1} .

Now, if μ lies between x_k to x_{k+1} , when we have even number of observations that is $x_1, x_2, \dots, x_k, x_{k+1}, \dots, x_n$, any number between x_k to x_{k+1} is called a median. For convenience many times we take the average of these two values, that is $x_k + x_{k+1}$ by 2. So, this we conclude that μ is a median of x_1, x_2, \dots, x_n . So, where we may take it to be $x_k + x_{k+1}$ by 2. So, we have μ_{ML} equal to the median of x_1, x_2, \dots, x_n . In both the cases we are getting median, we denote it by say m . So, now let us consider the important case, when both the parameters may be unknown. So, both μ and σ are unknown. In this case the likelihood function is equal to $\frac{1}{2\sigma^n}$ to the power n , e to the power minus $\sum |x_i - \mu|$ by σ .

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$$l(\mu, \sigma) = \log(\mu, \sigma, \mathcal{Z}) = -n \ln 2 - n \ln \sigma - \frac{\sum |x_i - \mu|}{\sigma}$$

$$l \text{ is maximized w.r.t } \mu \text{ when } \sum |x_i - \mu| \text{ is minimum}$$

$$\text{ie at } \mu = \text{Med}(x_1, \dots, x_n)$$

$$\text{So } \hat{\mu}_{ML} = M$$

$$\frac{\partial l}{\partial \sigma} = -\frac{n}{\sigma} + \frac{\sum |x_i - \mu|}{\sigma^2} \text{ is attaining the maximum}$$

$$\text{value at } \sigma = \frac{1}{n} \sum |x_i - \mu|$$

$$\text{So } \hat{\sigma}_{ML} = \frac{1}{n} \sum |x_i - M| \quad (\text{mean deviation about median})$$

So, we take the log here; that is equal to minus $n \log 2$, minus $n \log \sigma$ minus $\sum |x_i - \mu|$ by σ . So, as before the maximization with respect to μ will occur, when $\sum |x_i - \mu|$ is minimum, and we have already shown that this is occurring when μ is a median. So, l is maximized with respect to μ , when $\sum |x_i - \mu|$ is minimized; that is at μ equal to median of x_1, x_2, \dots, x_n . So, μ_{ML} is equal to the median which we are calling M . Now, you look at the solution for σ , if we consider the derivative of l with respect to σ , we get minus n by σ plus $\sum |x_i - \mu|$ by σ^2 . And as before if we argue, this is attaining the maximum value at σ is equal to $\frac{1}{n} \sum |x_i - \mu|$.

Now, you have already obtained the solution form, if you substitute it here you get the maximum value of, the maximized value of likelihood function for sigma equal to $\frac{1}{n} \sum_{i=1}^n |x_i - M|$. So, sigma head M L is equal to $\frac{1}{n} \sum_{i=1}^n |x_i - M|$ which is nothing, but the mean deviation about median. So, today friends we have discussed various probability models, and we have discussed the maximum likelihood estimators for those models. I have tried to cover various cases here. And another thing is that we will take up some different cases, where either the maximum likelihood estimator is not unique, it may not exist. And then we will consider the large sample properties of the maximum likelihood estimators in the next class.