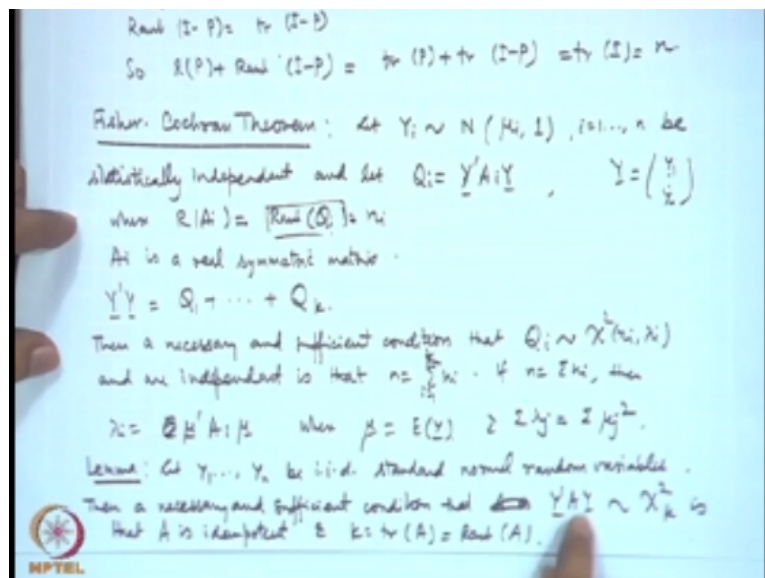


Statistical Methods of Scientists and Engineers
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Lecture - 18
Multivariate Analysis – III

So, we continue our discussion on the multivariate normal distribution and its properties. We have seen various characterizing properties, which also helped us in giving an alternative definition of the multivariate normal distribution. Now, we are trying to see its connections with chi square distribution as in the case of univariate normal distribution.

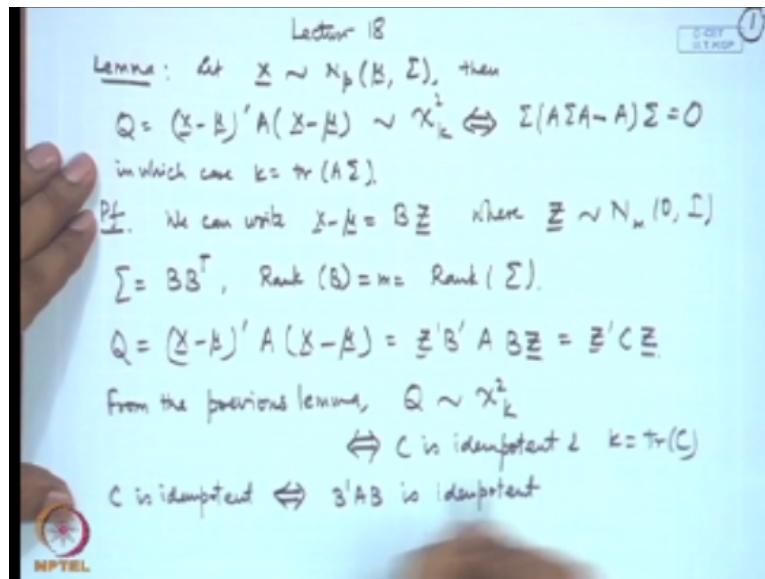
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For that purpose, I stated Fisher-Cochran Theorem and another Lemma, which is saying that $Y' A Y$ will have a chi square distribution. So, this is giving a necessary information condition that if we are having standard random variables, then if I considered $Y' A Y$ that is a quadratic form. This will have chi square k . We know that $Y' A Y$ has a chi square k .

But, if I consider any A here, then for idempotent matrix this will be true. Now, let us consider further results on this. The next result is that if X has a $N_p(\mu, \Sigma)$ distribution, then let us consider say Q that $= (X - \mu)' A (X - \mu)$, then that follows chi square k , this is if and only if $\Sigma A \Sigma - A \Sigma$ is null and in this case, you will have $k = \text{trace of } A \Sigma$.

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Let us look at the proof of this, so you can write $X - \mu = BZ$, if you remember the representation that I obtained for necessary and sufficient condition for the multivariate normal distribution, we were able to write a multivariate normal as $\mu + BZ$, where Z is a vector consisting of the standard normal independent random variables of dimension N . So, let us consider the decomposition of Σ as $B B^T$, rank of B is m , which is also the rank of Σ .

And the quadratic form Q that is $(X - \mu)' A (X - \mu)$. So, since $X - \mu = BZ$, this becomes $Z' B' A B Z$ that we can write as $Z' C Z$ and this $B' A B$, we can write as sum matrix C . Now, if we implement this result that if I am having a collection of a standard normal random variables, then $Y' A Y$ has a chi square k , if and only if A is idempotent.

So that condition will be applied to C and also the trace and rank of A will be $= k$ here. So, if we apply this result, Q will follow chi square k if and only if C is idempotent and $k = \text{trace of } C$ that is rank of C . Now, C is idempotent this condition is equivalent to, so $C = B' A B$. So, $B' A B$ is idempotent so this you can write as $B' A B * B' A B = B' A B$.

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$$\Leftrightarrow (B'AB)(B'AB) = B'AB$$

$$\Leftrightarrow B'(ABR'A - A)B = 0$$

$$\Leftrightarrow B'(A\Sigma A - A)B = 0$$

$$\Leftrightarrow BB'(A\Sigma A - A)BB' = 0$$

$$\Leftrightarrow \Sigma(A\Sigma A - A)\Sigma = 0$$

$$k = \text{tr}(C) = \text{tr}(B'AB) = \text{tr}(AB'B) = \text{tr}(A\Sigma)$$

Remark: if Σ is non-singular then this condition reduces to

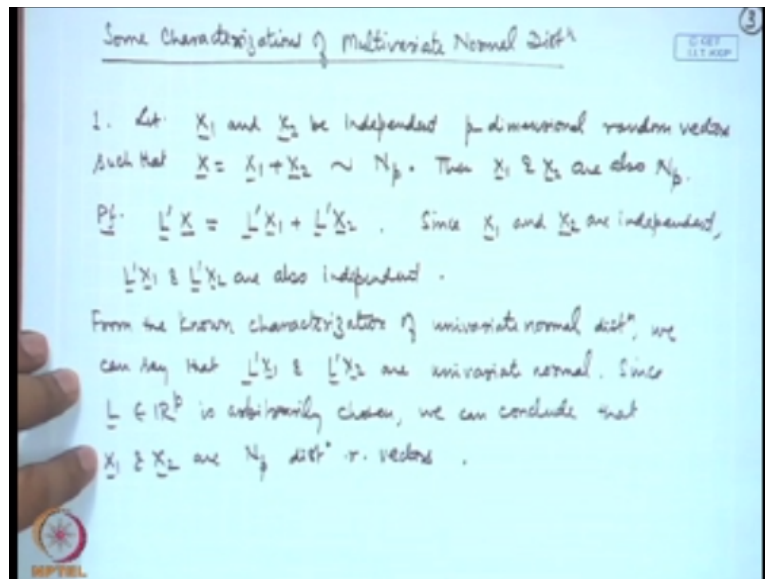
$$A\Sigma A = A$$

So I bring it to the left hand side, so we can write as $AB B' A - A B =$ a null matrix. Now, $B B'$ is Σ , so this becomes $B' A \Sigma A - A B =$ null matrix. Again, this is equivalent to I can pre-multiply by B and I can post-multiply by B' , this is equivalent. Now, a question is that why is this equivalent because if I am having this, I can consider here a transformation from here to get this thing here.

So, this will be implying $C \Sigma A \Sigma A - A \Sigma =$ null. Now, $k =$ trace of C that = trace of $B' A B$ that = trace $A B B'$ because of trace of some matrix $C \cdot D$ is same as trace of $D \cdot C$, so trace of $A \Sigma$. Now as a remark, let me mention here if Σ is non-singular, then I can multiply by Σ^{-1} and Σ^{-1} here, then this condition is $A \Sigma A = A$.

In that way, actually you can say that Σ is a generalized inverse of A that condition will be there. Now before going to, we will also discuss in detailed the noncentral chi square distribution, however, let me talk about certain characterizations of the multivariate normal distribution now. Some characterizations of multivariate normal distribution, let us consider let X_1 and X_2 be independent p -dimensional random vectors such that $X = X_1 + X_2$ follows N_p .

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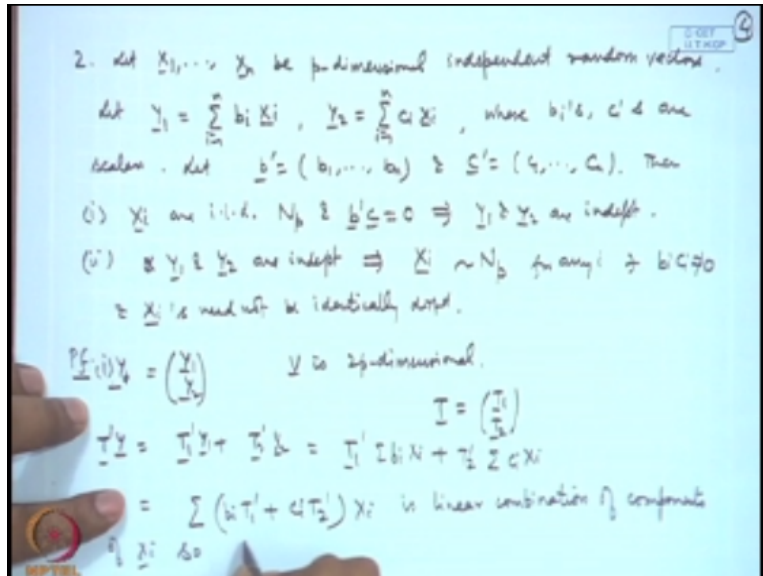


Then X_1 and X_2 are also N_p . Let us look at the proof of this, let us consider say a linear combination of the components of X . So that is becoming L prime $X_1 + L$ prime X_2 . Now, since X_1 and X_2 are independent, L prime X_1 and L prime X_2 are also independent. Now, there is a characterization of the univariate normal distribution in terms of the decomposed terms that means if I say X_1 and X_2 are univariate normal, such that $X_1 + X_2$ follows univariate normal, then each of X_1 and X_2 will be univariate normal.

So from this, we conclude that from the known characterization of, we can say that L prime X_1 and L prime X_2 are univariate normal. Now this L , I chose arbitrarily of p -dimension since L belongs to R^p is arbitrarily chosen, we can conclude that X_1 and X_2 are N_p distributed random vectors. A second characterization is generalization of this, which let me state in the full form here.

Let X_1, X_2, X_n be p -dimensional independent random vectors. Let us consider say $Y_1 = a$ linear combination of say $b_i X_i, i = 1$ to n and Y_2 as an another linear combination of the same where b_i 's and c_i 's, they are scalars. Let us consider say b as b_1, b_2, b_n and c as say c_1, c_2, c_n . Then, we will have the following that is X_i 's are IID N_p and b prime $c = 0$ implies Y_1 and Y_2 are independent.

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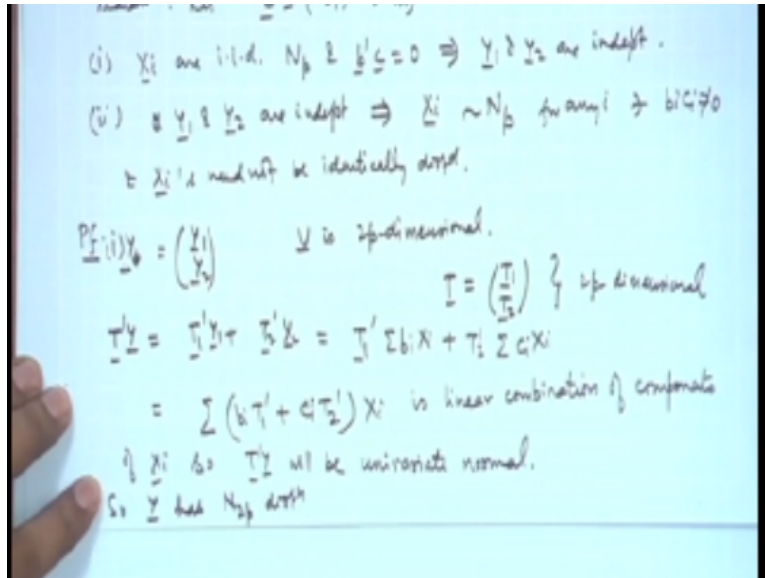


And secondly Y_1 and Y_2 are independent implies that X_i will follow N_p for any i such that $B_i C_i$ is not 0 and X_i 's need not be identically distributed. Let us look at the proof of this, so we can consider the vector Y_1, Y_2 let us call it say Y , I put them in the 2 dimensional form here. So this is now $2p$ -dimensional, so V is $2p$ -dimensional. If I consider linear combination of say T prime Y , then that will become say T_1 prime $Y_1 + T_2$ prime Y_2 , where $T = T_1, T_2$.

If I am assuming that X_i 's are independent random vectors, so in the first part if I am assuming that X_i 's are multivariate normal, then these are linear combinations of the, because what I have done here, Y_1 is a linear combination of X_i 's, so this is becoming T_1 prime sigma $b_i X_i + T_2$ prime sigma $c_i X_i$ that = Sigma $b_i T_1$ prime + $c_i T_2$ prime X_i . So, this is linear combination of components of X_i .

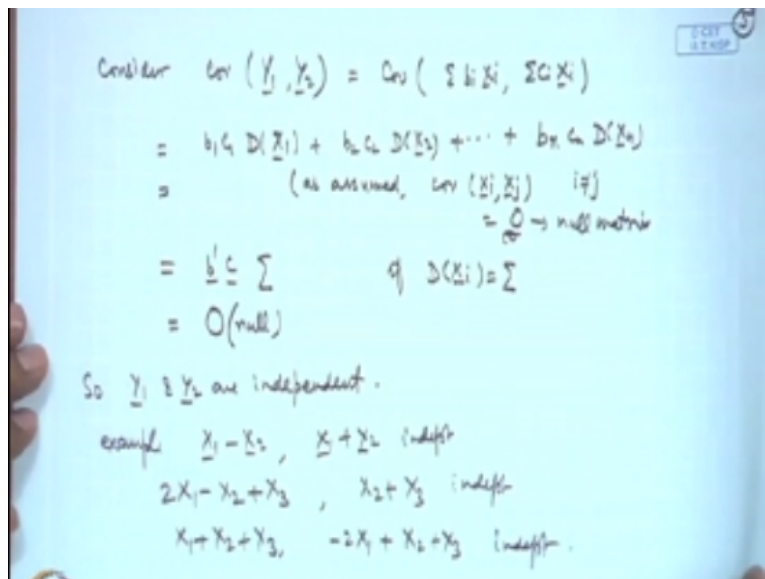
So, T prime Y will be univariate normal. So this is for any T , this is $2p$ -dimensional, so y has N_{2p} distribution that is $2p$ -dimensional multivariate normal distribution. Now, let us consider covariance matrix between Y_1 and Y_2 , now that will be = because I have written this as b prime X , see basically what we are getting here is covariance between Y_1 and Y_2 will become covariance between sigma $b_i X_i$ and sigma $c_i X_i$.

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That will consist of, since X_i 's are independent, this will reduce to $b_1 c_1$ dispersion matrix of $X_1 + b_2 c_2$ dispersion matrix of $X_2 + \dots + b_n c_n$ dispersion matrix of X_n . As we have assumed covariance terms between X_1, X_i, X_j for $i \neq j$, they will be null. So, this is nothing but $b'c$ sigma. If we are writing dispersion matrix of $X_i = \sigma_i$, then this = this.

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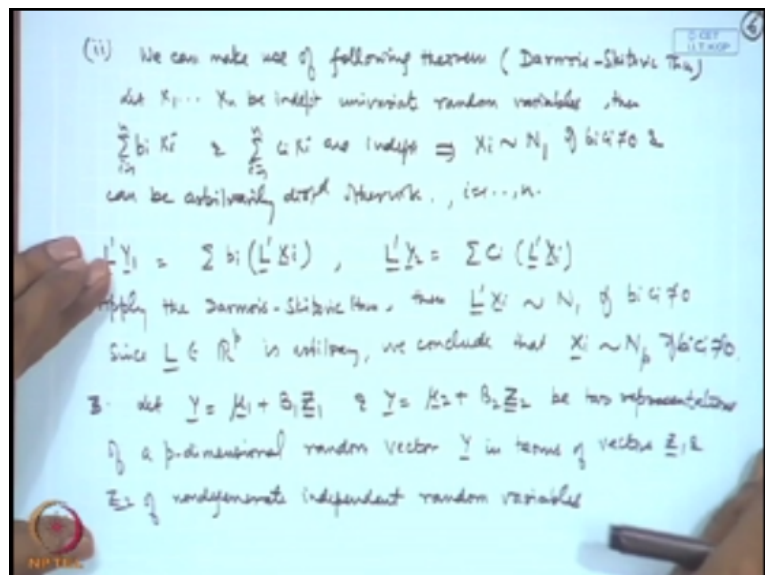


Now, if I am assuming here the $b'c = 0$, then this is simply = null. So, we will get Y_1 and Y_2 are independent. So, this result is proved that if X_i 's are independent and identically distributed p -dimensional multivariate normal distributions where $b_1 c_1 + b_2 c_2 + \dots + b_n c_n = 0$, then this Y_1 and Y_2 are independent. In particular, you may consider something like this.

For example, I take say $X_1 - X_2$ and $X_1 + X_2$, so then they will be independent. Suppose, I consider say $2X_1 - X_2 + X_3$ and say I take $X_2 + X_3$, then they are also independent because if I consider here, $2*0 - 1*1 + 1*1$, so that is going to be 0. If I consider say $X_1 + X_2 + X_3$ and I consider say $-2X_1 + X_2 + X_3$, then here the product is $-2+1+1$, so they are also independent so like that we can construct independent linear combinations here.

Let us look at the part B of this, in the second part, what we are saying is that if Y_1, Y_2 are independent, then X_i 's must be N_p here for any distinct. So let us look at this, so we can make use of, this is called actually Darmois-Skitovic theorem. Let X_1, X_2, X_n be independent univariate random variables, then $\sum_{i=1}^n b_i X_i, i = 1$ to n and $\sum_{i=1}^n c_i X_i, i = 1$ to n are independent, it implies that X_i 's will follow normal 1, if $b_i c_i$ is not 0.

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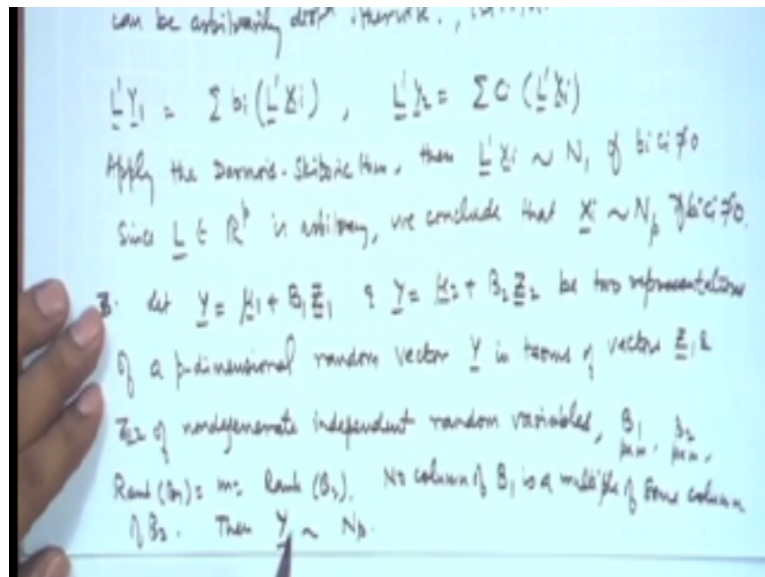


And can be arbitrarily distributed otherwise for $i = 1$ to n . So, let us consider say $L \text{ prime } Y_1$, so that $= \sum b_i L \text{ prime } X_i$ and similarly, $L \text{ prime } Y_2 = \sum c_i L \text{ prime } X_i$. On this, apply the Darmois-Skitovic theorem, then $L \text{ prime } X_i$ this will follow N_1 , if $b_i c_i$ is not 0. So, L is arbitrary vector in p -dimensional space, we conclude that X_i 's will follow N_p , if $b_i c_i$ is not 0.

Now, if you look at the statement this is again very powerful statement. What we are saying is that if I construct linear combinations of p -dimensional random vectors and if they are independent, then each of the terms in the linear combination will have a p -dimensional normal distribution. Of course, we are putting a condition here that $b_i c_i$ must be nonzero that means the corresponding term should be there.

A third characterization is based on the decomposition that I obtained and that we gave as an alternative definition of the multivariate normal distribution also. So, let us consider say $Y = \mu_1 + B_1 Z_1$, let us call it Z_1 and say $Y = \mu_2 + B_2 Z_2$. Suppose, these be 2 representations of a p -dimensional random vector in terms of say vectors Z_1 and Z_2 of nondegenerate independent random variables.

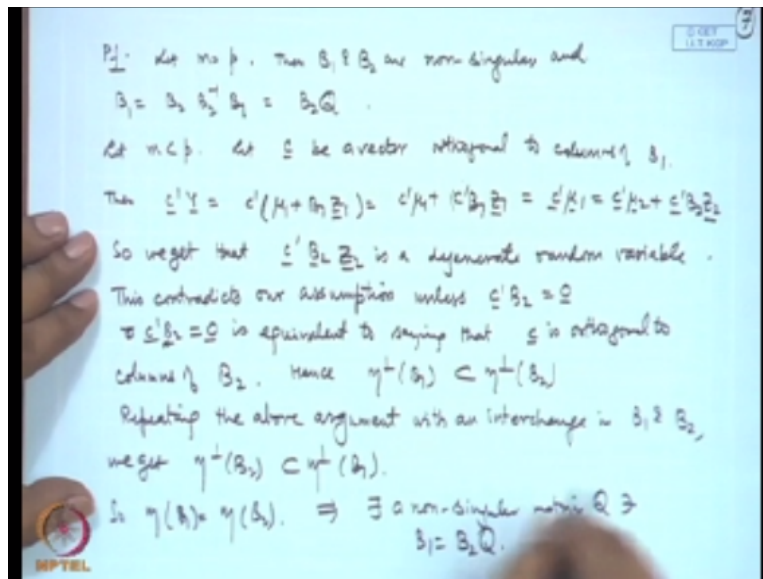
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And this B_1 is a p by m matrix, B_2 is p by m matrix, rank of B_1 is m and rank of B_2 is also m . We also assume that no column of B_1 is a multiple of some column of B_2 . Then, Y follows N_p , so now you see here. I am actually using the representation that I gave as an alternative definition of the multivariate normal distribution, but in that one, Z_1 and Z_2 are vector of IID standard normal variables.

Here, I am saying is that this is the vector of simply nondegenerate independent random variables and then, just by putting a condition on B_1 and B_2 , we are getting that Y must have a multivariate normal distribution. So, this is also very powerful characterization of a multivariate normal distribution. Let us consider say $m = p$, then B_1 and B_2 are nonsingular and then we can write $B_1 = B_2 B_2^{-1} B_1$.

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That we can write as a B_2 and some matrix, this term we can write as some Q . Let us assume say m is $<$ than p . Let C be a vector which is orthogonal to columns of B_1 and we write here $C^T Y$ that $= C^T \mu_1 + B_1 Z_1$, then that is becoming $C^T \mu_1 + C^T B_1 Z_1$, now this will become 0. So, this is simply $C^T \mu_1$ here. Now, that $= C^T \mu_2 + C^T B_2 Z_2$.

Now, what I am getting here $C^T B_2 Z_2 =$ now this is a scalar, so we are getting that $C^T B_2 Z_2$ is a degenerate random variable. Now, we assumed that this Z_1 and Z_2 are vectors of nondegenerate independent random variables. So, here I am getting this as a degenerate random variable. So, this is contradicting our assumption unless we have $C^T B_2 = 0$.

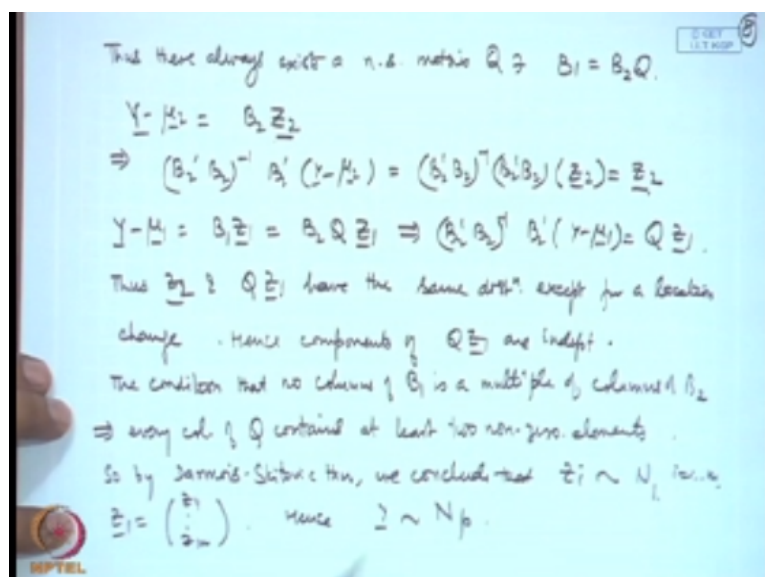
Now, if $C^T B_2 = 0$ is equivalent to saying that C is orthogonal to columns of B_2 . Now let us look at this, I started with C to be a vector which is orthogonal to the columns of B_1 and I am able to prove that C is now orthogonal to the columns of B_2 . So, this means that the orthogonal column space of say B_1 is a subspace of orthogonal column space of B_2 .

Now in this derivation, I have taken B_1, B_2 , now I started with C to be a vector orthogonal to the columns of B_1 in place of that, suppose I write B_2 here, then this statement will change here, here I will get $C^T \mu_2$ and here I will get $B_2 Z_2$, so this will become $C^T \mu_2$ and here then, I can write $C^T \mu_1 + C^T B_1 Z_1$. In that case, I will get the same statement in the reverse way.

So, repeating the argument with an interchange in B_1 and B_2 , we get orthogonal space of B_2 is a subspace of orthogonal space of B_1 , so that means they are same. Basically, column space are B_1 and column space are B_2 are same. This means that there exists a nonsingular matrix Q such that $B_1 = B_2Q$. So, I have written here if $m = p$, then I am able to write to that $B_1=B_2Q$ and if $m < p$ then also I am able to obtain a nonsingular matrix Q such that $B_1=B_2Q$.

So, this one and 2 give that all the time there will be a nonsingular matrix. Thus there always exist a nonsingular matrix Q such that $B_1 = B_2Q$. Now we make use of this, so let us write say $Y-\mu_2$ that = B_2Z_2 . So, this implies B_2 prime B_2 inverse B_2 prime $Y-\mu_2$ that = B_2 prime B_2 inverse B_2 prime B_2Z_2 that = Z_2 . So, if I consider now $Y-\mu_1$ that = B_1Z_1 that = B_2QZ_1 this implies that B_2 prime B_2 inverse B_2 prime $Y-\mu_1$ that will be = QZ_1 .

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So, what we are getting is that Z_2 and QZ_1 they have the same distribution except for a location change. Because both I am able to represent in terms of, see this is $Y-\mu_2$, so μ_2 is the translation here and here I am getting QZ_1 that is $Y-\mu_1$ here. So, components of QZ_1 they are independent. Now, the condition that no columns of B_1 is a multiple of columns of B_2 , then this implies that every column of Q contains at least 2 nonzero elements.

So by Darmois-Skitovic theorem, then we conclude that Z_i 's follows normal N_1 , $i = 1$ to n . So now $Z_1 =$ your components of this, let us call it as Z_{11}, Z_{12}, Z_{1n} . So, what you are getting here is that Y follows N_p . So, these are the 3 characterizations, Now, we move over to the actual density function.

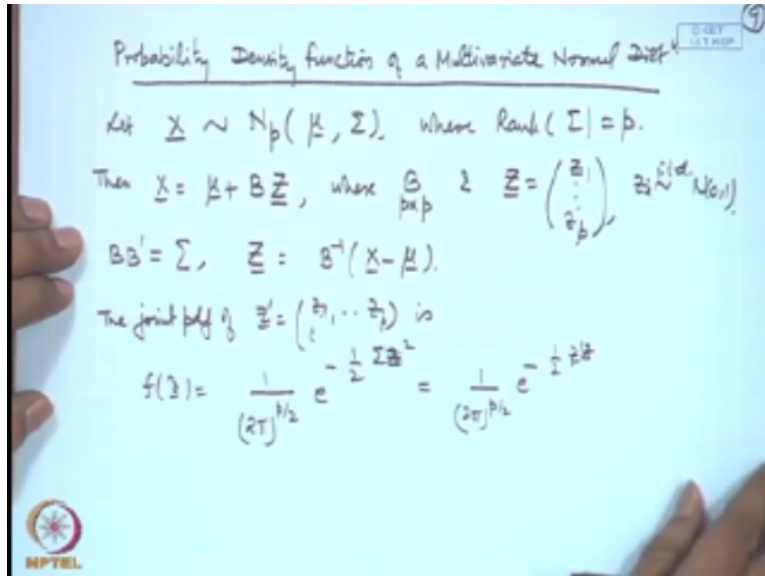
If you remember here in the case of one dimension and 2-dimension distribution, we always define a distribution and we talked about its probability and mass function and the probability density function. In the case of p -dimensional normal distribution, I have not yet actually defined the density function. So, one major reason is that when we talk about higher dimensions.

And if there is, for example here I mentioning σ as a variance-covariance matrix is positive semidefnite. So if it is a positive definite matrix, then it will have full rank, but if it is not a full rank that means the rank say $p-1$ or $p-2$ or in general I am saying $mn < p$ that means there will be some linear relationships among the variables there. If there are complete relationships there, in that case the density will exist on a subspace.

It will not exist on the fully space that was on p -dimensional space. So that is the reason that I gave the definition of the multivariate normal distribution in terms of its linear combinations and then in terms of an alternative representation like $\mu + BZ$ where Z is a collection of m independent univariate normal random variables. So, there m was the rank. So that means I am able to actually define in terms of alternative you can say characterization of the multivariate distribution.

I do not necessarily have to write the density function, but now I will write the density function for the full space that means when I considered the full rank, then we talk about the density function and actually, the representation that I have given, it will be exactly used for deriving the density function. So, we talk about probability density function of a multivariate normal distribution.

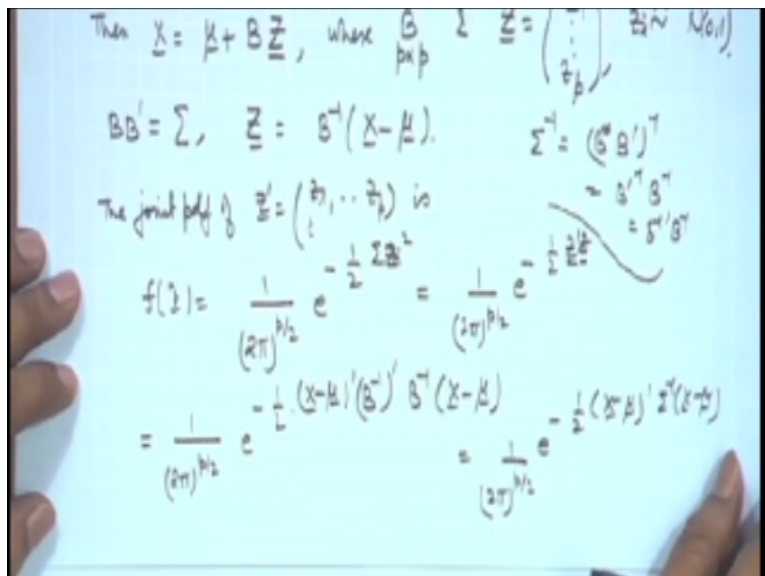
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So, let us consider X following $N_p(\mu, \Sigma)$ and I consider full rank, rank of $\Sigma = p$. if rank of $\Sigma = p$, then we can write $X = \mu + BZ$, where B is p by p and Z is a vector of independent, these are IID normal $0, 1$. So, if that is happening and also this $BB' = \Sigma$ and this $Z = B^{-1}(X - \mu)$. Now, if I have independent normal random variables, I can write down the density function.

So, the joint pdf of $Z = z_1, z_2, \dots, z_p$, so $Z' = z_1, z_2, \dots, z_p$ that is nothing but let me use a notation $f(Z)$. So that $f(Z) = \frac{1}{(2\pi)^{p/2}} e^{-\frac{1}{2} Z' \Sigma^{-1} Z}$. So that will be $f(Z) = \frac{1}{(2\pi)^{p/2}} e^{-\frac{1}{2} Z' B^{-1} B^{-1} Z}$. Let me use capital letters here, usually we write small letters for denoting the value of the random variable, but here for the sake of convenience, I am using the capital letters here.

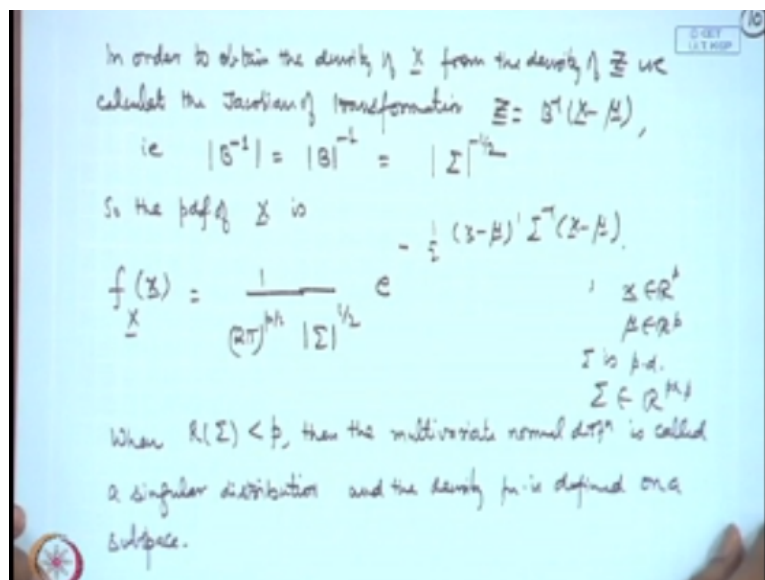
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Now, this Z is given in terms of this, so we write it here that = $1/2 \pi$ to the power $p/2$ e to the power $-1/2$, now $Z =$ this term here, so it is becoming $X - \mu$ prime B inverse prime B Inverse $X - \mu$. Now, if am assuming this $B B$ prime = σ , then σ inverse = $B B$ prime inverse that = B prime inverse B inverse that will be = B inverse prime B inverse. So we can use this, so this is simply becoming $1/2 \pi$ to the power $p/2$ e to the power $-1/2$ $X - \mu$ prime σ inverse $X - \mu$.

Now, if I am obtaining the distribution of X from here, then I have to calculate the Jacobian here. So, what will be the Jacobian term here? In order to obtain the density of X from the density of Z, we calculate the Jacobian of transformation that is $Z = B$ inverse $X - \mu$. So that is given by determinant of B inverse which is same as determinant of B inverse, which is also the determinant of sigma to the power $-1/2$.

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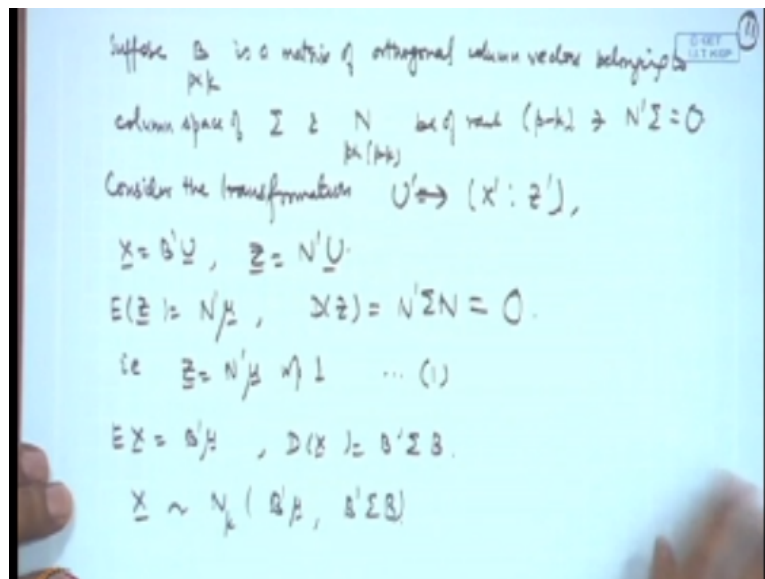


So, the pdf of X is given by that = $1/2 \pi$ to the power $p/2$ determinant of sigma to the power $-1/2$ e to the power $-1/2$ $X - \mu$ prime sigma inverse $X - \mu$. Here, X belongs to R^p , μ belongs to R^p and sigma is positive definite matrix. Sigma is R^p/p that is p/p positive definite matrix. When rank of sigma is $< p$, then the multivariate normal distribution is called a singular distribution and the density function is defined on a subspace.

Suppose B that is p/k is a matrix of orthogonal column vectors belonging to column space of sigma and N that is $p/p-k$ be of rank say $p-k$ such that N prime sigma is null matrix. So, let us consider the transformation, U prime going to X Z prime, where X is B prime U, $Z = N$ prime

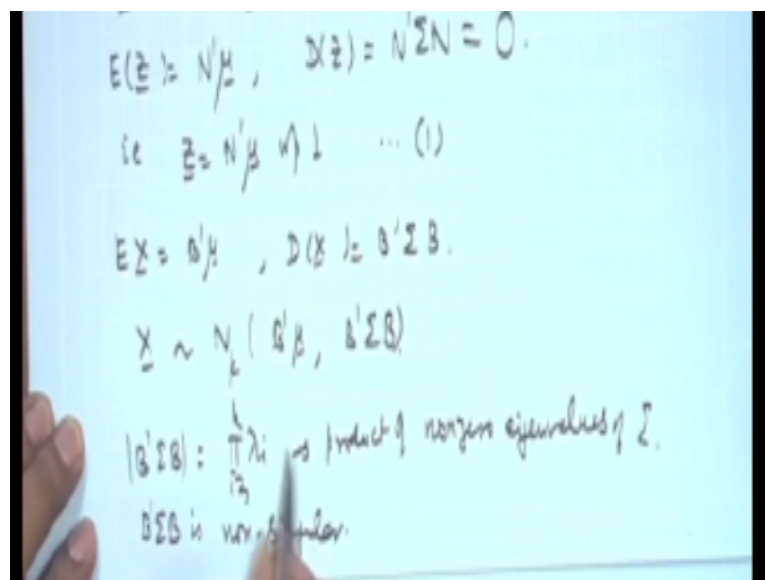
U. Then, expectation of $Z = N$ prime μ , dispersion matrix of $Z = N$ prime σ N that is becoming null.

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That means $Z = N$ prime μ with probability 1 that is degenerate (0) (44:23) and expectation of $X = B$ prime μ , dispersion matrix of $X = B$ prime σ B . So, X follows N_k B prime μ B prime σ B . So, we can write actually B prime σ B can be written as a product of nonzero Eigen values of σ . So, B prime σ B is nonsingular.

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So, X will have density $1/2 \pi$ to the power k/p B prime σ B to the power $1/2$ e to the power $-1/2$ $X - B$ prime μ B prime σ B to the power -1 $X - B$ prime μ , so this description 1 and 2 that describes the density. If you consider say $X - B$ prime μ , B prime

sigma B Inverse X-B prime mu then that is U-mu B B prime sigma B Inverse B prime U-mu = U-mu, a generalized inverse of this U-mu for some twice of sigma g inverse.

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So \underline{X} has density

$$\frac{1}{(2\pi)^{k/2} |B' \Sigma B|^{1/2}} e^{-\frac{1}{2} (\underline{X} - B' \underline{\mu}) (B' \Sigma B)^{-1} (\underline{X} - B' \underline{\mu})}$$

$$= \frac{1}{(2\pi)^{k/2} |B' \Sigma B|^{1/2}} e^{-\frac{1}{2} (\underline{U} - \underline{\mu}) B (B' \Sigma B)^{-1} B' (\underline{U} - \underline{\mu})}$$

$$= \frac{1}{(2\pi)^{k/2} |B' \Sigma B|^{1/2}} e^{-\frac{1}{2} (\underline{U} - \underline{\mu}) \Sigma^{-1} (\underline{U} - \underline{\mu})}$$

for some choice of \underline{U}

So the density is

$$\frac{1}{(2\pi)^{k/2} \sqrt{\frac{1}{|B' \Sigma B|}}} e^{-\frac{1}{2} (\underline{U} - \underline{\mu})' \Sigma^{-1} (\underline{U} - \underline{\mu})}$$

So the density is actually $1/2 \pi$ to the power $k/2$ product of the determinant the Eigen values $i = 1$ to k , e to the power $-$. So, this is actually density of on a subspace. It is not a density on the full space when the rank of sigma is not full. Now before going to the estimation, let us consider one or 2 applications of this conditional distribution or linear combinations etc. One example of a multivariate normal distribution let me write here.

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Example: $\underline{\mu} = \begin{pmatrix} 4 \\ 3 \\ 2 \\ 1 \end{pmatrix}$, $\Sigma = \begin{pmatrix} 3 & 0 & 2 & 2 \\ 0 & 1 & 1 & 0 \\ 2 & 1 & 1 & -2 \\ 2 & 0 & -2 & 4 \end{pmatrix} = \begin{pmatrix} \Sigma_{11} & \Sigma_{12} \\ \Sigma_{21} & \Sigma_{22} \end{pmatrix}$

Let $\underline{X} \sim N_4(\underline{\mu}, \Sigma)$.

$\underline{X} = \begin{pmatrix} X_1 \\ X_2 \\ X_3 \\ X_4 \end{pmatrix} = \begin{pmatrix} X^{(1)} \\ X^{(2)} \end{pmatrix} \rightarrow L$ $\underline{\mu} = \begin{pmatrix} \mu^{(1)} \\ \mu^{(2)} \end{pmatrix}$

Find conditional distⁿ of $X^{(2)}$ from $X^{(1)} = \begin{pmatrix} 3 \\ 2 \end{pmatrix}$.

$\sim N_2 \left(\mu^{(2)} + \Sigma_{21} \Sigma_{11}^{-1} (\underline{x}^{(1)} - \mu^{(1)}), \Sigma_{22} - \Sigma_{21} \Sigma_{11}^{-1} \Sigma_{12} \right)$

$\begin{pmatrix} 2 \\ 1 \end{pmatrix} + \begin{pmatrix} 2 & 1 \\ 2 & 0 \end{pmatrix} \begin{pmatrix} 1/2 & 0 \\ 0 & 1 \end{pmatrix} \begin{pmatrix} -1 \\ -1 \end{pmatrix}$, $\begin{pmatrix} 1 & -2 \\ -2 & 4 \end{pmatrix} = \begin{pmatrix} 2 & 1 \\ 2 & 0 \end{pmatrix} \begin{pmatrix} 1/2 & 0 \\ 0 & 1 \end{pmatrix} \begin{pmatrix} 2 & 2 \\ 1 & 0 \end{pmatrix}$

Let us consider say $\mu = 4, 3, 2, 1$ and I consider sigma as $3, 0, 2, 2, 0, 1, 1, 0, 2, 1, 9, -2, 2, 0, -2, 4$. So let us take say X following $N_4 \mu \sigma$, let us consider some partitioning of this, say it = X_1, X_2, X_3, X_4 , which I am actually writing as a X_1 and X_2 , okay. That is this is 2

dimensional and this is 2 dimensional here. Let us define say conditional distribution of say X_2 given $X_1 = \text{say } 3, 2$.

Now, we have discussed the conditional distribution of one component giving the second component. So this follow N_2 and if I considered the corresponding decomposition of μ as a μ_1 and μ_2 , then this will become $\mu_2 + \sigma_{21}$, so I am partitioning this as σ_{11} , σ_{12} , σ_{21} , σ_{22} . So, if I considered this, then this term is σ_{11} , this is σ_{12} , this is σ_{21} , and this is σ_{22} .

So this will become, so let us calculate these terms here. So this one is now 2, $1 + \sigma_{21}$ is this term here, 2, 1, 2, 0, σ_{11} inverse is the inverse of this that is $1/3, 1, 0, 0$ and then you have $X_1 - \mu_1$, so 3, 2 - μ_1 , so that will become -1, -1 and here I will get 9, -2, -2, 4, $-\sigma_{21}$ that is 2, 1, 2, 0, σ_{11} inverse * σ_{12} that is 2, 2, 1, 0 that is the dispersion term here.

So, I will get here X_2 given $X_1 = 3, 2$ as $N_2(17/3, 11/3, 20/3, -10/3, -10/3, 8/3)$. So, I am able to obtain the conditional distribution of X_2 given $X_1 = \text{a certain number}$. So, this is quite interesting here, you can obtain and you can actually look at this, this is 4, 3, 2, 1 and here you see X_2 given some value of X_1 . So, here you can see that there is a dramatic change here, this is $17/2$ which is bigger than 5 itself, this is around 4.

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$$\underline{X} = \begin{pmatrix} X_1 \\ X_2 \end{pmatrix} = \begin{pmatrix} \underline{X}^{(1)} \\ \underline{X}^{(2)} \end{pmatrix} \sim N_2 \left(\begin{pmatrix} \mu_1 \\ \mu_2 \end{pmatrix}, \begin{pmatrix} \sigma_{11} & \sigma_{12} \\ \sigma_{21} & \sigma_{22} \end{pmatrix} \right)$$

Find conditional dist of $\underline{X}^{(2)}$ given $\underline{X}^{(1)} = \begin{pmatrix} 3 \\ 2 \end{pmatrix}$.

$$\sim N_2 \left(\underbrace{\mu_2 + \sigma_{21} \sigma_{11}^{-1} (\underline{x}^{(1)} - \mu_1)}_{\text{mean}}, \underbrace{\sigma_{22} - \sigma_{21} \sigma_{11}^{-1} \sigma_{12}}_{\text{variance}} \right)$$

$$\begin{pmatrix} 2 \\ 1 \end{pmatrix} + \begin{pmatrix} 2 & 1 \\ 2 & 0 \end{pmatrix} \begin{pmatrix} 1/3 & 0 \\ 0 & 1 \end{pmatrix} \begin{pmatrix} -1 \\ -1 \end{pmatrix}, \quad \begin{pmatrix} 9 & -2 \\ -2 & 4 \end{pmatrix} - \begin{pmatrix} 2 & 1 \\ 2 & 0 \end{pmatrix} \begin{pmatrix} 1/3 & 0 \\ 0 & 1 \end{pmatrix} \begin{pmatrix} 2 & 2 \\ 0 & 1 \end{pmatrix}$$

$$\underline{X}^{(2)} \text{ given } \underline{X}^{(1)} = \begin{pmatrix} 3 \\ 2 \end{pmatrix} \text{ is } N_2 \left(\begin{pmatrix} 17/3 \\ 11/3 \end{pmatrix}, \begin{pmatrix} 20/3 & -10/3 \\ -10/3 & 8/3 \end{pmatrix} \right)$$

And whereas the original means of X_2 was only 2, 1. So, if X_1 is given 3, 2 then it has increase the means of X_2 and similarly, there is substantial change in the value of the

variance-covariance terms here. Let us also define in the same, $A =$ say 1, 2 and let us consider say $B = 1, -2, 2, -1$. What is the distribution of say AX_1 ? So, according to AX_1 will have normal with mean.

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$$A = [1 \ 2] \quad B = \begin{bmatrix} 1 & -2 \\ 2 & -1 \end{bmatrix}$$

$$AX_1 \sim N(A\mu_1, A\Sigma_1 A^T) = N(10, 7)$$

$$BX_2 \sim N_2(B\mu_2, B\Sigma_2 B^T)$$

$$N\left(\begin{pmatrix} 0 \\ 3 \end{pmatrix}, \begin{pmatrix} 33 & 16 \\ 36 & 32 \end{pmatrix}\right)$$

$$\text{Cov}(AX_1, BX_2) = A\Sigma_1 B^T = (0, 6)$$

So $A\mu_1$ because A is a scalar here, A is a row vector here. So this will become a scalar and then you will have $A\Sigma_1 A^T$. So, you can calculate this, this value is simply 10 and this is 7. Similarly, suppose I consider BX_2 , so BX_2 is actually = 2 dimensional vector here that is following normal $B\mu_2, B\Sigma_2 B^T$. So, if you can calculate this, these terms have to be 0, 3, 33, 16, 36, 32.

Let us also consider say covariance between AX_1 and BX_2 , then this will become = $A\Sigma_1 B^T$, so that = 0, 6. So, in this example I have shown you a direct application of the distribution theory of the multivariate normal distribution and various things were considered here. Let me give one more exercise here. Let us consider say $\Sigma = 4, 1, 2, 1, 9, -3, 2, -3, 25$, okay.

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$$\Sigma = \begin{pmatrix} 1 & 1 & 2 \\ 1 & 4 & -3 \\ 2 & -3 & 15 \end{pmatrix}$$

$$\rho = \begin{pmatrix} 1 & \text{cov}(x_1, x_2) & \dots \\ \vdots & \ddots & \ddots \\ \vdots & \vdots & 1 \end{pmatrix} \rightarrow \text{Corr}(X)$$

Find $V^{1/2} \rightarrow$ diagonals are standard deviations

$$V^{1/2} \rho V^{1/2} = \Sigma$$

$$V^{1/2} = \begin{pmatrix} 2 & 0 & 0 \\ 0 & 3 & 0 \\ 0 & 0 & 5 \end{pmatrix}, \quad V^{-1/2} = \begin{pmatrix} 1/2 & 0 & 0 \\ 0 & 1/3 & 0 \\ 0 & 0 & 1/5 \end{pmatrix}$$

$$P = V^{-1/2} \Sigma V^{-1/2} = \begin{bmatrix} 1 & 1/6 & 1/5 \\ 1/6 & 1 & -1/5 \\ 1/5 & -1/5 & 1 \end{bmatrix}$$

So, here I considered rho as say 1, 1, 1 and here I will consider correlation between, so this is actually correlation matrix, correlation matrix of X, okay. So that means these term will be denoting correlation between X1, X2 not covariance, it is correlation terms here. So, consider find V1/2, where these diagonals are standard deviations and find rho and also show that V1/2 rho V1/2 = sigma for this particular case.

See this is an interesting thing because you can do the manipulation with the variance-covariance matrix because of the positive semidefiniteness of the matrix, this is very important because it has a spectral decomposition, you can have a gram decomposition as V V transpose etc. So, lots of nice properties are coming here. Let us consider V1/2 here will become = 2, 3, 5 that is the standard deviation matrix here.

Let us consider V -1/2 so that will be 1/2, 1/3, 1/5, 0, 0, 0, 0, 0, 0 and rho = V to the power -1/2 sigma V to the power -1/2 that = 1, 1/6, 1/5, 1/6, 1, -1/5, 1/5, -1/5 and 1. I mentioned about the uniqueness of the sigma 21, 11 inverse term actually. So see there is a problem, suppose we are calculating the inverses, then sometimes the inverses will not exist or with the little variation, the inverse may vary too much.

So that means it is an example of an unstable matrix. Let us take one case here at least, let me give you example of one such problem. Let us take say A = 4, 4.001, 4.001, 4.002 and B = say 4, 4.001, 4.001, 4.002001, you can see here that in A and B, 3 terms are exactly the same, the 4th term I have modified only by adding 0.000001 okay only that much difference is there.

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$$A = \begin{pmatrix} 4 & 4.01 \\ 4.01 & 4.01 \end{pmatrix}, \quad B = \begin{pmatrix} 4 & 4.01 \\ 4.01 & 4.0201 \end{pmatrix}$$

$$A^{-1} = -10^6 \begin{pmatrix} 4.002 & -4.001 \\ -4.001 & 4 \end{pmatrix}, \quad |A| = -10^{-6}$$

$$B^{-1} = \frac{10^4}{3} \begin{pmatrix} 4.00201 & -4.001 \\ -4.001 & 4 \end{pmatrix}, \quad |B| = 3 \times 10^{-6}$$

$$A^{-1} \approx -3B^{-1} \quad \text{Although } A \approx B$$

⊗ This is an example of unstable system

Let us look at say A inverse, A inverse = -10 to the power 6, 4.002, -4.001, -4.001, 4 and if I look at B inverse then that = 10 to the power 6/3, 4.002001, -4.001, -4.001, 4. So you can see that there is a dramatic change in the value here. Actually, determinant of A is turning out to be -10 to the power -6, whereas determinant of B = 3*10 to the power -6. So, there is substantial change in the value.

So, we are getting that A inverse is approximately -3B inverse. If you look at A and B, there is hardly any difference here. In fact, the 3 terms are exactly the same in the 4th term, I am considering the change after 5 decimal places. In the 6th decimal place, there is a minor change by 0.000001, but if you look at the inverses here, A is almost same as B, but if you look at the inverse, so then you are getting substantial change.

So, this is an example of unstable system. The reason is that if I look at this that they are almost linearly dependent here, see A if you look at this is dependent, so this is almost dependent here because there is a small change. In that case, a small change in the value of one term makes a huge change in the value of B inverse. In the next class, I will be talking about the estimation of the parameters of multivariate normal distribution.

We will also discuss the noncentral chi square distribution etc because that concept is coming here and it will be also used in finding out the distributions of the statistics there, so that I will be taking up in the next lecture.