

**Introduction to Econometrics**  
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**Lecture 55**

**Relaxing the Assumptions of CLRM-Autocorrelation and Heteroscedasity Part-6**  
 So, that means, if you now compare the Goldfeld Quandt test and the BPG test then you are actually arriving at same type of conclusion.

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The screenshot shows the STATA interface with the following data:

Source	SS	df	MS	Number of obs
Model	41886.7134	1	41886.7134	30
Residual	2361.15325	28	84.3269018	
Total	44247.8667	29	1525.78851	

	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
income	.0100632	.0041187	2.44	0.021	-.0016265 .0184999
_cons	-.7426146	.7529284	-0.99	0.332	-2.284918 .7996892

	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
income	.6377846	.0286167	22.29	0.000	.579166 .6964031
_cons	9.290387	5.231386	1.78	0.087	-1.4257 20.00632

Command: `hettest`

The screenshot shows the STATA interface with the following test results:

```

.bptest
-----
Breusch-Pagan / Cook-Weisberg test for heteroscedasticity
H0: Constant variance
Variables: fitted values of consumption
-----
chi2(1) = 5.21
Prob > chi2 = 0.024
    
```

Command: `hettest`

So, in STATA they have a readymade command to conduct this test and that is basically I look, I will explain this, so once you run your original model `reg consumption on income` then the command for BPG and Goldfeld Quandt test is `het test` and see the value is 5.21

exactly the same value what we have derived same value, manually whatever we have calculated STATA is also reporting the same chi square value and what is the probability that the calculated value is greater than the tabulated one look at the p value 0.0224, so that means, if you multiply this by 100 you will get 2.24 which is greater than 1 but, less than 5, so that means, it is significant at 5 percent level. So, you can reject your null hypothesis at 5 percent level of significance. So, this is the Briush Pagan and Goldfeld Quandt test.

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The slide content is as follows:

**BPG Test:**  $y_i = \beta_0 + \beta_1 x_{1i} + \beta_2 x_{2i} + \beta_3 x_{3i} + \dots + \beta_k x_{ki} + u_i \dots \textcircled{1}$

**step 1:** estimate model  $\textcircled{1}$  and get the RSS  

$$\hat{\sigma}_{ML}^2 = \frac{RSS}{n}$$

$$\left( \frac{RSS}{n-k} \right) = \frac{\lambda^2}{\sigma_{OLS}^2}$$

**step 2:**  $\frac{u_i}{\hat{\sigma}_{ML}^2} = \hat{\beta}$

**step 3:**  $\hat{\beta} = \lambda_0 + \lambda_1 x_{1i} + \lambda_2 x_{2i} + \dots + \lambda_k x_{ki} + \epsilon_i \dots \textcircled{2}$   
 and collect ESS from  $\textcircled{2}$   $H_0: \lambda_1 = \lambda_2 = \dots = \lambda_k = 0$

**step 4:**  $\frac{1}{2} ESS \sim \chi_{df}^2$  where  $m$  is the total number of parameters in the original model  
 $\chi_{cal}^2 > \chi_{tab}^2 \Rightarrow$  reject  $H_0$   
 $\chi_{cal}^2 = \frac{1}{2}(10 \cdot 42) = 5 \cdot 21$   
 $\chi_{tab}^2(5\%) = 3 \cdot 23 \Rightarrow \chi_{cal}^2 > \chi_{tab}^2$  at 5% sig

But, there are 2 limitations of this BPG test also. First of all the chi square what I said, this chi square this ESS by 2 is asymptotically follows chi square distribution that means, this particular test statistic follow the chi square distribution only in large sample only in large sample also the sigma square tilde, what we got here to adjust your  $U_i$  hat square and that is you see that is the definition of sigma square from the maximum likelihood method and maximum likelihood method is applicable when you have a large sample, but here you have only 30 observations but, these 30 observations may not actually define a large sample. So, that is limitation number 1 half of ESS asymptotically follow chi squared distribution that means, this particular test statistic you can derive only when your sample size is large.

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**Limitations of BPG test**

- (1)  $\frac{1}{2} ESS \xrightarrow{asymp} \chi^2 \Rightarrow$  this test is applicable only in large sample
- (2)  $U_i \sim N \Rightarrow$  this test largely depends on the normality assumption of  $U_i$  in the original model

**White's general het. test**

-  $y_i = \alpha + \beta_1 x_{1i} + \beta_2 x_{2i} + U_i \dots \text{--- (1)}$

-  $U_i^2 = \lambda_0 + \lambda_1 x_{1i} + \lambda_2 x_{2i} + \lambda_3 x_{1i}^2 + \lambda_4 x_{2i}^2 + \lambda_5 x_{1i} x_{2i} + \epsilon_i$

-  $n \times R^2 = \chi^2_{df=k}$  (total number of parameters in the auxiliary reg. excluding the intercept.)

$H_0: \lambda_1 = \lambda_2 = \dots = \lambda_5 = 0$

F test

So, limitations of BPG test first one is that half of ESS follows asymptotically chi square distribution that means, this implies this test is applicable only in large sample. second limitation is that in your original model  $U_i$  actually follows a normal distribution so these test largely depend on the normality assumption of the error term in the original model, so that means this implies this test largely depends on the normality assumption of  $U_i$  in the original model.

To overcome these 2 limitations, there is one more test which is called White General Heteroscedasticity test in short I will write het test, how is this test suppose your original model is  $y_i$  equals to  $\alpha$  plus  $\beta_1 x_{1i}$  plus  $\beta_2 x_{2i}$  plus  $U_i$  this is your original model and here what they say the test is very simple you estimate this model and then, collect  $U_i$  hat square and then you regress this  $U_i$  hat square on  $\lambda_1 x_{1i}$  plus  $\lambda_2 x_{2i}$  plus  $\lambda_3 x_{1i}^2$  plus  $\lambda_4 x_{2i}^2$  plus  $\lambda_5 x_{1i} x_{2i}$  plus  $\epsilon_i$  and then, get this is step 1 this is step 2 and then from here you get  $R^2$  and in step 3 if you multiply these  $R^2$  square that will follow again  $\chi^2$  distribution with the degrees of freedom equals to  $k$  this is let us say equation 1, this is equation 2 which is also known as auxiliary regression.

$k$  equals to total number of explanatory variable in or I will say the total number of parameters in the auxiliary regression but excluding the intercept, so that means from the auxiliary regression alternatively, you can check instead of constructing  $n$  multiplied by  $R^2$  square, you can also test  $\lambda_1$  equals to  $\lambda_2$  equals to dot dot dot  $\lambda_5$  equals



Now, why White's test is called general Heteroscedasticity test can you think of, see the specification of a white test you have  $U_i$  hat squared equals to all the explanatory variable it is square and cross product, if this cross product terms are significantly they are 0 if you have more number of explanatory variables you will have more cross product.

So, if the cross-product terms are 0, so that means it is a pure test of Heteroscedasticity. if the cross product terms are not 0 then, this is a test of Heteroscedasticity as well as model misspecification that is why it is called a General test because you are testing Heteroscedasticity as well as model misspecification. Now, you can take the same data set and again you can manually conduct this White's General Heteroscedasticity test and that you take as an assignment, so conduct this test at home conduct.

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The screenshot displays the Stata 16.0 interface with the following content:

**Command History:**

```

1  reg consumption income
2  gen sigmaq=2361.15/30
3  predict sigmaid
4  gen sigmaq=sq2
5  gen p=sigmaqsigmaq
6  reg p income
7  reg consumption income
8  hettest

```

**Regression Results:**

	income	_cons					
	.0100632	-.0041187	2.44	0.021	-.0016265	-.0184999	
		-.7426146	-.7529284	-0.99	0.332	-2.284918	-.7996892

**Model Summary:**

Source	SS	df	MS	Number of obs =	F(1, 28) =
Model	41886.7134	1	41886.7134	30	496.72
Residual	2361.15325	28	84.3269018		Prob > F = 0.0000
Total	44247.8667	29	1525.78851		R-squared = 0.9466
					Adj R-squared = 0.9417
					Root MSE = 9.183

**Parameter Estimates:**

consumption	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
income	-.6377846	-.0286167	22.29	0.000	-.791566	-.4964031
_cons	9.290307	5.231306	1.78	0.087	-1.4257	20.00632

**White's Test Results:**

```

. hettest
-----
Breusch-Pagan / Cook-Weisberg test for heteroskedasticity
Ho: Constant variance
Variables: fitted values of consumption

chi2(1) = 5.21
Prob > chi2 = 0.0224

```

**Command:** reg consumption inco

STATA 16.0

History

```

1  reg consumption income
2  gen sigmaq=281.55/30
3  predict resid
4  gen uq=u^2
5  gen p=sigmaq*sigmaq
6  reg p income
7  reg consumption income
8  hettest
9  reg consumption income

```

Command: `hettest`

Breusch-Pagan / Cook-Weisberg test for heteroskedasticity  
 Ho: Constant variance  
 Variables: fitted values of consumption

chi2(1) = 5.21  
 Prob > chi2 = 0.0224

Command: `reg consumption income`

Source	SS	df	MS	F(1, 28)	Prob > F
Model	41886.7134	1	41886.7134	496.72	0.0000
Residual	2361.5525	28	84.3269018	R-squared = 0.9466	
Total	44247.8667	29	1525.78851	adj R-squared = 0.9467	Root MSE = 9.183

consumption	Coeff.	Std. Err.	t	P> t	[95% Conf. Interval]
income	-.6377846	.0286167	22.29	0.000	-.5791566 - .6964031
_cons	9.290307	5.231386	1.78	0.087	-1.4257 20.00632

Command: `imtest,white`

STATA 16.0

History

```

1  reg consumption income
2  gen sigmaq=281.55/30
3  predict resid
4  gen uq=u^2
5  gen p=sigmaq*sigmaq
6  reg p income
7  reg consumption income
8  hettest
9  reg consumption income
10 imtest,white

```

Command: `imtest,white`

White's test for Ho: homoskedasticity  
 against Ha: unrestricted heteroskedasticity

chi2(2) = 5.33  
 Prob > chi2 = 0.0696

Cameron & Trivedi's decomposition of IM-test

Source	chi2	df	p
Heteroskedasticity	5.33	2	0.0696
Skewness	1.56	1	0.2117
Kurtosis	0.00	1	0.9647
Total	6.89	4	0.1417

Command: `imtest,white`

Now, here I will just show you how to do it in STATA. so in STATA once again if you run the original model, reg consumption and then income after that you have to do imtest white, what is the command I am saying imtest white, that is the command for white's general Heteroscedasticity test and from there or you can do or you can see that chi square value is 5.33 and the p value is 0.0696 that means it is significant only a 10 percent level.

But, once again you have to remember that even the White's General Heteroscedasticity test is also a large sample test, which may not be applicable in a small sample in this case, because you have only 30 number of observations, why this is a large sample test?

Because see in your auxiliary regression and  $U_i$  hat squared is a function of all your explanatory variable it square and cross product, so with 2 explanatory variable with only 2

explanatory variable you are having 5 parameters, so you can imagine if you have 5 then all the 5 explanatory variable is square and cross product that means 5c2 terms cross product times cross product terms would be there, so many explanatory variable will appear in the auxiliary regression and that will eat up lot of degrees of freedom, so that is why you need to have a large sample to conduct this particular test also.

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Now, the question is, we have detected Heteroscedasticity but what is the solution and how will you solve. So, the solution is in your model  $y_i$  equals to  $\alpha$  plus  $\beta x_i$  plus  $u_i$  and you are saying that  $\sigma_i^2 = \sigma_i^2$  equals to  $\sigma_i^2$  so, that means variance of  $u_i$  in STATA I will write variance of  $U_i$  equal to  $\sigma_i^2$ , so there are 2 cases when  $\sigma_i^2$  is known to you somehow, you know the error variance  $\sigma_i^2$  is known to you, so in that case if you simply divide this equation in STATA of estimating that original equation, you just divide the equation by  $\sigma_i$  plus  $\beta x_i$  by  $\sigma_i$  plus  $U_i$  by  $\sigma_i$  you divide the equation and then run OLS in the modified regression, if you run OLS in this modified regression, that is called generalised least square method.

Now, the question is with OLS is not applicable in the original model, how come we are able to apply OLS in the transform model, because if you calculate the variance of these  $U_i$  by  $\sigma_i$  then that is nothing but expectation of  $U_i$  divided by  $\sigma_i^2$ , which is nothing but  $1 / \sigma_i^2$  into expectation of  $U_i^2$  and what is expectation of  $U_i^2$  that is also  $\sigma_i^2$  equals to  $1 / \sigma_i^2$  into  $\sigma_i^2$  equals to 1. So,

that means variance is constant in the modified regression while it was not in the original regression.

Now, if you look at the procedure that means what I am doing in the original regression in this regression while doing OLS we are trying to minimise  $\hat{U}_i$  square and here what I am trying to minimise in this generalised model, we are trying to minimise  $U_i$  divided by  $\sigma_i$  square, so this method is also known as weighted least square method. Where,  $1/\sigma_i$  square is basically is the weight, so that means higher the variance, lower is the weight you are attaching to that particular error term.

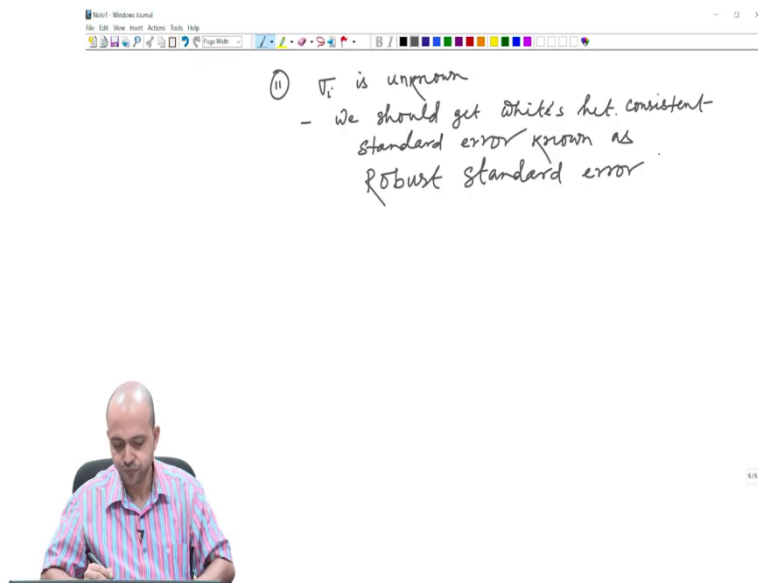
Now, how it is different from OLS? This is  $y_i$  in OLS, what was happening, so this was where error term this is my error term suppose, this one is  $U_1$  and this is  $U_4$ , so when you are minimising summation  $U_i$  square that means you are basically minimising summation  $U_1$  square plus  $U_2$  square plus dot dot dot  $U_n$  square, so that means, you are attaching same weight for all these error terms but, here in the WLS or Weighted Least Square what I am doing am attaching  $1/\sigma_i$  as the weight.

That means, higher that means, when your distance is too far from the predicted line I will attach only the lower weight, so weight is inversely proportional to the distance of the error term from this so that means, higher the value of the error term lower the weight I am attaching to that particular error term and that should be realist that should be justified also because, this error term contributes less towards constructing this regression line, why should I attach too much weight to this, why should I attach similar weight to  $U_4$  with that of  $U_1$ ,  $U_1$ , because this is too close to the predicted line.

So, that means this is contributing more, this is more important than this error term if you put equal weight then  $U_4$  will dominate the summation  $\hat{U}_i$  square that is not actually happened that is the problem of OLS that is why in WLS I am saying am attaching weight and weight is actually inversely proportional to the value of this. Higher the  $U$  lower would be the weight attached to it that is the logic but, the question is this is all right, you can transform the equation when  $\sigma_i$  square is known. But, when it is unknown, they not will do?



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When  $\sigma_i$  is unknown we should get White's Heteroscedasticity consistent standard error actually because, ultimately, your standard error will get affected in presence of Heteroscedasticity, so you have to get this and this particular standard error White Heteroscedasticity consistent standard error that is known as Robust standard error.

And nowadays in all these statistical software's they routinely compute the Robust standard error which is basically White Heteroscedasticity consistent standard error that means assuming there is a presence of heteroscedasticity what is the standard error in large difference between the OLS standard error and this Robust standard error indicates that your data is basically suffering from Heteroscedasticity problem and if you go to your data set once again.

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Stata 16.0

```

History
1 reg consumption income
2 gen sigmaq = 2361.15/30
3 predict sigmaid
4 gen sigmaq^2
5 gen p = sigmaq/sigmaid
6 reg income
7 reg consumption income
8 hettest
9 reg consumption income
10 intest,white

```

	Total	44247.8667	29	1525.78851	Root MSE	=	9.183
consumption	Cof.	Std. Err.	t	P> t	[95% Conf. Interval]		
income	.6377846	.0260317	22.29	0.000	-.579186	-.6964831	
_cons	9.290367	5.231366	1.78	0.087	-1.4257	20.0632	

White's test for H0: homoskedasticity  
against H1: unrestricted heteroskedasticity

chi2(2) = 5.33  
Prob > chi2 = 0.0696

Cameron & Trivedi's decomposition of IM-test

Source	chi2	df	p
Heteroskedasticity	5.33	2	0.0696
Skewness	1.56	1	0.2117
Kurtosis	0.00	1	0.9647
Total	6.89	4	0.1417

Command  
reg consumption income,robust

Stata 16.0

```

History
1 reg consumption income
2 gen sigmaq = 2361.15/30
3 predict sigmaid
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8 hettest
9 reg consumption income
10 intest,white
11 reg consumption income,robust

```

Prob > chi2 = 0.0696

Cameron & Trivedi's decomposition of IM-test

Source	chi2	df	p
Heteroskedasticity	5.33	2	0.0696
Skewness	1.56	1	0.2117
Kurtosis	0.00	1	0.9647
Total	6.89	4	0.1417

Linear regression

Number of obs = 30  
F(1, 28) = 457.18  
Prob > F = 0.0000  
R-squared = 0.9466  
Root MSE = 9.183

	Robust					
consumption	Cof.	Std. Err.	t	P> t	[95% Conf. Interval]	
income	.6377846	.0260283	21.38	0.000	-.576684	-.6988851
_cons	9.290367	4.43783	2.09	0.045	-1.998248	18.38079

Command

Stata/16.0

File Edit Data Graphics Statistics User Window Help

History

```

1  reg consumption income
2  gen sigmaeq=281.15/30
3  predict uresid
4  gen uresq=u^2
5  gen p=nsq/sigmaeq
6  reg p income
7  reg consumption income
8  hettest
9  reg consumption income
10 intestatwhite
11 reg consumption income,robust

```

Command: `reg consumption income`

Cameron & Trivedi's decomposition of H-test

Source	chi2	df	p
Heteroskedasticity	5.33	2	0.0696
Skewness	1.56	1	0.2117
Kurtosis	0.60	1	0.9647
Total	6.89	4	0.1417

`. reg consumption income,robust`

Linear regression

	Number of obs =	30
F(1, 28)	=	457.18
Prob > F	=	0.0000
R-squared	=	0.9466
Root MSE	=	9.183

	Robust					
consumption	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
income	-.6377846	.6298283	21.38	0.000	-.5766884	-.6988851
_cons	9.290367	4.43783	2.09	0.045	-.1998248	18.38079

Command: `reg consumption income`

Stata/16.0

File Edit Data Graphics Statistics User Window Help

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10 intestatwhite
11 reg consumption income,robust
12 reg consumption income

```

Command: `reg consumption income`

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`. reg consumption income`

Source	SS	df	MS	Number of obs =	30
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Residual	2361.15325	28	84.3269018	R-squared	= 0.9466
Total	44247.8667	29	1525.78851	Adj R-squared	= 0.9447
				Root MSE	= 9.183

	Robust					
consumption	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
income	-.6377846	.6298283	21.38	0.000	-.5766884	-.6988851
_cons	9.290367	5.231386	1.78	0.087	-1.4257	20.00632

Command: `reg consumption income`

So, if you run this model `reg consumption` then income after that you have to put a Robust command to get a Robust standard error. Now, you see the standard error for the income coefficient is 0.0298 and in your original model, if you run simply `reg consumption` on income then your standard error is basically 0.0286 so there is not much of a difference in the standard error that means, the presence of Heteroscedasticity is not so severe and that was prominent from the White Heteroscedasticity test also that you can reject the null only at 10 percent level, this is how you can get a Robust standard error to solve the Heteroscedasticity problem, if at all anything is there. So, with this we are basically closing our discussion on heteroscedasticity.

Now, you see, we have discussed autocorrelation we have discussed multicollinearity and now, we have discussed Heteroscedasticity that means, we have relaxed only 3 assumptions out of 10 we mentioned in the context of classical linear regression model, that means, we assume other 7 assumptions are maintained out of those 7 assumptions, there is 1 more assumption which is also very important known as.

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(ii)  $\sigma^2$  is unknown  
 - we should get White's het. consistent standard error known as Robust standard error

$y_i = \alpha + \beta x_i + u_i$   
 non-stochastic / exogenous  
 $\text{Cov}(x_i, u_i) = 0$   
 But in case  $\text{Cov}(x_i, u_i) \neq 0$   
 $\Rightarrow x_i$  is endogenous

- Endogeneity Problem  
 Instrumental Variable Estimation Technique

Let us have  $y$  equals to  $\alpha$  plus  $\beta x$  plus  $u$  where we assume this  $x$  is basically non stochastic or exogenous that means, covariance of these  $x$  and  $u$  is actually 0, this is not correlated with this error term but in case covariance between  $x$  and  $u$  is not equal to 0, then that is called  $x$  is endogenous and it will lead to endogeneity problem.

So, this endogeneity problem if you have in your data set, then your OLS method is not applicable, you have to use instrumental variable estimation technique but, this IV technique or instrumental variable estimation technique is beyond the scope of this basic econometrics the course what we are dealing with that is that comes under the next course, which is applied econometrics.

But, you should aware of this problem also, this is also one of the assumption that we maintained in the context of classical linear regression model that all our explanatory variables, they are exogenous in nature and that means, they are not correlated with the error term. So, with this we are closing our discussion of this module that means relaxing the assumption of classical linear regression model Autocorrelation, Multicollinearity and

Heteroscedasticity detection of the problem consequences and the solution we had discussed.

Thank you.